UNIVERSITY OF CALIFORNIA

Los Angeles

Probability and Symmetry in Computational Linear Algebra

A dissertation submitted in partial satisfaction of the requirements for the degree Doctor of Philosophy in Mathematics

by

Man-Chung Yeung

1997
The dissertation of Man-Chung Yeung is approved.

Bjorn Engquist

Stott Parker

Eitan Tadmor

Tony F. Chan, Committee Chair

University of California, Los Angeles

1997
TABLE OF CONTENTS

1 Introduction .................................................. 1
  1.1 Gaussian Elimination ................................. 2
  1.2 Krylov subspace methods ............................ 6
  1.3 Contributions of this thesis ......................... 10
    1.3.1 Probabilistic analysis of GE without pivoting ...... 10
    1.3.2 Symmetry in preconditioned Krylov subspace methods ... 12
    1.3.3 ML(k)BiCGSTAB ................................. 14

2 Probabilistic Analysis of Gaussian Elimination Without Pivoting and Its Application ......................... 19
  2.1 Properties of a Gaussian Matrix ...................... 19
  2.2 Real Gaussian Elimination ............................ 21
    2.2.1 Density Functions of $u_{pq}$ and $l_{pq}$ .......... 21
    2.2.2 Probability of Small Pivot ....................... 29
    2.2.3 Probability of Large Growth Factor ............... 30
    2.2.4 Numerical Experiments ............................ 34
  2.3 Complex Gaussian Elimination ........................ 37
    2.3.1 Density Functions of $u_{pq}$ ................... 38
2.3.2 Density Functions of $l_{pq}$ ........................................... 43

2.4 Application ............................................................... 45

3 Preserving Symmetry in Preconditioned Krylov Subspace Methods ........................................... 54

3.1 Symmetric preconditioning in GMRES ............................. 55

3.2 Truncated iterative methods .......................................... 60

3.3 Symmetric preconditioning in Bi-CG ............................ 65

3.4 Numerical Experiments .................................................. 67

3.4.1 Truncated iterative methods ...................................... 68

3.4.2 Breakdown behavior of Bi-CG ................................. 72

3.5 Conclusions .............................................................. 76

4 ML($k$)BiCGSTAB: A BiCGSTAB Variant Based on Multiple Lanczos Starting Vectors ................................. 77

4.1 A Lanczos method for $k$ Left Starting Linear Independent Vectors .................................. 78

4.2 ML($k$)BiCG: A BiCG Variant Based on Multiple Starting Vectors .................................. 82

4.3 ML($k$)BiCGSTAB: A BiCGSTAB Variant Based on Multiple Starting Vectors Lanczos ....................... 87

4.4 Numerical Experiments .................................................. 107

4.5 Appendix ................................................................. 115
Bibliography ........................... 120
LIST OF FIGURES

2.1 (a) Distribution of $u_{12,12}$: observed(+), predicted(-).  (b) Distribution of $u_{31,31}$: observed(+), predicted(-). 

2.2 (a) Distribution of $l_{13,12}$: observed(+), predicted(-).  (b) Distribution of $l_{30,25}$: observed(+), predicted(-).

2.3 (a) Overlap of Figures 1(a) and 1(b).  (b) Percentage frequency distributions of $\rho_L$ (dashed) and $\rho_U$ (solid).  $n = 25$.

2.4 Percentage frequency distributions of $\rho_L$ (dashed) and $\rho_U$ (solid).

(a) $n = 50$.  (b) $n = 100$.

2.5 (a) Distribution of $u_{12,12}^R$: observed(+), predicted(-).  (b) Distribution of $|u_{12,12}|$: observed(+), predicted(-).

2.6 (a) Distribution of $u_{31,31}^R$: observed(+), predicted(-).  (b) Distribution of $|u_{31,31}|$: observed(+), predicted(-).

2.7 (a) Distribution of $l_{13,12}^R$: observed(+), predicted(-).  (b) Distribution of $|l_{13,12}|$: observed(+), predicted(-).

2.8 Observed distribution (+) of the elements of $U_{22}$ vs. the $N(0,1.77)$ distribution (-).  $k = 20$, $n = 51$.  

vi
2.9 (a) Observed distribution (\( \cdot \)) of the (12, 12)-th entry of \( \tilde{U}_{22} \) vs. the predicted distribution \( 0.7f_{u_{12,12}}(0.7t) (\cdot) \). (b) Observed distribution (\( \cdot \)) of the (31, 31)-th entry of \( \tilde{U}_{22} \) vs. the predicted distribution 0.8\( f_{u_{31,31}}(0.8t) (\cdot) \). The functions \( f_{u_{12,12}}(t) \) and \( f_{u_{31,31}}(t) \) are defined in Theorem 2.1. 

2.10 (a) Observed distribution (\( \cdot \)) of the (13, 12)-th entry of \( \tilde{L}_{22} \) vs. the predicted distribution \( f_{l_{13,12}}(t) (\cdot) \). (b) Observed distribution (\( \cdot \)) of the (30, 29)-th entry of \( \tilde{L}_{22} \) vs. the predicted distribution \( f_{l_{30,29}}(t) (\cdot) \). The functions \( f_{l_{13,12}}(t) \) and \( f_{l_{30,29}}(t) \) are defined in Theorem 2.2.

2.11 Percentage frequency distributions of \( \rho_{L_{22}} \) (dashed) and \( \rho_{\tilde{U}_{22}} \) (solid). (a) \( k = 20, n = 45 \). (b) \( k = 20, n = 70 \). See Figures 2.3 and 2.4 for the meaning of \( \alpha \).

2.12 Percentage frequency distributions of \( \rho_{L_{22}} \) (dashed) and \( \rho_{\tilde{U}_{22}} \) (solid). (a) \( k = 20, n = 120 \). (b) \( k = 20, n = 220 \). See Figures 2.3 and 2.4 for the meaning of \( \alpha \).

3.1 Minimum cosines in right-preconditioned Bi-CG (solid line) and symmetrically-preconditioned Bi-CG (dashed line) for the \( Re = 1 \) problem.

4.1 Example 1: with no preconditioning

4.2 Example 2: with no preconditioning

4.3 Example 2: with ILU(0) preconditioning
4.4 Example 3: with ILU(0) preconditioning . . . . . . . . . . . . . . 112

4.5 Number of matvec's / $10n$ vs. $k$ for the matrices HOR131 (a) and ORSIRR1 (b). "o" denotes no convergence within $10n$ matvec's.

For $k = 1$, we plotted the number of matvec's for BiCGSTAB. Note that there is a dramatic improvement in performance as $k$ increases from 1, but that for $k \geq 30$, the performance is not sensitive to $k$. . 114
LIST OF TABLES

2.1 Probabilities of small pivot. ........................... 35

2.2 12500 matrices from Class 1 were selected in this experiment. B: the number of matrices which make GE without pivoting performed on $U_{22}$ breakdown; EP: the empirical probability of breakdown when GE without pivoting performed on $U_{22}$ ........................... 52

3.1 Mat-Vec’s for convergence for right-preconditioned methods. ...... 70

3.2 Mat-Vec’s for convergence for symmetric right-preconditioned methods. ........................... 71

3.3 Frequencies of minimum cosines for right-preconditioned (first row of each pair of rows) and symmetrically-preconditioned (second row of each pair of rows) Bi-CG. ........................... 74

3.4 Frequencies of minimum cosines when $r^*_0$ is chosen randomly. ... 75

3.5 Steps and minimum cosines for the driven cavity problem. .......... 75

4.1 Average cost per step of the preconditioned ML(k)BiCGSTAB. ... 107
4.2 Comparison of Methods on a Representative Group of Matrices from the Harwell-Boeing Collection. ML(25), ML(50) and ML(100) stand for ML(25)BiCGSTAB, ML(50)BiCGSTAB and ML(100)BiCGSTAB respectively. The numbers in the table are number of matvec’s. "-" means no convergence within 20n matvec’s for BiCG and 10n for the other methods, “b” denotes breakdown, “o” denotes overflow.
ACKNOWLEDGMENTS

I wish to express my sincere gratitude to:

My thesis advisor, Prof. Tony Chan, for his patience, excellent guidance and encouragement throughout the completion of the thesis. His careful reading of this thesis and his constructive suggestions are also very much appreciated.

My committee members, Profs Bjorn Engquist, Stott Parker and Eitan Tadmor for their involvement in this project.

Prof. Alan Edelman for providing me with several important references and suggestions on an earlier draft of Chapter 2 which improves the derivations of equations (2.1) and (2.5). In addition, I would like to thank Profs Gene Golub and Nick Trefethen for their valuable comments and suggestions on the earlier draft of Chapter 2.

Dr. Edmond Chow and Prof. Youcef Saad for their kindly help and collaboration in the research of Chapter 3. In addition, sincere thanks go to Prof. H. A. van der Vorst for his valuable comments on an earlier draft of Chapter 4.
The faculty, staff and students of the Department of Mathematics for their help and supports.

The author was supported by the National Science Foundation under contract ASC-92-01266, ONR under contract ONR-N00014-92-J-1890 and by the Army Research Office under contract DAAL-03-9-C-0047 (Univ. of Tenn. subcontract ORA 4466.04, Amendment 1 and 2).


ABSTRACT OF THE DISSERTATION

Probability and Symmetry in Computational Linear Algebra

by

Man-Chung Yeung

Doctor of Philosophy in Mathematics

University of California, Los Angeles, 1997

Professor Tony F. Chan, Chair

This thesis is divided into two parts. In the first part (Chapter 2), we study
Gaussian elimination by means of probabilistic analysis and in the second one
(Chapters 3 and 4), we study how to preserve symmetry in preconditioned Krylov
subspace methods and develop a new Krylov subspace solver for the solution of a
nonsymmetric linear system $Ax = b$.

The numerical instability of Gaussian elimination is proportional to the size
of the $L$ and $U$ factors that it produces. The worst case bounds are well known.
For the case without pivoting, breakdowns can occur and it is not possible to
provide a priori bounds for $L$ and $U$. For the partial pivoting case, the worst
case bound is $O(2^n)$, where $n$ is the size of the system. Yet, these worst case
bounds are seldom achieved, and in particular Gaussian elimination with partial pivoting is extremely stable in practice. Surprisingly, there has been relatively little theoretical study of the "average" case behaviour. The purpose of our study in Chapter 2 is to provide a probabilistic analysis of the case without pivoting. The distribution we use for the entries of $A$ is the normal distribution with mean 0 and unit variance. We first derive the distributions of the entries of $L$ and $U$. Based on this, we prove that the probability of the occurrence of a pivot less than $\epsilon$ in magnitude is $O(\epsilon)$. We also prove that the probabilities $\text{Prob}(||U||_\infty/||A||_\infty > n^{2.5})$ and $\text{Prob}(||L||_\infty > n^3)$ decay algebraically to zero as $n$ tends to infinity. Numerical experiments are presented to support the theoretical results. By combining our theoretical results and the observations of Trefethen and Schreiber [57] on random matrices, we propose a new method of using GE which may be particularly suited to certain kinds of parallel computers, which greatly reduces data movements while avoiding breakdown and instability effectively.

In Chapter 3, we consider the problem of solving a linear system $Ax = b$ when $A$ is nearly symmetric and when the system is preconditioned by a symmetric positive definite matrix $M$. In the symmetric case, one can recover symmetry by using $M$-inner products in the Conjugate Gradient Algorithm. Inner products can also be used in the nonsymmetric case, and near symmetry can be preserved similarly. We explore the implementation issues associated with this technique and compare a few different options.
We present in Chapter 4 a variant of the popular BiCGSTAB method for solving nonsymmetric linear systems. The method, which we denote by ML($k$)BiCGSTAB, is derived from a variant of the BiCG method based on a Lanczos process using multiple ($k > 1$) starting left Lanczos vectors. Compared with the original BiCGSTAB method, our new method produces a residual polynomial which is of lower degree after the same number of steps but which also requires fewer matrix-vector products to generate, on average requiring only $1 + 1/k$ matvec's per step. Empirically, it also seems to be more stable and faster convergent. The new method can be implemented as a $k$-term recurrence and can be viewed as a bridge connecting the Arnoldi-based FOM/GMRES methods and the Lanczos-based BiCGSTAB methods.
CHAPTER 1

Introduction

One of the most frequently encountered tasks in numerical computation is solving a square and usually large linear system

\[ Ax = b. \]

There are two types of approach for the solution of the system: direct methods, which are basically variations of Gaussian elimination (GE), and iterative ones, which come in many flavours. Direct methods are chosen when \( A \) is dense and unstructured. When \( A \) does not fall into this category, iterative methods are often of interest. Among iterative methods, Krylov subspace methods feature short term recurrences and therefore require low cost and low memory. In this Chapter, we give a description of GE and Krylov subspace methods in §1.1 and §1.2 respectively and summarize the contributions of this thesis in §1.3.
1.1 Gaussian Elimination

GE is the most common general method for solving an square, dense, unstructured linear system

\[ Ax = b, \]

(1.1)

where \( A \) is an \( n \times n \) real or complex matrix and \( b \) is a real or complex vector of length \( n \). The basic idea of the method is to convert the system (1.1) to an equivalent triangular system. The conversion is built on the following theorem [31].

**Theorem 1.1** If all the principal submatrices of \( A \) are nonsingular, then there is a unit lower triangular matrix \( L \) and an upper triangular matrix \( U \) with nonzero diagonal entries such that

\[ A = LU. \]

If \( A \) is nonsingular, then there is a permutation matrix \( P \), a unit lower triangular matrix \( L \) with no entry bigger than 1 in absolute value and an upper triangular matrix \( U \) with nonzero diagonal entries such that

\[ PA = LU. \]

If \( A \) is nonsingular, then there are two permutation matrices \( P \) and \( Q \), a unit lower triangular matrix \( L \) and an upper triangular matrix \( U \), where no entry in \( L \) or \( U \) is bigger than 1 in absolute value and all the diagonal entries of \( U \) are
nonzero, such that

\[ PAQ = LU. \]

Corresponding to the different cases in Theorem 1.1, there are three fundamental versions of GE:

- **GE without pivoting**: Compute \( A = LU \) → Solving \( Ly = b \) → Solving \( Ux = y; \)

- **GE with partial pivoting**: Compute \( PA = LU \) → Solving \( PLy = b \) → Solving \( Ux = y; \)

- **GE with complete pivoting**: Compute \( PAQ = LU \) → Solving \( PLy = b \) → Solving \( UQx = y; \)

For the details of the implementations of these GE versions, one is referred to [31].

The conditions in Theorem 1.1 are only sufficient. In some cases, the \( LU \) factorization of \( A \) need not exist. For example, one can not find \( l_{ij} \) and \( u_{ij} \) such that

\[
\begin{bmatrix}
1 & 2 & 3 \\
2 & 4 & 7 \\
1 & 6 & 3
\end{bmatrix}
= 
\begin{bmatrix}
1 & 0 & 0 \\
l_{21} & 1 & 0 \\
l_{31} & l_{32} & 1
\end{bmatrix}
\begin{bmatrix}
u_{11} & u_{12} & u_{13} \\
0 & u_{22} & u_{23} \\
0 & 0 & u_{33}
\end{bmatrix}.
\]

To see this, one can equate entries to get

\[ u_{11} = 1; \]

\[ l_{21}u_{11} = 2 \rightarrow l_{21} = 2/u_{11} = 2; \]

\[ l_{31}u_{11} = 1 \rightarrow l_{31} = 3/u_{11} = 1; \]
\[ u_{12} = 2; \]
\[ l_{21}u_{12} + u_{22} = 4 \rightarrow u_{22} = 0. \]

But when one then looks at the (3, 2)-th entry one obtains the contradictory equation \( 6 = l_{31}u_{12} + l_{32}u_{22} = 4. \)

In practice, if some diagonal entry \( u_{i_0i_0} \) of \( U \) is zero or close to zero, say, \( u_{i_0i_0} \) is less than the machine precision in absolute value, then the \( LU \) factorization of \( A \) can not be computed, or in other words, GE breaks down. This is because \( u_{i_0i_0} \) will be used as a denominator in the next step of the computations. Among the three versions of GE above, it is always true that GE with partial or complete pivoting has less chance to break down than GE without pivoting. This can be understood in part from Theorem 1.1 since the condition that all the principal submatrices of \( A \) are nonsingular implies that \( A \) is nonsingular but not true conversely. In the literature of GE, the diagonal elements of \( U \) are called pivot elements.

When GE is performed on the system (1.1) in floating point arithmetic, the computed \( LU \) factors \( \tilde{L} \) and \( \tilde{U} \) are produced. Then, by solving two corresponding triangular systems, we obtain the solution \( \tilde{x} \) to (1.1). The computed solution \( \tilde{x} \) satisfies

\[ (A + E) \tilde{x} = b \]

with

\[ |E| \leq nu \left( 3 |A| + 5 |\tilde{L}| |\tilde{U}| \right) + O(u^2) \]
in the case of GE without pivoting and

\[ |E| \leq n u \left( 3 |A| + 5 \hat{P}^T |\hat{L}| |\hat{U}| \right) + O(u^2) \]

in the case of GE with partial pivoting, where \( \hat{P} \) is the computed analogs of \( P \).

Here \( u \) is the unit roundoff and for any matrix \( M \), we use \( |M| \) to denote the matrix obtained by taking the absolute value of the elements of \( M \) [31, P.106, P.115]. From this, it follows that

\[ \|E\|_\infty \leq n u \|A\|_\infty \left( 3 + 5 \|\hat{L}\|_\infty \frac{\|\hat{U}\|_\infty}{\|A\|_\infty} \right) + O(u^2). \]

We define

\[ \rho_L = \|L\|_\infty, \quad \rho_U = \|U\|_\infty / \|A\|_\infty \]

(1.2)

and call \( \rho_L \) and \( \rho_U \) the growth factors. It is obvious that the smaller the growth factors are, the more accurate the computed solution \( \hat{x} \) is.

GE with partial pivoting is extremely stable in practice. However, this stability can not be guaranteed. The worst case examples are well known: the "growth factor" can be as large as \( O(2^n) \) (and can occur in practical applications [24]).

Theoretical studies on the numerical stability of GE have been made since 1940s by a great number of authors, for example, Turing [59], von Neumann and Goldstine [61], [62], Wilkinson [63], [64], and so on. Recently, Trefethen and Schreiber [57] considered the topic from a statistical viewpoint.
1.2 Krylov subspace methods

A Krylov subspace method is an iterative method for the solution of (1.1) when $A$ is sparse, or probably structured. At the $l$-th iteration of the method, the approximate solution $x_l$ is defined by the Petrov-Galerkin condition

$$x_l \in x_0 + K_l(A, r_0) \equiv x_0 + \text{span}\{r_0, Ar_0, \ldots, A^{l-1}r_0\}$$

and

$$b - Ax_l \perp L_l,$$

where $L_l$ is a subspace in $\mathbb{R}^n$ of dimension $l$ and where $r_0 = b - Ax_0$ is the initial residual for the initial guess $x_0$. The subspace $K_l(A, r_0)$ is called a Krylov subspace. The existing Krylov subspace methods in the literature are defined with special choices of $L_l$. For example, we obtain the BiCG method [22, 39, 50] when we choose

$$L_l = K_l(A^T, r_0^*),$$

where $r_0^*$ is an arbitrary vector, and the FOM method [47, 48, 50] when

$$L_l = K_l(A, r_0).$$

The classical Conjugate Gradient method (CG) [36] is a special version of BiCG where $A$ is symmetric.

Let $v$ be a vector and $\nu_{(A,v)}$ be the grade of $v$ with respect to $A$, i.e., the degree of the minimal polynomial of $v$ with respect to $A$. Given two vectors $v_0$ and $v_0^*$,
we define

\[ W_l = \begin{bmatrix}
  v_0^T v_0 & v_0^T A v_0 & \cdots & v_0^T A^{l-1} v_0 \\
  v_0^T A v_0 & v_0^T A^2 v_0 & \cdots & v_0^T A^l v_0 \\
  \vdots & \vdots & & \vdots \\
  v_0^T A^{l-1} v_0 & v_0^T A^l v_0 & \cdots & v_0^T A^{2(l-1)} v_0
\end{bmatrix}, \quad l = 1, 2, \cdots, \nu, \]

where \( \nu = \min\{\nu_{(A,v_0)}, \nu_{(A^T,v_0^*)}\} \). The derivation of BiCG is based on the Lanczos process \([38]\) which is figured out by the following theorem.

**Theorem 1.2** Given \( v_0 \) and \( v_0^* \). If the principal submatrices of \( W_\nu \) are nonsingular, there exist \( n \times \nu \) matrices \( V_\nu = [v_0, v_1, \cdots, v_{\nu-1}] \) and \( V_\nu^* = [v_0^*, v_1^*, \cdots, v_{\nu-1}^*] \) of rank \( \nu \), whose first columns are \( v_0 \) and \( v_0^* \) respectively, and a \( \nu \times \nu \) tridiagonal matrix \( T_\nu \) with all the entries \( t_{i+1,j} = 1 \) in its lower subdiagonal, such that

\[
(v_l, v_{l'}^*) = \delta_{l,l'}, \quad l, l' = 0, 1, \cdots, \nu - 1,
\]

where \( \delta_{l,l'} = 0 \) if \( l \neq l' \) and \( \delta_{l,l'} \neq 0 \) if \( l = l' \), and

\[
AV_\nu = V_\nu T_\nu
\]

and

\[
A^T V_\nu^* = V_\nu^* T_\nu.
\]

Such \( V_\nu, V_\nu^* \) and \( T_\nu \) are unique.

In practice, the Lanczos process could face breakdown - divisions by 0 - called *Lanczos breakdown*, when the Lanczos vectors \( v_l \) and \( v_l^* \) are exactly or nearly linear.
dependent. Techniques called look-ahead have been developed to cure this kind of breakdown, see, for instance, [10, 11, 26, 27, 32, 34, 44, 45, 56] and the references given therein. Moreover, in the actual derivation of the BiCG algorithm, the matrix $T_\nu$ is decomposed into two triangular matrices and as a result, BiCG is a two-term recurrence method. Since such a decomposition of $T_\nu$ need not exist, BiCG has another potential breakdown called pivot breakdown in addition to the Lanczos one. Several methods have been produced to remedy pivot breakdowns, e.g., CSBCG [6, 7] and QMR [29].

FOM is derived from the Arnoldi process [3] which is characterized in the theorem below.

**Theorem 1.3** Given $v_0$ with $\|v_0\|_2 = 1$. There exist an $n \times \nu$ orthonormal matrix $V_\nu = [v_0, v_1, \ldots, v_{\nu-1}]$ of rank $\nu$, whose first column is $v_0$, and a $\nu \times \nu$ Hessenberg matrix $H_\nu$ such that

$$AV_\nu = V_\nu H_\nu,$$

where $\nu = \nu_{(A,v_0)}$. Such $V_\nu$ and $H_\nu$ are unique.

Mathematically, no breakdown occurs in the Arnoldi process. Similarly to the case of BiCG, the matrix $H_\nu$ is factorized into two triangular matrices in deriving the FOM algorithm and therefore, pivot breakdown is possible in FOM. GMRES [47, 48, 50] is an improved version of FOM which not only remedies pivot breakdowns in FOM but minimizes the residual at each iteration. FOM and GMRES are long-term recurrence methods involving all previously computed
iterates and residuals, unless truncated or restarted, otherwise they are practically impossible. However, both truncated and restarted treatments lose the convergence history.

BiCG and QMR require to perform a matrix-by-vector product with the transpose matrix $A^T$ at each iteration. Since $A^T$ may not be available in some applications [50, p.214], methods avoiding transpose product have been developed. Examples of these include TFiQMR [14], CGS [55], BiCGSTAB [60], BiCGSTAB2 [33], TFQMR [25] and BiCGSTAB($k$) [53]. In principle, these methods compute a residual characterized by a product of the BiCG residual with a polynomial of the matrix $A$. Typical examples are CGS and BiCGSTAB. The residuals $r_{l}^{CGS}$ and $r_{l}^{BiCGSTAB}$ of CGS and BiCGSTAB respectively at the $l$-iteration are defined by

$$r_{l}^{CGS} = \psi_l(A)r_{l}^{BiCG}$$

and

$$r_{l}^{BiCGSTAB} = \phi_l(A)r_{l}^{BiCG},$$

where $\psi_l(\lambda)$ is the polynomial of degree $l$ such that $r_{l}^{BiCG} = \psi_l(A)r_0$ and where $\phi_l(\lambda) \equiv (\rho_l \lambda + 1)\phi_{l-1}(\lambda)$ and $\phi_0(\lambda) \equiv 1$. Transpose-free methods which cure breakdown have also been developed, e.g., [12, 30].
1.3 Contributions of this thesis

In this thesis, we study GE without pivoting from the point of view of probability (Chapter 2), how to preserve symmetry in preconditioned Krylov subspace methods (Chapter 3) and also develop a new Krylov subspace solver to the system (1.1) (Chapter 4).

1.3.1 Probabilistic analysis of GE without pivoting

Recently, Trefethen and Schreiber (TS) [57] studied the numerical stability of GE with pivoting from the viewpoint of statistics. Among their many results, they observed that for many distributions of matrices, the matrix elements after the first few steps of Gaussian elimination with (partial or complete) pivoting are approximately normally distributed. They also found that, for $n \leq 1024$, the average growth factor (normalized by the standard deviation of the initial matrix elements) is within a few percent of $n^{2/3}$ for the partial pivoting case and approximately $n^{1/2}$ for the complete pivoting case. After having performed more extensive experiments, Edelman and Mascarenhas [20] further suggest that the growth factor in the partial pivoting case may grow more like $n^{1/2}$ than $n^{2/3}$.

Following TS, we study the probability of small pivots and large growth factors in Chapter 2. However, we will only consider the case without pivoting. We are doing so for three reasons. The first is quite obvious: the non-pivoting case is far
easier to analyze than the pivoting case. In particular, we are able to derive in
closed form the density functions of the elements of the $LU$ factors and probabilistic
bounds for the occurrence of small pivots and the growth factors. The second reason
is that, with the advent of parallel computing, there is more incentive to trade off
the stability of partial pivoting for the higher performance of simpler but possibly
less stable forms of GE, including no pivoting, see, for instance, [42, 43]. Finally,
we are hoping that our results for GE without pivoting will be useful in the analysis
of, as well as providing a basis of comparison for, the partial pivoting case.

The class of random matrices we chose to study is that of (complex) Gaussian
matrices. By definition, a (complex) Gaussian matrix is a random matrix with
independent and identically distributed elements which are $(\tilde{N}(0,1)) N(0,1)$, the
normal distribution with mean 0 and variance 1. This choice is motivated by the
empirical results of TS mentioned earlier. Matrices of this type have also been
studied by Edelman [17], [19], who derived the expected singular values.

On some kinds of parallel computers, say, systolic systems [37], avoiding data
movements is necessary [40]. Based on our theoretical results and TS's observa-
tions, we suggest that we may use GE in such a way that, in the first few steps, we
use GE with partial (or complete) pivoting and then use the version without pivot-
ing (or with pairwise, neighbor pivoting). Our experimental results illustrate that
this idea may minimize data movements while avoiding breakdown and instability
of GE.
1.3.2 Symmetry in preconditioned Krylov subspace methods

Consider the solution of the linear system (1.1) by a preconditioned Krylov subspace method. Assume at first that $A$ is symmetric positive definite (SPD) and let $M$ be an SPD matrix that is a preconditioner for the matrix $A$. Then one possibility is to solve either the left-preconditioned system

$$M^{-1}Ax = M^{-1}b$$

or the right-preconditioned system

$$AM^{-1}u = b, \quad x = M^{-1}u.$$  \hspace{1cm} (1.4)

Both of these systems have lost their symmetry. A remedy exists when the preconditioner $M$ is available in factored form, e.g., as an incomplete Choleski factorization,

$$M = LL^T,$$

in which case a simple way to preserve symmetry is to split the preconditioner between left and right, i.e., solve

$$L^{-1}AL^{-T}u = L^{-1}b, \quad x = L^{-T}u.$$  \hspace{1cm} (1.4)

This can also be done when $M$ can be factored as $M = M^{1/2} \times M^{1/2}$. Unfortunately, the requirement that $M$ be available in factored form is often too stringent.
However, this remedy is not required. As is well-known, we can preserve symmetry by using a different inner product. Specifically, we observe that $M^{-1}A$ is self-adjoint for the $M$-inner product,

$$(x, y)_M = (Mx, y) = (x, My)$$

since we have

$$(M^{-1}Ax, y)_M = (Ax, y) = (x, Ay) = (x, M(M^{-1}A)y) = (x, M^{-1}Ay)_M.$$ 

Indeed, this is how the preconditioned conjugate gradient (PCG) and symmetrically preconditioned conjugate residual algorithms may be derived [31, 35, 50].

The question we now raise is the following. When $A$ is only nearly symmetric, it is often the case that there exists a preconditioner $M$ which is SPD. In this situation, the use of either of the forms (1.3) or (1.4) is just as unsatisfactory as in the fully symmetric case. Indeed, whatever degree of symmetry was available in $A$ is now entirely lost. Although the above remedy based on $M$-inner products is always used in the symmetric case, it is rather surprising that this problem is seldom ever mentioned in the literature for the nearly symmetric case. In the nonsymmetric case, when $M$ exists in factored form, some form of balancing can also be achieved by splitting the preconditioner from left and right. However, there does not seem to have been much work done in exploiting $M$-inner products even when $M$ is not available in factored form. This dichotomy between the treatment of the symmetric and the nonsymmetric cases is what motivated the study in Chapter
3, where we suppose $A$ is nearly symmetric and show how alternative inner products may be used to preserve the degree of symmetry of $A$ in preconditioned Krylov subspace methods.

Ashby, Manteuffel, and Saylor [4] have fully considered the case of using alternate inner products when the matrix $A$ is symmetric. Work of which we are aware that consider the use of alternate inner products when $A$ is near-symmetric are Young and Jea [65] and Meyer [41]. In the latter, the $A^TM^{-1}A$ inner product (with $M$ SPD) is used with ORTHOMIN and ORTHODIR.

1.3.3 ML($k$)BiCGSTAB

BiCGSTAB is a popular Krylov subspace method for the iterative solution of nonsymmetric linear systems. Its main features are that it is transpose-free, makes more efficient use of matrix-vector products when compared to BiCG and is more stable than CGS, but it faces potential breakdown. In Chapter 4, we introduce a new variant of BiCGSTAB which inherits all of these nice features and reduces the probability of breakdown to zero. In addition, the key new ingredient of our method is the use of multiple starting left Lanczos vectors which has the desirable effect of lowering the cost per step and increasing the robustness.

BiCGSTAB is derived from BiCG which is a Lanczos based Krylov subspace method. In BiCG, the residual vector $r_l$ at the $l$-th step lies in a Krylov subspace
$K_{l+1}(A, r_0)$ and is chosen to be

$$r_l \perp L_l = K_l(A^T, r_0^*) .$$

In our variant of the BiCG method, which we denote by ML($k$)BiCG, $r_l$ is still in $K_{l+1}(r_0, A)$ but is now chosen to be

$$r_l \perp L_l = \text{span}\{p_1, p_2, \cdots, p_l\},$$

where $p_{jk+i} \equiv A^T j q_i, i = 1, 2, \cdots, k; j = 0, 1, 2, \cdots,$ and $q_i$'s are $k$ arbitrary vectors. In the context of Krylov subspaces, $L_l$ is the union of $k$ Krylov subspaces $K_{j+1}(A^T, q_s)$ and $K_j(A^T, q_{s'})$, where $1 \leq s \leq i < s' \leq k$ and $jk + i = l$, generated from $k$ linearly independent starting vectors $q_i$.

Our motivation for using multiple starting Lanczos vectors is to mitigate somewhat the ill-conditioning of $K_l(A^T, r_0^*)$ for large $l$ by replacing a high degree Krylov polynomial corresponding to one starting vector by a union of lower degree Krylov polynomials generated from different, independent starting vectors. We think this leads to better stability and robustness of the resulting iterative method and, by choosing the starting vectors randomly, we may avoid breakdown in practice. We derive an efficient implementation of this idea, requiring only memory of the previous $k$ iterates (i.e. a $k$-term recurrence).

But we consider the major contribution in Chapter 4 to be an extension of BiCGSTAB, which we denote by ML($k$)BiCGSTAB, using multiple starting Lanczos vectors. The derivation is similar to, but rather more complicated than,
that of deriving BiCGSTAB from BiCG. Not only does our method inherits from ML(k)BiCG the increased stability of using multiple starting Lanczos vectors, it also inherits from BiCGSTAB the advantages of being transpose-free and also more efficient in using matvec per step than ML(k)BiCG. Specifically, after \( l = jk + i \) steps, where \( i = 1, \ldots, k \) and \( j = 0, 1, \ldots \), the residual vector \( \tilde{r}_l \) can be written as
\[
\tilde{r}_l = \psi_{j+1}(A)\phi_l(A)r_0,
\]
where \( \phi_l \) is the degree \( l \) polynomial corresponding to the residual vector \( r_l \) of ML(k)BiCG, and \( \psi_{j+1} \) is a degree \( j + 1 \) smoothing polynomial. Thus, the “degree” of \( \tilde{r}_l \) is \( l + j + 1 \). To compute \( \tilde{r}_l \), exactly \( l + j + 1 \) matvec’s with \( A \) (no \( A^T \)) are required. Thus, the cost per step on the average is \( 1 + 1/k \) matvec’s.

Both ML(k)BiCG and ML(k)BiCGSTAB can be implemented efficiently as \( k \)-term recurrences.

A way to view our method is through the conditioning of the collection of vectors to which the residual vector is required to be orthogonal, which we believe controls the stability of the method. For FOM/GMRES, these vectors are mutually orthogonal and thus is perfectly conditioned. For BiCG, these vectors are the Lanczos vectors and they can be ill-conditioned. For our new ML(k)BiCG, these vectors consist of a union of \( k \) sets of Lanczos vectors, generated by (arbitrary) initial vectors which can be made orthogonal, thus making them better conditioned than in the BiCG case. Our new ML(k)BiCGSTAB method, being a product method derived from ML(k)BiCG, inherits this increased stability. Thus, ML(k)BiCGSTAB can be viewed as an attempt to merge the advantages of
FOM/GMRES (stability) and BiCGSTAB (short recurrence, transpose-free and efficient use of matvec's) while avoiding their disadvantages: FOM/GMRES (long recurrence) and BiCGSTAB (imperfect stability).

Our ML(k)BiCGSTAB method has a close relationship to the Lanczos-type method for multiple starting vectors proposed recently by Aliaga, Boley, Freund and Hernández (ABFH) [1]. In fact, even though we have developed our methods independently of the ABFH framework, our BiCG extension ML(k)BiCG can be easily derived from the ABFH framework, using one right Lanczos vector and k left Lanczos vectors. Even so, ML(k)BiCG deserves some interest on its own right, because we believe it is the first attempt in using multiple starting vector Lanczos-type methods for solving a single linear system.\footnote{Freund and Malhotra [28] did consider a QMR-type method based on the ideas in [1] for linear systems with multiple right-hand-sides, but there the number of right and left Lanczos vectors are the same.} The main application of nonequal left and right starting vectors cited in [1] is for computing transfer functions in multi-input multi-output time invariant linear dynamical systems. A more significant difference between our present results and those in [1] is that we have derived a BiCGSTAB variant based on multiple starting vectors, with the advantages stated earlier. We believe this is the first product Krylov method based on multiple starting vectors.

The origin of the ideas behind the new methods presented here can be traced to Ruhe's vectorwise implementation of the block Lanczos method [46], in the same way that [1] can be considered an extension of Ruhe's method to non-Hermitian
matrices and non-equal left and right starting vectors plus look-ahead. As such, we believe our methods inherit the advantage of faster convergence of the underlying block Lanczos method.

Gutknecht [33], Sleijpen and Fokkema [53] recently generalized BiCGSTAB to versions called BiCGSTAB2 and BiCGSTAB(\(k\)) respectively, in which they replaced the GMRES(1) part in BiCGSTAB with GMRES(\(k\)). The purpose of doing so is to increase the robustness of BiCGSTAB. From our experimental results, ML(\(k\))BiCGSTAB appears to be a good alternative to achieve this goal exploiting the robustness of GMRES(\(k\))/FOM(\(k\)).
CHAPTER 2

Probabilistic Analysis of Gaussian Elimination Without Pivoting and Its Application

Let $X$ be an $n \times n$ square (complex) Gaussian matrix and let $X = LU$ be the $LU$ factorization of $X$. We derive the density functions of the entries of $L$ and $U$ respectively and prove that the probability of the occurrence of a pivot less than $\varepsilon$ in magnitude is $O(\varepsilon)$.\footnote{We note that Foster [23] has studied the probability of large diagonal elements in the QR factorization of a rectangular matrix $A$.} We also derive bounds on the probabilities of large growth factors. In particular, we prove that the probabilities $\text{Prob}(\|U\|_{\infty}/\|A\|_{\infty} > n^{2.5})$ and $\text{Prob}(\|L\|_{\infty} > n^3)$ decay algebraically to zero as $n$ tends to infinity. From our experimental results, we observe that the probabilities $\text{Prob}(n \leq \|U\|_{\infty} < n^{1.5})$ and $\text{Prob}(n \leq \|U\|_{\infty}/\|A\|_{\infty} < n^{1.5})$ tend to one as $n$ goes to infinity. This indicates that our theoretical bounds are not the tightest possible but not too loose either.

2.1 Properties of a Gaussian Matrix

We present some properties involving Gaussian matrices which will be used in the following sections. These properties can be derived easily from classical facts...
that may be found in standard texts or papers, for example, \([2, 9, 18, 52, 58]\).

Let \(N(0,1)\) refer to the normal distribution with mean 0 and variance 1 and let \(\tilde{N}(0,1)\) the complex distribution of \(x + yi\) where \(x\) and \(y\) are independent and identically distributed (iid) \(N(0,1)\). By definition, an (complex) Gaussian matrix is a random matrix with iid elements which are \((\tilde{N}(0,1)) N(0,1)\).

**Proposition 2.1** Let \(v\) and \(X\) be two independent (complex) Gaussian matrices of sizes \(n \times 1\) and \(n \times n\) respectively. Let \(H\) be the Householder matrix such that

\[
Hv = (s, 0, \cdots, 0)^T
\]

where \(s \geq 0\) and let

\[
Y = X(PH)^T,
\]

where \(P\) is a permutation matrix such that \(PHv = (0, \cdots, 0, s)^T\). Then the entries \(s, y_{ij}, i, j = 1, \cdots, n\), are independent and \(s^2\) is \((\chi^2_{2n}) \chi^2_n\) while all \(y_{ij}\) are \((\tilde{N}(0,1)) N(0,1)\). Moreover, Let

\[
X = QR
\]

be the QR factorization of \(X\) obtained by performing the standard Householder transformations and let

\[
w = Q^{-1}v.
\]

Then the entries \(r_{ij}, w_i, i = 1, \cdots, n, j = i, \cdots, n\) are independent and \(r_{ii}\) is

\[(\chi^2_{2(n-i+1)}) \chi^2_{n-i+1}, i = 1, \cdots, n\] while all others are \((\tilde{N}(0,1)) N(0,1)\).
2.2 Real Gaussian Elimination

In this section, we suppose all the matrices appearing have real entries.

2.2.1 Density Functions of $u_{pq}$ and $l_{pq}$

Let $X$ be an $n \times n$ Gaussian matrix and let $X = LU$, where $L$ is an unit lower triangular matrix and $U$ is an upper triangular matrix, be the $LU$ factorization of $X$. The $(p,q)$-th ($p \leq q$) entry $u_{pq}$ of $U$ and the entries of $X$ have the following relation.

Lemma 2.1 Let $X = LU$ be the $LU$ factorization of $X$. Then

$$u_{pq} = x_{pq} - x_{p*}^T X_{p-1}^{-1} x_{*q},$$

where

$$x_{p*} = (x_{p1}, \cdots, x_{pp-1})^T,$$

$$x_{*q} = (x_{1q}, \cdots, x_{p-1q})^T,$$

and $X_{p-1}$ is the $(p - 1) \times (p - 1)$ leading principal submatrix of $X$.

Proof. Permuting the $p$-th and $q$-th columns of $X$ and $U$ simultaneously on both sides of $X = LU$ and then comparing the corresponding blocks, we find

$$\begin{bmatrix} X_{p-1} & x_{*q} \\ x_{p*}^T & x_{pq} \end{bmatrix} = \begin{bmatrix} L_{p-1} & 0 \\ I_{p*}^T & 1 \end{bmatrix} \begin{bmatrix} U_{p-1} & u_{*q} \\ 0 & u_{pq} \end{bmatrix}$$

\(^2\)Since they just form a set of measure zero, we ignore matrices for which the Gaussian elimination fails.
where
\[ l_{p*} = (l_{p1}, \cdots, l_{pp-1})^T, \]
\[ u_{*q} = (u_{1q}, \cdots, u_{p-1q})^T \]
and where \( L_{p-1} \) and \( U_{p-1} \) are the \((p-1) \times (p-1)\) leading principal submatrices of \( L \) and \( U \) respectively. It follows that

\[ X_{p-1} = L_{p-1} U_{p-1}, \quad I_{p*} = x_{p*}^T U_{p-1}^{-1}, \]

\[ u_{*q} = l_{p-1}^{-1} x_{*q}, \quad u_{pq} = x_{pq} - I_{p*}^T u_{*q} \]

and these imply the desired equation. \( \square \)

Let \( H \) be an \((p-1) \times (p-1)\) orthogonal matrix, e.g., the one \( PH \) in Proposition 2.1, such that

\[ x_{p*}^T H = (0, \cdots, 0, s) = \eta^T \]

with \( s \geq 0 \). Then

\[ u_{pq} = x_{pq} - \eta^T (X_{p-1}H)^{-1} x_{*q} \]

\[ \equiv x_{pq} - \eta^T Y^{-1} x_{*q}. \]

By Proposition 2.1, the entries \( s, x_{pq}, x_{iq} \) and \( y_{ij}, i, j = 1, \cdots, p-1 \), are mutually independent and all \( x_{pq}, x_{iq} \) and \( y_{ij}, i, j = 1, \cdots, p-1 \), are \( N(0,1) \) while \( s^2 \) is \( \chi^2_{p-1} \). We now decompose \( Y \) as

\[ Y = QR, \]
where \( Q \) is an \((p-1) \times (p-1)\) orthogonal matrix and \( R \) an \((p-1) \times (p-1)\) upper triangular matrix with positive diagonal elements. We then further have

\[
\begin{align*}
    u_{pq} &= x_{pq} - \eta^T R^{-1} Q^T x_{sq} \\
    &= x_{pq} - \eta^T R^{-1} w \\
    &= x_{pq} - \frac{sw_{p-1}}{r_{p-1p-1}}.
\end{align*}
\]  

Again, by Proposition 2.1, the variables \( s, x_{pq}, w_i \) and \( r_{ij}, i \leq j, i, j = 1, \ldots, p-1 \), are independent. \( s^2 \) is \( \chi^2_{p-1} \) and \( r^2_{ii} \) is \( \chi^2_{p-1} \), \( i = 1, \ldots, p - 1 \) and all others are \( N(0,1) \).

Set

\[
v = \frac{s^2}{(p-1)r_{p-1p-1}^2}
\]

and

\[
z = \frac{1}{\sqrt{1 + (p-1)v}} x_{pq} - \frac{\sqrt{(p-1)v}}{\sqrt{1 + (p-1)v}} w_{p-1},
\]

then \( v \) is \( F_{p-1,1} \), the \( F \) distribution with \( p - 1 \) and \( 1 \) degrees of freedom, and \( z \) is \( N(0,1) \) and (2.1) becomes

\[
    u_{pq} = z\sqrt{1 + (p-1)v}.
\]  

Since the variables \( v \) and \( z \) are independent and their density functions are known, it is straightforward to determine the density function of \( u_{pq} \).
Theorem 2.1 Suppose $X$ is an $n \times n$ Gaussian matrix and let $X = LU$ be the LU factorization of $X$. Then the density function of the $(p,q)$-th entry of $U$ is

$$f_{u_{pq}}(t) = \frac{\sqrt{2}}{\pi} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} \left( \sum_{i=0}^{\left\lfloor p/2 \right\rfloor} \xi_{i,p} t^{-2i-2} + (-1)^{\left\lfloor p/2 \right\rfloor} \zeta_{p} t^{-p+1} \exp \left( -\frac{1}{2} t^2 \right) \phi_{p}(t) \right)$$

(2.3)

where

$$\xi_{i,p} = \begin{cases} (-1)^i \prod_{j=0}^{i-1} (p-2j-3) & i > 0 \\ 1 & i = 0 \end{cases}$$

$$\zeta_{p} = \begin{cases} (p-3)!! & p > 3 \\ 1 & p = 2, 3 \end{cases}$$

$$\phi_{p}(t) = \left( \int_{0}^{t} \exp \left( \frac{1}{2} x^2 \right) dx \right)^{p-1-2\left\lfloor (p-1)/2 \right\rfloor}$$

and where $-\infty < t < \infty$, $2 \leq p \leq q$.

Proof. Since the variables $v$ and $z$ in (2.2) are $F_{p-1,1}$ and $N(0,1)$ respectively, the density functions of them are given as follows,

$$f_{u}(t) = \begin{cases} \frac{1}{\sqrt{\pi}} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} (p-1)^{(p-1)/2} t^{(p-3)/2} (1 + (p-1)t)^{-p/2} & t > 0 \\ 0 & t \leq 0 \end{cases}$$

24
and

\[ f_z(t) = \frac{1}{\sqrt{2\pi}} \exp(-t^2/2). \]

Since \( v \) and \( z \) are independent, their joint density function is given by

\[ f(v, z) = f_v(v) f_z(z) \]

\[ = \begin{cases} 
    c v^{(p-3)/2} (1 + (p-1)v)^{-p/2} \exp(-z^2/2) & v > 0 \\
    0 & \text{otherwise},
\end{cases} \]

where \( c = \frac{1}{\sqrt{2\pi}} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} (p-1)^{(p-1)/2} \). Thus, the distribution function \( F_{u_{pq}}(\alpha) \) of \( u_{pq} \) is

\[ F_{u_{pq}}(\alpha) = c \int \int_{u_{pq} \leq \alpha} v^{(p-3)/2} (1 + (p-1)v)^{-p/2} \exp(-z^2/2) \, dv \, dz \]

\[ = c \int_{\sqrt{1+(p-1)v} \leq \alpha} v^{(p-3)/2} (1 + (p-1)v)^{-p/2} \exp(-z^2/2) \, dv \, dz \]

\[ = c \int_0^\infty dv \int_{-\infty}^{\sqrt{1+(p-1)v}} v^{(p-3)/2} (1 + (p-1)v)^{-p/2} \exp(-z^2/2) \, dz. \]

Letting \( z = t/\sqrt{1+(p-1)v} \) we find

\[ F_{u_{pq}}(\alpha) = c \int_{-\infty}^\alpha dt \int_0^\infty v^{(p-3)/2} (1 + (p-1)v)^{-(p+1)/2} \exp\left(\frac{1}{2} \frac{t^2}{1 + (p-1)v}\right) \, dv. \]

Letting \( s = 1/(1 + (p-1)v) \) this can then be written

\[ F_{u_{pq}}(\alpha) = \frac{1}{\sqrt{2\pi}} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} \int_{-\infty}^\alpha dt \int_0^1 (1-s)^{(p-3)/2} \exp\left(-\frac{1}{2} t^2 s\right) \, ds. \]
Thus
\[ f_{u_{pq}}(t) = \frac{1}{\sqrt{2\pi}} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} \int_0^1 (1 - s)^{(p-3)/2} \exp\left(-\frac{1}{2} t^2 s\right) ds. \]

Finally, letting \( w = |t|\sqrt{1-s} \) we have
\[ f_{u_{pq}}(t) = \frac{\sqrt{2}}{\pi} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} \exp\left(-\frac{1}{2} t^2 \right) t^{p+1} \int_0^t w^{p-2} \exp\left(\frac{1}{2} w^2 \right) dw. \quad (2.4) \]

Since
\[
\int_0^t w^{p-2} \exp\left(\frac{1}{2} w^2 \right) dw = \exp\left(\frac{1}{2} t^2 \right) \sum_{i=0}^{\left\lfloor \frac{p+1}{2} \right\rfloor} (-1)^i \prod_{j=0}^{i-1} (p - 2j - 3) t^{p-2i-3} + \\
(-1)^{\left\lfloor (p-1)/2 \right\rfloor} (p-3)!! \left( \int_0^t \exp\left(\frac{1}{2} x^2 \right) dx \right)^{p-1-2\left\lfloor (p-1)/2 \right\rfloor}
\]
by integration by parts, where we define 0!! = (-1)!! = 1, the desired result follows.

\( \Box \)

Similar to the derivation of the density function of \( u_{pq} \), we first establish a relation between \( l_{pq} \) and the entries of \( X \) and then simplify it. Let \( X = LU \) and \( X^T = \tilde{L}\tilde{U} \) be the \( LU \) factorizations of \( X \) and \( X^T \) respectively. Set \( \tilde{D} = \text{diag}(\tilde{u}_{11}, \ldots, \tilde{u}_{nn}) \). Thus, \( X^T = \tilde{L}\tilde{D}\tilde{D}^{-1}\tilde{U} \). So \( X = (\tilde{D}^{-1}\tilde{U})^T (\tilde{L}\tilde{D})^T \). Note that
\[
(\tilde{D}^{-1}\tilde{U})^T \text{ is unit lower triangular and } (\tilde{L}\tilde{D})^T \text{ upper triangular. By the uniqueness of the } LU \text{ factorization of } X, \text{ we have}
\]
\[
L = (\tilde{D}^{-1}\tilde{U})^T.
\]

Hence
\[
l_{pq} = \tilde{u}_{aq}/\tilde{u}_{qq}
\]
for $1 \leq q < p \leq n$. By Lemma 2.1,

$$
\tilde{u}_{qp} = x_{pq} - x_{*q}^T X_{q-1}^{-T} x_{p*}
$$

and

$$
\tilde{u}_{qq} = x_{qq} - x_{*q}^T X_{q-1}^{-T} x_{q*},
$$

where

$$
x_{p*} = (x_{p1}, \ldots, x_{pq-1})^T,
$$

$$
x_{q*} = (x_{q1}, \ldots, x_{qq-1})^T,
$$

$$
x_{*q} = (x_{1q}, \ldots, x_{q-1q})^T
$$

and $X_{q-1}$ is the $(q-1) \times (q-1)$ leading principal submatrix of $X$. We now let $H$ be an $(q-1) \times (q-1)$ orthogonal matrix such that

$$
x_{*q}^T H = (0, \ldots, 0, s) \equiv \eta^T
$$

with $s \geq 0$. Then

$$
l_{pq} = \frac{x_{pq} - \eta^T (X_{q-1}^T H)^{-1} x_{p*}}{x_{qq} - \eta^T (X_{q-1}^T H)^{-1} x_{q*}}
$$

$$
= \frac{x_{pq} - \eta^T Y^{-1} x_{p*}}{x_{qq} - \eta^T Y^{-1} x_{q*}}.
$$

As in the case of $u_{pq}$, all the entries in the above expression are mutually independent and $s^2$ is $\chi^2_{q-1}$ while others are $N(0,1)$. Let

$$
Y = QR
$$

27
be the QR factorization of $Y$ where $R$ has positive diagonal elements. Then the expression can be reduced to

$$ l_{pq} = \frac{x_{pq} - \eta^T R^{-1} Q^T x_{p^*}}{x_{qq} - \eta^T R^{-1} Q^T x_{q^*}} $$

$$ = \frac{x_{pq} - \eta^T R^{-1} w}{x_{qq} - \eta^T R^{-1} \mu} $$

$$ = \frac{r_{q-1q-1} x_{pq} - s w_{q-1}}{r_{q-1q-1} x_{qq} - s \mu_{q-1}}. $$

The entries $x_{pq}, x_{qq}, w_i, \mu_i$ and $r_{ij} (i < j)$ are $N(0,1)$ while $s^2$ is $\chi^2_{q-1}$ and $r_{ii}^2$ is $\chi^2_{q-2}$, where $i = 1, \ldots, q - 1, j = 2, \ldots, q - 1$. They are all independent.

Set

$$ v = \frac{r_{q-1q-1}}{\sqrt{r_{q-1q-1}^2 + s^2}} x_{pq} - \frac{s}{\sqrt{r_{q-1q-1}^2 + s^2}} w_{q-1} $$

and

$$ z = \frac{r_{q-1q-1}}{\sqrt{r_{q-1q-1}^2 + s^2}} x_{qq} - \frac{s}{\sqrt{r_{q-1q-1}^2 + s^2}} \mu_{q-1}, $$

then $l_{pq} = v/z$. Since both $v$ and $z$ are $N(0,1)$ and they are independent, $l_{pq}$ has Cauchy distribution.

**Theorem 2.2** Suppose $X$ is an $n \times n$ Gaussian matrix and let $X = LU$ be the LU factorization of $X$. Then the density function of the $(p,q)$-th entry of $L$ is

$$ f_{l_{pq}}(t) = \frac{1}{\pi} \frac{1}{1 + t^2} $$

where $-\infty < t < \infty$ and $1 \leq q < p \leq n$.  

28
2.2.2 Probability of Small Pivot

In practice, if one of the pivot elements $u_{pp}$ is zero or smaller in magnitude than a preset tolerance $\epsilon$, Gaussian elimination will fail. In this section, we describe the probability of the occurrence of such a situation.

To make the statements below neatly, we use a shorthand notation here. For given $\epsilon > 0$ and $1 \leq p \leq n$, we define

$$E_{p,\epsilon} = \{ X \in R^{n \times n} | |u_{pp}| < \epsilon \}.$$ 

Then the event that at least one $u_{pp}$ has $|u_{pp}| < \epsilon$ is naturally denoted by $\bigcup_{p=1}^{n} E_{p,\epsilon}$.

**Lemma 2.2** Suppose $X$ is an $n \times n$ Gaussian matrix and let $X = LU$ be the LU factorization of $X$. Given $\epsilon > 0$ and $1 \leq p \leq n$. Then

$$\text{Prob}(E_{p,\epsilon}) \leq \frac{\sqrt{2}}{\pi} \frac{\Gamma(p/2)}{\Gamma((p+1)/2)} \epsilon.$$ 

**Proof.** From (2.4), we have

$$f_{u_{pp}}(t) \leq \frac{\sqrt{2}}{\pi} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} t^{p+1} \int_{0}^{t} w^{p-2} dw = \frac{1}{\sqrt{2} \pi} \frac{\Gamma(p/2)}{\Gamma((p+1)/2)},$$

where $p \geq 2$, and from which the desired result follows.

For the case where $p = 1$, it is sufficient to note that

$$\text{Prob}(E_{1,\epsilon}) = \text{Prob}(|x_{11}| < \epsilon) = \frac{1}{\sqrt{2\pi}} \int_{-\epsilon}^{\epsilon} \exp\left(-\frac{1}{2}t^2\right) dt.$$
Theorem 2.3 Suppose $X$ is an $n \times n$ Gaussian matrix and let $X = LU$ be the

LU factorization of $X$. Then

$$Prob(\bigcup_{p=1}^{n} E_{p, \epsilon}) \leq \frac{\sqrt{2}}{\pi} \sum_{p=1}^{n} \frac{\Gamma(p/2)}{\Gamma((p+1)/2)} \epsilon. \quad (2.7)$$

Proof. Since $Prob(\bigcup_{p=1}^{n} E_{p, \epsilon}) \leq \sum_{p=1}^{n} Prob(E_{p, \epsilon})$, (2.7) follows by Lemma 2.2. \qed

The coefficient of $\epsilon$ is a rather slow-growing function of $n$. In fact, it is about

1800 even when $n = 10^6$. So, if $\epsilon$ is small enough, (2.7) will certainly give a

satisfying bound for the desirable probability. Moreover, the right hand side of

(2.7) is linear with $\epsilon$.

2.2.3 Probability of Large Growth Factor

It is possible that the growth factors $\rho_L$ and $\rho_U$, which are defined by (1.2)

in §1.1 of Chapter 1, can be very large because small pivots can appear. In this

section, we study the probabilities of the occurrence of large $\rho_L$ and $\rho_U$ and give

probabilistic bounds on the sizes of them in the following Theorem.

Theorem 2.4 Suppose $X$ is an $n \times n$ Gaussian matrix and let $X = LU$ be the LU

factorization of $X$. Then there exist numbers $1 > b > 0$ and $c > 0$, independent of

$n$, such that

$$Prob(\rho_U > r) \leq \frac{c}{r} n^{5/2} + \min \left( \frac{c}{r} n^{7/2}, \frac{1}{n} \right) + \frac{b^n}{n}$$

30
and

$$\text{Prob}(\rho_L > r) \leq \frac{c}{r} n^3$$

for any $r \geq 1$.

**Proof.** We first claim that there exists a $c_1 > 0$, independent of $n$, such that

$$\text{Prob}(|U|_\infty > r) \leq \frac{c_1}{r} n^{7/2}. \quad (2.8)$$

In fact, by (2.4), we have

$$f_{\text{up}}(t) \leq \frac{\sqrt{2}}{\pi} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} \exp \left(-\frac{1}{2} t^2\right) t^{-1} \int_0^t \exp \left(\frac{1}{2} w^2\right) dw.$$ 

$$= \frac{\sqrt{2}}{\pi \sqrt{\bar{p}}} \frac{\Gamma(p/2)}{\Gamma((p-1)/2)} \exp \left(-\frac{1}{2} t^2\right) \int_0^t \exp \left(\frac{1}{2} w^2\right) dw \frac{\sqrt{\bar{p}}}{t^2}.$$ 

Since

$$\lim_{k \to +\infty} \frac{1}{\sqrt{k}} \frac{\Gamma(k/2)}{\Gamma((k-1)/2)}$$

exists by Stirling’s formula

$$\lim_{x \to +\infty} \frac{\Gamma(x+1)}{x^x \exp(-x) \sqrt{2\pi x}} = 1,$$

and since

$$\lim_{t \to +\infty} t \exp \left(-\frac{1}{2} t^2\right) \int_0^t \exp \left(\frac{1}{2} w^2\right) dw = 1,$$

we can find a $c_2$ such that

$$f_{\text{up}}(t) \leq c_2 \sqrt{\bar{p}/t^2}.$$
Therefore
\[ P_{\text{prob}}(\|U\|_\infty > r) \leq \sum_{p=1}^{n} \sum_{q=p}^{n} P(|u_{pq}| > r/n) \]
\[ = \sum_{p=1}^{n} \sum_{q=p}^{n} \int_{|t| > r/n} f_{pq}(t) \, dt \]
\[ \leq \sum_{p=1}^{n} \sum_{q=p}^{n} \int_{|t| > r/n} \frac{c_2 \sqrt{p}}{t^2} \, dt \]
\[ \leq \frac{c_2 c_3}{r} n^{7/2} \]
for some \( c_3 > 0 \), independent of \( n \). The existence of \( c_3 \) is due to the existence of the limit
\[ \lim_{k \to +\infty} \frac{1}{k^{5/2}} \sum_{p=1}^{k} (k - p + 1) \sqrt{p} = \int_{0}^{1} (1 - t) \sqrt{t} \, dt. \]

We set \( c_1 = c_2 c_3 \) and then (2.8) is proven. For proving the first inequality in the theorem, we note that the expected value \( \mu \) and the variance \( \sigma^2 \) of the variable \( x_1 = \sum_{q=1}^{n} |x_{1q}| \) are
\[ \mu = n \sqrt{\frac{2}{\pi}} , \quad \sigma^2 = \left( 1 - \frac{2}{\pi} \right) n. \]

Setting \( \epsilon = n \sqrt{1 - \frac{2}{\pi}} \) in Chebyshev's inequality [15, p.183]
\[ \text{Prob}(|x_1 - \mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2} , \]
we have
\[ \text{Prob}(x_1 < nc_4) \leq \frac{1}{n} , \quad (2.9) \]
where $c_4 = \sqrt{\frac{2}{\pi}} - \sqrt{1 - \frac{2}{\pi}}$. Combining (2.8) and (2.9) we find

$$
\text{Prob}(\rho_U > r) = \text{Prob} (\|U\|_\infty > r\|A\|_\infty)
$$

$$
\leq \text{Prob} (\|U\|_\infty > rx_1)
$$

$$
= \text{Prob} (\|U\|_\infty > rx_1, x_1 \geq nc_4) + \text{Prob} (\|U\|_\infty > rx_1, nc_4 > x_1 > 1) + 
$$

$$
\text{Prob} (\|U\|_\infty > rx_1, x_1 \leq 1)
$$

$$
\leq \text{Prob} (\|U\|_\infty > nrc_4) + \min (\text{Prob} (\|U\|_\infty > r), \text{Prob} (x_1 < nc_4)) + 
$$

$$
\text{Prob} \left( |x_{1q}| \leq 1, \forall 1 \leq q \leq n \right)
$$

$$
\leq \frac{c_1}{c_4 r^n} + \min \left( \frac{c_1 n^{7/2}}{r n^{7/2}}, \frac{1}{n} \right) + \left( \frac{1}{\sqrt{2\pi}} \int_{-1}^{1} \exp \left( -\frac{1}{2} t^2 \right) dt \right)^n.
$$

Finally,

$$
\text{Prob}(\rho_L > r) \leq \sum_{p=2}^{n} \sum_{q=1}^{p-1} \text{Prob} \left( |l_{pq}| > \frac{r-1}{n-1} \right)
$$

$$
= \frac{1}{\pi} \sum_{p=2}^{n} \sum_{q=1}^{p-1} \int_{|t| > \frac{r-1}{n-1}} \frac{1}{1 + t^2} dt
$$

$$
\leq \frac{c_5}{n^3} \frac{1}{r^3}
$$

for some $c_5$. \qed
2.2.4 Numerical Experiments

In this subsection, we present numerical results to support Theorems 2.1 - 2.4. All our calculations have been carried out in MATLAB 4.2c on SUN workstations.

In our first experiment, 595000 Gaussian matrices of dimension \( n = 31 \) were selected at random. Then Gaussian elimination was applied to each of the matrices and then statistics on the elements \( l_{13,12}, l_{30,29}, u_{12,12} \) and \( u_{31,31} \) were accumulated. The data are plotted in Figures 2.1 and 2.2 together with the corresponding functions indicated in Theorems 2.1 and 2.2. In order to make clearer the difference between Figures 2.1(a) and 2.1(b), we present them together in Figure 2.3(a).

The purpose of our second experiment is to test formula (2.7). Gaussian matrices of several dimensions \( n \) were selected at random, with the sample size varying. A few tolerances \( \epsilon \) were used. The results are outlined in Table 2.1. The frequency column of the table provides the numbers of matrices which, in their LU factors, have at least one \( u_{pp} \) less than \( \epsilon \) in magnitude. By comparing with the empirical probabilities, we conclude that the bound given in (2.7) is a fairly tight one.

Finally, if we set \( r = n^\alpha, \alpha > 2.5 \) for \( \rho_U \) and \( \alpha > 3 \) for \( \rho_L \), in Theorem 2.4, then we can see that the probabilities \( \text{Prob}(\rho_L > n^\alpha) \) and \( \text{Prob}(\rho_U > n^\alpha) \) decrease with \( n \) increasing. In fact, empirically this is true even for smaller \( \alpha \), say, \( \alpha > 1.5 \) for both \( \rho_L \) and \( \rho_U \), as illustrated in Figures 2.3(b) and 2.4. In this experiment, we chose sample sizes to be 968500, 365500 and 98000 for \( n = 25, 50 \) and 100
<table>
<thead>
<tr>
<th>$n$</th>
<th>$\epsilon$</th>
<th>Sample Size</th>
<th>Frequency</th>
<th>Empirical probability</th>
<th>Theoretical bound</th>
</tr>
</thead>
<tbody>
<tr>
<td>25</td>
<td>$10^{-5}$</td>
<td>$10^6$</td>
<td>5</td>
<td>$5 \times 10^{-5}$</td>
<td>$5.8586 \times 10^{-5}$</td>
</tr>
<tr>
<td>50</td>
<td>$10^{-3}$</td>
<td>$10^4$</td>
<td>90</td>
<td>0.009</td>
<td>0.0085</td>
</tr>
<tr>
<td>50</td>
<td>$10^{-4}$</td>
<td>$10^4$</td>
<td>8</td>
<td>$8 \times 10^{-4}$</td>
<td>$8.4853 \times 10^{-4}$</td>
</tr>
<tr>
<td>50</td>
<td>$10^{-5}$</td>
<td>$10^4$</td>
<td>0</td>
<td>0</td>
<td>$8.4852 \times 10^{-5}$</td>
</tr>
<tr>
<td>50</td>
<td>$10^{-5}$</td>
<td>$10^5$</td>
<td>9</td>
<td>$9 \times 10^{-5}$</td>
<td>$8.4852 \times 10^{-5}$</td>
</tr>
<tr>
<td>75</td>
<td>$10^{-3}$</td>
<td>$10^4$</td>
<td>89</td>
<td>0.0089</td>
<td>0.0105</td>
</tr>
<tr>
<td>75</td>
<td>$10^{-4}$</td>
<td>$10^4$</td>
<td>8</td>
<td>$8 \times 10^{-4}$</td>
<td>0.0011</td>
</tr>
<tr>
<td>75</td>
<td>$10^{-5}$</td>
<td>$10^4$</td>
<td>0</td>
<td>0</td>
<td>$1.0519 \times 10^{-4}$</td>
</tr>
<tr>
<td>100</td>
<td>$10^{-3}$</td>
<td>$10^4$</td>
<td>115</td>
<td>0.0115</td>
<td>0.0122</td>
</tr>
</tbody>
</table>

Table 2.1: Probabilities of small pivot.

Figure 2.1: (a) Distribution of $u_{12,12}$: observed(+), predicted(-). (b) Distribution of $u_{31,31}$: observed(+), predicted(-).
Figure 2.2: (a) Distribution of $l_{13,12}$: observed(+) predicted(-). (b) Distribution of $l_{30,29}$: observed(+), predicted(-).

Figure 2.3: (a) Overlap of Figures 1(a) and 1(b). (b) Percentage frequency distributions of $\rho_L$ (dashed) and $\rho_U$ (solid). $n = 25$.  

36
Figure 2.4: Percentage frequency distributions of $\rho_L$ (dashed) and $\rho_U$ (solid). (a) $n = 50$. (b) $n = 100$.

respectively. In each sample, we calculated $\rho_L$ and $\rho_U$ for each matrix $X$. Then the data of $\rho_L$ and $\rho_U$ were grouped into ten classes respectively. In the case of $\rho_L$, for example, the first class consists of matrices $X$ with $n^0 \leq \rho_L < n^{0.5}$ and the second class with $n^{0.5} \leq \rho_L < n^1$, the third one with $n^1 \leq \rho_L < n^{1.5}$ and so on. The number of matrices in each class was then divided by the corresponding sample size to get the percentage frequency to the class. The distributions have been plotted in the form of histograms. Empirically, there is a tendency that $\text{Prob}(n \leq \rho_L < n^{1.5})$ and $\text{Prob}(n \leq \rho_U < n^{1.5})$ tend to one as $n$ goes to infinity.

2.3 Complex Gaussian Elimination

All the results that were derived for the real case in §2.2.1 have straightforward analogues in the complex case.
2.3.1 Density Functions of $u_{pq}$

Let $X$ be an $n \times n$ complex Gaussian matrix and let $X = LU$ be the complex $LU$ factorization of $X$.

By applying the procedure of deriving (2.1) to the complex case now with unitary transforms replacing the corresponding orthogonal transforms, the following simple expression for the $(p,q)$-th $(2 \leq p \leq q)$ entry $u_{pq}$ of $U$ can be obtained,

$$u_{pq} = x_{pq} - \frac{sw_{p-1}}{r_{p-1p-1}},$$

where all the variables on the right-hand side are independent and $r_{p-1p-1}^2(r_{p-1p-1} \geq 0), s^2(s \geq 0), w_{p-1}$ and $x_{pq}$ are $\chi^2_2, \chi^2_{2(p-1)}, \tilde{N}(0,1)$ and $\tilde{N}(0,1)$ respectively.

Set

$$v = \frac{1}{p-1} \frac{s^2}{r_{p-1p-1}^2} = \frac{s^2/2(p-1)}{r_{p-1p-1}^2/2}$$

and

$$z = \frac{1}{\sqrt{1 + (p-1)v}} x_{pq} - \frac{\sqrt{(p-1)v}}{\sqrt{1 + (p-1)v}} w_{p-1},$$

then

$$u_{pq} = z \sqrt{1 + (p-1)v}.$$

The variables $v$ and $z$ are again independent and they are $F_{2(p-1),2}$ and $\tilde{N}(0,1)$ respectively.

**Theorem 2.5** Suppose $X$ is an $n \times n$ complex Gaussian matrix and let $X = LU$ be the $LU$ factorization of $X$. Then the density functions of $u_{pq}^R, u_{pq}^I$ and $|u_{pq}|$, the
real part, imaginary part and absolute value of the \((p, q)\)-th \((2 \leq p \leq q)\) entry \(u_{pq}\) of \(U\) respectively, are

\[
f_{u_{pq}}(t) = f_{u_{pq}}^i(t) = \frac{p-1}{\sqrt{2\pi}} \sum_{i=0}^{\infty} \frac{(-1)^i}{2^i i!} B(p-1, i+3/2) t^{2i},
\]

where \(-\infty < t < \infty\) and

\[
f_{|u_{pq}|}(t) = (-1)^p 2^p (p-1)! t^{1-2p} \left( \exp \left( -\frac{t^2}{2} \right) \left( 1 - p + \frac{t^2}{2} \right) - \sum_{i=0}^{p-1} (-1)^i \frac{1-p+i}{2^i i!} t^{2i} \right),
\]

where \(0 < t\).

**Proof.** Since \(v\) and \(z^R\) are independent and they are \(F_{2(p-1),2}\) and \(N(0,1)\) respectively, their joint density is

\[
f(v, z) = f_v(v) f_{z^R}(z)
\]

\[
= \begin{cases} 
\frac{1}{\sqrt{2\pi}} (p-1)^p v^{p-2} (1 + (p-1)v)^{-p} \exp(-z^2/2) & v > 0 \\
0 & \text{otherwise},
\end{cases}
\]

where \(z^R\) is the real part of \(z\). Thus the distribution function \(F_{u_{pq}}(\alpha)\) of \(u_{pq}^R\) is given by

\[
F_{u_{pq}}(\alpha) = \int_{u_{pq}^R \leq \alpha} f(v, z) dv dz
\]

\[
= \int_{z \sqrt{1+(p-1)v} \leq \alpha} \frac{1}{\sqrt{2\pi}} (p-1)^p v^{p-2} (1 + (p-1)v)^{-p} \exp(-z^2/2) dv dz
\]

\[
= \frac{1}{\sqrt{2\pi}} (p-1)^p \int_0^\infty dv \int_{-\infty}^{\sqrt{1+(p-1)v}} v^{p-2} (1 + (p-1)v)^{-p} \exp(-z^2/2) dz.
\]

39
Letting \( z = t/\sqrt{1 + (p-1)v} \),

\[
F_{u_{pq}}(\alpha) = \frac{1}{\sqrt{2\pi}} (p-1)^p \int_{0}^{\infty} dv \int_{-\infty}^{\alpha} v^{p-2} (1 + (p-1)v)^{-p-1/2} \exp\left(-\frac{1}{2} \frac{t^2}{1 + (p-1)v}\right) dt
\]

\[
= \frac{1}{\sqrt{2\pi}} (p-1)^p \int_{-\infty}^{\alpha} dt \int_{0}^{\infty} v^{p-2} (1 + (p-1)v)^{-p-1/2} \exp\left(-\frac{1}{2} \frac{t^2}{1 + (p-1)v}\right) dv.
\]

Letting \( s = \frac{1}{1 + (p-1)v} \),

\[
F_{u_{pq}}(\alpha) = \frac{1}{\sqrt{2\pi}} (p-1) \int_{-\infty}^{\alpha} dt \int_{0}^{1} s^{1/2}(1 - s)^{p-2} \exp\left(-\frac{1}{2} t^2 s\right) ds.
\]

Thus

\[
f_{u_{pq}}(t) = \frac{p-1}{\sqrt{2\pi}} \int_{0}^{1} s^{1/2}(1 - s)^{p-2} \exp\left(-\frac{1}{2} t^2 s\right) ds
\]

\[
= \frac{p-1}{\sqrt{2\pi}} \sum_{i=0}^{\infty} \frac{1}{i!} \left(-\frac{1}{2} t^2\right)^i \int_{0}^{1} s^{i+1/2}(1 - s)^{p-2} ds
\]

\[
= \frac{p-1}{\sqrt{2\pi}} \sum_{i=0}^{\infty} \frac{(-1)^i}{2i!} B(p - 1, i + 3/2) t^{2i}.
\]

To obtain the density of \(|u_{pq}|\), we notice that

\[
|u_{pq}| = |z| \sqrt{1 + (p-1)v}.
\]

Set

\[w = zR^2 + zI^2,\]

then

\[|u_{pq}| = \sqrt{w} \sqrt{1 + (p-1)v}.
\]
$w$ is $\chi^2_v$ and independent of $v$. Since the joint density of $w$ and $v$ is

$$f(w, v) = \begin{cases} 
\frac{1}{2}(p-1)^p v^{p-2} (1 + (p-1)v)^{-p} \exp(-w/2) & v, w > 0 \\
0 & \text{otherwise},
\end{cases}$$

the distribution function $F_{|u_{pq}|}(\alpha)$ of $|u_{pq}|$ is

$$F_{|u_{pq}|}(\alpha) = \iint_{|u_{pq}| \leq \alpha} f(w, v) dw dv$$

$$= \iint_{w(1+(p-1)v) \leq \alpha} \frac{1}{2}(p-1)^p v^{p-2} (1 + (p-1)v)^{-p} \exp(-w/2) dw dv$$

$$= \frac{1}{2}(p-1)^p \int_0^\infty dv \int_0^{\alpha^2/(1+(p-1)v)} v^{p-2} (1 + (p-1)v)^{-p} \exp(-w/2) dw$$

$$= (p-1)^p \int_0^\infty v^{p-2} (1 + (p-1)v)^{-p} \left(1 - \exp\left(-\frac{1}{2} \frac{\alpha^2}{1 + (p-1)v}\right)\right) dv.$$  

Letting $s = \frac{1}{1 + (p-1)v},$

$$F_{|u_{pq}|}(\alpha) = 1 - (p-1) \int_0^1 (1 - s)^{p-2} \exp \left(-\frac{\alpha^2}{2} s\right) ds.$$  

Therefore,

$$f_{|u_{pq}|}(t) = F'_{|u_{pq}|}(t)$$

$$= t(p-1) \int_0^1 (1 - s)^{p-2} s \exp \left(-\frac{t^2}{2} s\right) ds$$

$$= t(p-1) \int_0^1 (1 - s)^{p-2} s \sum_{i=0}^{\infty} (-1)^i \frac{t^{2i}}{2^i i!} s^i ds$$

41
\[ \begin{align*}
&= t(p-1) \sum_{i=0}^{\infty} (-1)^i \frac{t^{2i}}{2^i i!} \int_0^1 (1-s)^{p-2} s^{i+1} ds \\
&= t(p-1) \sum_{i=0}^{\infty} (-1)^i \frac{t^{2i}}{2^i i!} B(p-1, i+2) \\
&= t(p-1) \sum_{i=0}^{\infty} (-1)^i \frac{t^{2i}}{2^i i!} \frac{(p-2)!(i+1)!}{(i+p)!} \\
&= t(p-1)! \sum_{i=0}^{\infty} (-1)^i \frac{t^{2i}}{2^i} \frac{i+1}{(i+p)!}.
\end{align*} \]

Since
\[ \sum_{i=0}^{\infty} \frac{x^i}{(i+p)!} = \frac{d}{dx} \left( \sum_{i=0}^{\infty} \frac{x^{i+p}}{(i+p)!} \right) \]
\[ = \frac{d}{dx} \left( x^{1-p} \sum_{i=0}^{\infty} \frac{x^{i+p}}{(i+p)!} \right) \]
\[ = \frac{d}{dx} \left( x^{1-p} \left( \exp(x) - \sum_{i=0}^{p-1} \frac{x^i}{i!} \right) \right) \]
\[ = x^{-p} \exp(x)(1-p+x) - \sum_{i=0}^{p-1} \frac{i-p+1}{i!} x^{i-p}, \]

we have
\[ f_{\nu \rho i}(t) = (-1)^p 2^p (p-1)! t^{1-2p} \left( \exp \left( -\frac{t^2}{2} \right) \left( 1 - \frac{t^2}{2} \right) - \sum_{i=0}^{p-1} (-1)^i \frac{1-p+i}{2^i i!} t^{2i} \right). \]
2.3.2 Density Functions of $l_{pq}$

The expressions for the density functions of $l_{pq}$ are far simpler than those of $u_{pq}$ as shown in the real case. The situation is the same in the complex case. A similar derivation of (2.5) with unitary transforms replacing orthogonal transforms gives

$$l_{pq} = \frac{r_{q-1q-1}x_{pq} - sw_{q-1}}{r_{q-1q-1}x_{qq} - s\mu_{q-1}},$$

where the variables on the right-hand side are independent and $r_{q-1q-1}^2(r_{q-1q-1} > 0), s^2(s > 0)$ are $\chi^2_2, \chi^2(q-1)$ respectively and others are $\tilde{N}(0, 1)$.

Set

$$v = \frac{r_{q-1q-1}}{\sqrt{r_{q-1q-1}^2 + s^2}} x_{pq} - \frac{s}{\sqrt{r_{q-1q-1}^2 + s^2}} w_{q-1}$$

and

$$z = \frac{r_{q-1q-1}}{\sqrt{r_{q-1q-1}^2 + s^2}} x_{qq} - \frac{s}{\sqrt{r_{q-1q-1}^2 + s^2}} \mu_{q-1},$$

then $l_{pq} = v/z$. Both $v$ and $z$ are $\tilde{N}(0, 1)$ and independent. Thus,

$$l_{pq} = \frac{vR + izI}{zR + izI}$$

$$= \frac{vRxR + vIxI + iz(vIzR - vRzI)}{zR^2 + zI^2}$$

and therefore,

$$l_{pq}^R = \frac{vRxR + vIxI}{zR^2 + zI^2}$$

$$= \frac{vRxR + vIxI}{\sqrt{zR^2 + zI^2}} \sqrt{zR^2 + zI^2}$$

43
\[ \equiv \frac{\xi}{\sqrt{\eta}} \]

\[ = \frac{1}{\sqrt{2}} \frac{\xi}{\sqrt{\eta/2}} \]

\[ \equiv \frac{1}{\sqrt{2}} \zeta. \]

\( \eta \) and \( \xi \) are \( \chi^2 \) and \( N(0,1) \) respectively and independent. Hence \( \zeta \) has the student's \( t \) distribution with 2 degrees of freedom. Thus,

\[ f_{l_{pq}}(t) = \frac{1}{2} (1 + t^2)^{-3/2}, \]

where \(-\infty < t < \infty\). Moreover,

\[ |l_{pq}| = |v|/|z| = \sqrt{\frac{|v|^2/2}{|z|^2/2}} = \sqrt{\xi}, \]

where \( \xi \) has \( F_{2,2} \) distribution. Hence

\[ f_{|l_{pq}|}(t) = \begin{cases} \frac{2t}{(1 + t^2)^2} & t > 0 \\ 0 & \text{otherwise} \end{cases}. \]

We summarize the above results about \( l_{pq} \) in the following theorem.

**Theorem 2.6** Suppose \( X \) is an \( n \times n \) complex Gaussian matrix and let \( X = LU \) be the LU factorization of \( X \). Then the density functions of \( l_{pq}^R, l_{pq}^I \) and \( |l_{pq}| \), the real part, imaginary part and absolute value of the \((p,q)\)-th \((1 \leq q < p \leq n)\) entry \( l_{pq} \) of \( L \) respectively, are

\[ f_{l_{pq}^R}(t) = f_{l_{pq}^I}(t) = \frac{1}{2} (1 + t^2)^{-3/2}, \]
where $-\infty < t < \infty$ and

$$f_{\text{exp}}(t) = \frac{2t}{(1+t^2)^2},$$

where $0 < t$.

Figures 2.5-2.7 plot the empirical densities of some entries of $L$ and $U$ and their corresponding predicted densities in Theorems 2.5 and 2.6.

![Graphs](image)

Figure 2.5: (a) Distribution of $u_{12,12}^R$: observed(+), predicted(-). (b) Distribution of $|u_{12,12}|$: observed(+), predicted(-).

2.4 Application

In some kinds of parallel computations, such as in a systolic system [37], Gaussian elimination without pivoting is preferred to GE with pivoting because the latter version is difficult to implement due to data movement which can not be determined a priori and, as a result, leads to high overhead both in time and
Figure 2.6: (a) Distribution of $u_{31,31}^R$: observed(+) predicted(-). (b) Distribution of $|u_{31,31}|$: observed(+), predicted(-).

Figure 2.7: (a) Distribution of $l_{13,12}^R$: observed(+), predicted(-). (b) Distribution of $|l_{13,12}|$: observed(+), predicted(-).
memory [40]. However, as we mentioned earlier, there are two difficulties about
the implementation of GE without pivoting, that is, breakdown and instability.
Recently, Lê, Parker and Pierce (LPP) [40, 42, 43] suggested an effective method
to overcome these difficulties in which they first randomized the original matrix
and then applied GE without pivoting to the resulting matrix. More precisely, let
A be an $n \times n$ matrix and $V$ a random $n \times n$ matrix from some special class. Their
basic idea is that they perform GE without pivoting on the product matrix $VA$
instead of $A$. Experimental results have shown that this indeed greatly reduces
the probability of breakdown and improves the potential ill-conditioning of $A$.

In this section, we describe an alternative method to that of LPP's to cure
breakdown and instability in GE without pivoting. Based on eight random classes
of matrices, Trefethen and Schreiber [57] observed by experiments that, after the
first few steps of GE with partial (or complete) pivoting, the remaining matrix ele-
ments are approximately normally distributed, regardless of what class the matrix
comes from. On the other hand, as we have seen in §2.2, if GE without pivoting is
applied to a Gaussian matrix, then the probability of the occurrence of a pivot less
than $\epsilon$ in magnitude is $O(\epsilon)$ and $\text{Prob}(n \leq \rho_L < n^{1.5}) \approx 1$, $\text{Prob}(n \leq \rho_U < n^{1.5}) \approx 1$
when $n$ is large. All of these observations suggest that we may use GE in such a
way that, in the first few steps, we use GE with partial (or complete) pivoting and
then use the version without pivoting (or with pairwise, neighbor pivoting). The
goal of the use of GE with pivoting in the early steps is to make the elements ap-
proximately normally distributed so that the chances of breakdown and instability could be reduced when a simpler form of GE is performed later on. We have done some experiments on this idea based on two random classes of matrices:

- Class 1: independent elements from the discrete distribution with $P(-1) = P(1) = 0.5$
- Class 2: independent elements from the discrete distribution with $P(0) = P(1) = 0.5$

Since both the results from these two classes are analogous, we will only show those for Class 1 below. The scheme we adopted in our experiments is the following. We perform GE with complete pivoting on $A$ in the first $k$ steps to get the decomposition,

$$PAQ = \begin{bmatrix} L_{11} & 0 \\ L_{21} & I \end{bmatrix} \begin{bmatrix} U_{11} & U_{12} \\ 0 & U_{22} \end{bmatrix},$$

where $P$ and $Q$ are permutation matrices, $U_{11}$ and $L_{11}$ are $k \times k$ upper and unit lower triangular matrices respectively and $I$ is the $(n-k) \times (n-k)$ identity matrix. Based on the observations of Trefethen and Schreiber, we may expect that the elements in $U_{22}$ are approximately normally distributed. Then we perform GE without pivoting on $U_{22}$ to get the $LU$ decomposition of $U_{22}$,

$$U_{22} = \tilde{L}_{22} \tilde{U}_{22}.$$
Thus we obtain the following form of $LU$ decomposition for $A$,

\[
PAQ = \begin{bmatrix}
L_{11} & 0 \\
L_{21} & L_{22}
\end{bmatrix} \begin{bmatrix}
U_{11} & U_{12} \\
0 & \tilde{U}_{22}
\end{bmatrix}.
\]

Although the entries of $U_{22}$ are not independent, we may still expect, according to the theory we derived in earlier sections, that the probability of the occurrence of small pivots and large growth factors is very small in the $LU$ factorization of $U_{22}$.

![Figure 2.8](image)

**Figure 2.8**: Observed distribution (+) of the elements of $U_{22}$ vs. the $N(0, 1.77)$ distribution (-). $k = 20, n = 51$.

Figure 2.8 plots the empirical distribution of the elements of $U_{22}$ together with the normal distribution $N(0, 1.77)$, where we chose $k = 20$ and $n = 51$. We can see that the elements of $U_{22}$ are indeed almost normally distributed, even though it is obvious that they are correlated. Based on this observation, we can expect that the elements of $\tilde{L}_{22}$ and $\tilde{U}_{22}$ would obey the distributions described in Theorems 2.1 and 2.2. In fact, Figures 2.9 and 2.10 illustrate this point, where we plotted the empirical distributions of the elements $\tilde{l}_{13,12}, \tilde{l}_{30,29}, \tilde{u}_{12,12}$ and $\tilde{u}_{31,31}$ of $\tilde{L}_{22}$ and $\tilde{U}_{22}$.
Figure 2.9: (a) Observed distribution (+) of the (12, 12)-th entry of $\tilde{U}_{22}$ vs. the predicted distribution $0.7f_{u_{12,12}}(0.7t)$ (-). (b) Observed distribution (+) of the (31, 31)-th entry of $\tilde{U}_{22}$ vs. the predicted distribution $0.8f_{u_{31,31}}(0.8t)$ (-). The functions $f_{u_{12,12}}(t)$ and $f_{u_{31,31}}(t)$ are defined in Theorem 2.1.

Figure 2.10: (a) Observed distribution (+) of the (13, 12)-th entry of $\tilde{L}_{22}$ vs. the predicted distribution $f_{l_{13,12}}(t)$ (-). (b) Observed distribution (+) of the (30, 29)-th entry of $\tilde{L}_{22}$ vs. the predicted distribution $f_{l_{30,29}}(t)$ (-). The functions $f_{l_{13,12}}(t)$ and $f_{l_{30,29}}(t)$ are defined in Theorem 2.2.
Figure 2.11: Percentage frequency distributions of $\rho_{L_{22}}$ (dashed) and $\rho_{U_{22}}$ (solid).
(a) $k = 20, n = 45$. (b) $k = 20, n = 70$. See Figures 2.3 and 2.4 for the meaning of $\alpha$.

Figure 2.12: Percentage frequency distributions of $\rho_{L_{22}}$ (dashed) and $\rho_{U_{22}}$ (solid).
(a) $k = 20, n = 120$. (b) $k = 20, n = 220$. See Figures 2.3 and 2.4 for the meaning of $\alpha$. 

51
Table 2.2: 12500 matrices from Class 1 were selected in this experiment. B: the number of matrices which make GE without pivoting performed on $U_{22}$ breakdown; EP: the empirical probability of breakdown when GE without pivoting performed on $U_{22}$

<table>
<thead>
<tr>
<th>$k$</th>
<th>$n = 64$</th>
<th></th>
<th></th>
<th>$n = 128$</th>
<th></th>
<th></th>
<th>$n = 256$</th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>12359</td>
<td>9.8872 x 10^{-1}</td>
<td>12378</td>
<td>9.9024 x 10^{-1}</td>
<td>12368</td>
<td>9.8944 x 10^{-1}</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>5</td>
<td>9273</td>
<td>7.4184 x 10^{-1}</td>
<td>9408</td>
<td>7.5264 x 10^{-1}</td>
<td>9324</td>
<td>7.4592 x 10^{-1}</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>10</td>
<td>926</td>
<td>7.408 x 10^{-2}</td>
<td>936</td>
<td>7.488 x 10^{-2}</td>
<td>935</td>
<td>7.48 x 10^{-2}</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>15</td>
<td>21</td>
<td>8.16 x 10^{-3}</td>
<td>14</td>
<td>1.12 x 10^{-3}</td>
<td>19</td>
<td>1.52 x 10^{-3}</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>20</td>
<td>1</td>
<td>8 x 10^{-5}</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>25</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td>0</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

$\hat{U}_{22}$ respectively versus their corresponding (scaled) predicted distributions. Both the observed and predicted distributions match very well in these two figures.

Table 2.2 shows some information on the situation of breakdown. In this experiment, 12500 matrices were selected from Class 1 for each matrix order $n$. From the table, we can see that the probability of breakdown is almost the same as $n$ varies for each fixed $k$ while it decreases as $k$ increases for $n$ fixed. If this were true for the general case, then we would be able to avoid breakdown in GE without pivoting with little data movement by choosing small $k$, regardless of how large the matrix size is.

In the case of the Gaussian matrix, we have seen that the probabilities $\text{Prob}(m \leq \rho_L < n^{1.5})$ and $\text{Prob}(n \leq \rho_U < n^{1.5})$ tend to one as $n$ goes to infinity. The situation seems to be the same for the Class 1 matrices as shown in Figures 2.11 and 2.12.

More precisely, for $k$ fixed, we may have

$$\lim_{n \to +\infty} \text{Prob}(n \leq \rho_{L_{22}} < n^{1.5}) = 1$$
and

\[ \lim_{n \to +\infty} \text{Prob}(n \leq \rho_{\tilde{U}_{22}} < n^{1.5}) = 1, \]

where \( \rho_{\tilde{L}_{22}} \equiv \|\tilde{L}_{22}\|_\infty \) and \( \rho_{\tilde{U}_{22}} \equiv \|\tilde{U}_{22}\|_\infty / \|U_{22}\|_\infty \).

We end this section with the conclusion that the method we discussed here may be a good choice in solving GE related problems to avoid breakdown, instability and data movements in large quantities. We can not say any more on this idea at this point because we only did a few experiments on some special classes. We mention it just for the references of the GE users.
CHAPTER 3

Preserving Symmetry in Preconditioned Krylov Subspace Methods

We consider the problem of solving a linear system

\[ Ax = b \]

by a preconditioned Krylov subspace method when \( A \) is nearly symmetric and when the system is preconditioned by a symmetric positive definite matrix \( M \). In the symmetric case, one can recover symmetry by using \( M \)-inner products in the conjugate gradient (CG) algorithm. This idea can also be used in the nonsymmetric case, and near symmetry can be preserved similarly. Like CG, the new algorithms are mathematically equivalent to split preconditioning, but do not require \( M \) to be factored. Better robustness in a specific sense can also be observed. When combined with truncated versions of iterative methods, tests show that this is more effective than the common practice of forfeiting near-symmetry altogether.

This chapter is organized as follows. In §1, it is shown how alternative inner products may be used to preserve symmetry in GMRES. §2 considers the use of truncated iterative methods when the preconditioned system is close to being symmetric. This has been hypothesized by many authors, for example, Axelsson [5] and Meyer [41]. In §3 we consider the symmetrically preconditioned Bi-CG
algorithm. §4 shows the results of tests on a Navier-Stokes problem. This problem is parameterized by the Reynolds number, and thus nearness to symmetry. We conclude this chapter in §5.

### 3.1 Symmetric preconditioning in GMRES

When \( A \) is nearly symmetric, split preconditioning may be used to preserve the original degree of symmetry. Alternatively, left-preconditioning with the \( M \)-inner product, or right-preconditioning with the \( M^{-1} \)-inner product may be used. We refer to these latter two preconditionings as *symmetric*.

Consider first the left-preconditioned version of the Arnoldi algorithm based on the classical Gram-Schmidt process, with preconditioner \( M \), which we assume to be SPD.

**Algorithm 2.1. Left-Preconditioned Arnoldi-Classical Gram-Schmidt**

1. Choose a vector \( v_1 \) of norm 1.
2. For \( j = 1, 2, \cdots, m \)
   3. \( w_j = Av_j \);
   4. \( z = M^{-1}w_j \);
   5. Compute \( h_{ij} = (z, v_j) \) for \( i = 1, 2, \cdots, j \);
   6. \( z = z - \sum_{i=1}^{j} h_{ij}v_i \);
   7. \( h_{j+1,j} = \|z\|_2 \);
   8. If \( h_{j+1,j} = 0 \) then stop;
   9. \( v_{j+1} = z/h_{j+1,j} \);
10. End
To implement this procedure using $M$-inner products, we would first compute the scalars $h_{ij}$ in Line 5,

$$h_{ij} = (z_j, v_i)_M = (Mz_j, v_i) = (w_j, v_i), \quad i = 1, \ldots, j.$$  \quad (3.1)

Then we would modify the vector $z_j$ to obtain the next Arnoldi vector (before normalization),

$$\hat{z}_j = z_j - \sum_{i=1}^j h_{ij}v_i.$$ \quad (3.2)

To complete the orthonormalization step we must normalize the final $\hat{z}_j$. Because of the $M$-orthogonality of $\hat{z}_j$ versus all previous $v_i$'s we observe that

$$((\hat{z}_j, \hat{z}_j)_M = (z_j, \hat{z}_j)_M = (M^{-1}w_j, \hat{z}_j)_M = (w_j, \hat{z}_j).$$ \quad (3.3)

Thus, the desired $M$-norm can be computed according to (3.3) and then computing

$$h_{j+1,j} = (\hat{z}_j, w_j)^{1/2} \quad \text{and} \quad v_{j+1} = \hat{z}_j / h_{j+1,j}.$$ \quad (3.4)

One potentially serious difficulty with the above procedure is that the inner product $(\hat{z}_j, \hat{z}_j)_M$ as computed by (3.3) may be negative in the presence of round-off. Indeed, loss of $M$-orthogonality between $z_j$ and the previous $v_i$'s causes the first equality in (2.3) to be invalid in the presence of rounding, which may result in a negative inner product $(w_j, \hat{z}_j)$.

There are two remedies. First, we can compute this $M$-norm explicitly at the expense of an additional matrix-vector multiplication with $M$, i.e., from

$$((\hat{z}_j, \hat{z}_j)_M = (M\hat{z}_j, \hat{z}_j).$$
As was pointed out earlier, this is undesirable, since the operator $M$ is often not available explicitly. Indeed in many cases, only the preconditioning operation $M^{-1}$ is available from a sequence of operations, as is the case for multigrid preconditioning. Another difficulty with computing $h_{ij}$ with (3.1) is that it is not immediately amenable to a modified Gram-Schmidt implementation. Indeed, consider the first step of a hypothetical modified Gram-Schmidt step, which consists of $M$-orthonormalizing $z$ against $v_1$,

$$h_{i1} = (z, v_1)_M, \quad \hat{z} = z - h_{i1}v_1.$$  

As was observed, the inner product $(z, v_1)_M$ is equal to $(w, v_1)$ which is computable. Now we need $M\hat{z}$ to compute $(z - h_{i1}v_1, v_i)_M$ in the modified Gram-Schmidt process. However, no such vector is available, and we can only compute the $(z, v_i)_M$ of classical Gram-Schmidt.

An alternative is to save the set of vectors $Mv_i$ (again, not computed by multiplying by $M$) which would allow us to accumulate inexpensively both the vector $\hat{z}_j$ and the vector $\tilde{w}_j$ via the relation

$$\hat{w}_j \equiv M\hat{z}_j = w_j - \sum_{i=1}^{j} h_{ij}Mv_i,$$

which is obtained from (3.2). Now the inner product $(\hat{z}_j, \hat{z}_j)_M$ is given by

$$(\hat{z}_j, \hat{z}_j)_M = (M^{-1}\hat{w}_j, M^{-1}\hat{w}_j)_M = (M^{-1}\hat{w}_j, \tilde{w}_j).$$

In this form, this inner product is guaranteed to be nonnegative as desired. This leads to the following algorithm.
Algorithm 2.2. Arnoldi-Classical Gram-Schmidt and $M$-inner products

1. Choose a vector $w_1$ such that $v_1 = M^{-1}w_1$ has $M$-norm 1.
2. For $j = 1, 2, \cdots, m$
3. $w = Av_j$;
4. Compute $h_{ij} = (w, v_i)$ for $i = 1, 2, \cdots, j$;
5. $\hat{w} = w - \sum_{i=1}^{j} h_{ij}w_i$;
6. $z = M^{-1}\hat{w}$;
7. $h_{j+1,j} = (z, \hat{w})^{1/2}$; If $h_{j+1,j} = 0$ then stop;
8. $w_{j+1} = \hat{w}/h_{j+1,j}$;
9. $v_{j+1} = z/h_{j+1,j}$;
10. End

As is noted, the above algorithm requires that we save two sets of vectors: the $v_j$'s and the $w_i$'s. The $v_j$'s form the needed Arnoldi basis, and the $w_i$'s are required when computing the vector $\hat{w}_j$ in Line 5. If we do save these two sets of vectors we can also easily formulate the algorithm with the modified Gram-Schmidt version of the Arnoldi procedure.

The situation for right-preconditioning with the $M^{-1}$-inner product is much simpler, mainly because $M^{-1}z$ is available when $z$ needs to be normalized in the $M^{-1}$ norm. It is only necessary to note that $M^{-1}z$ is normally computed at the next iteration in the standard Arnoldi algorithm. Again, both the $v$'s and the $w$'s need to be saved, where $w_j = M^{-1}v_j$ in this case.
The additional storage of the $w$'s, however, makes the algorithm naturally flexible, i.e., it accommodates the situation where $M$ varies at each step as when $M^{-1}v$ is the result of some unspecified computation. If $M^{-1}$ is not a constant operator, then a basis for the right-preconditioned Krylov subspace cannot be constructed from the $v$'s alone. However, the vectors $w_j = M_j^{-1}v_j$ do form a basis for this subspace, where $M_j^{-1}$ denotes the preconditioning operation at the $j$-th step. The use of this extra set of vectors is exactly how the standard flexible variant of GMRES is implemented [49].

We can now write preconditioned GMRES algorithms using these implementations of alternative inner products in the Gram-Schmidt process. In the left-preconditioned case, we can minimize the $M$-norm of the preconditioned residual vector $M^{-1}(b - Ax)$ by simply minimizing the 2-norm of the projected system in the standard GMRES algorithm. To be formal, we state the following theorem without proof.

**Theorem 3.1** The approximate solution $x_m$ obtained from the left-preconditioned GMRES algorithm with $M$-inner products minimizes the residual $M$-norm $\|M^{-1}(b - Ax)\|_M$ over all vectors of the affine subspace $x_0 + K_m$ in which

$$K_m = \text{span}\{z_0, M^{-1}Az_0, \ldots, (M^{-1}A)^{m-1}z_0\}$$  \hspace{1cm} (3.5)

where $z_0 = M^{-1}r_0$. Also, the approximate solution $x_m$ obtained from the right-preconditioned GMRES algorithm with $M^{-1}$-inner products minimizes the residual $M^{-1}$-norm $\|b - Ax\|_{M^{-1}}$ over the same affine subspace.
We can show that both left and right symmetric preconditioning are mathematically equivalent to split preconditioning. All that must be noticed is that minimizations of

\[ \|L^{-1}(b - Ax)\|_2 = \|b - Ax\|_{M^{-1}} = \|M^{-1}(b - Ax)\|_M \]

are the same, and that the minimizations are over the same subspace in each of the left, right, and split preconditioning options [50, §9.3.4]. We emphasize in particular that it is the split preconditioned residual that is minimized in all three algorithms.

### 3.2 Truncated iterative methods

Truncated iterative methods are an alternative to restarting, when the number of steps required for convergence is large and the computation and storage of the Krylov basis becomes excessive. When \( A \) is exactly symmetric, a three-term recurrence governs the vectors in the Arnoldi process, and it is only necessary to orthogonalize the current Arnoldi vector against the previous two vectors. If \( A \) is nearly symmetric, an incomplete orthogonalization against a small number of previous vectors may be advantageous over restarted methods. The advantage here may offset the cost of maintaining the extra set of vectors to maintain the initial degree of symmetry. The incomplete Arnoldi procedure outlined below stores only the previous \( k \) Arnoldi vectors, and orthogonalizes the new vectors against them.
It differs from the full Arnoldi procedure only in Line 4, which would normally be a loop from 1 to \( j \). It can be considered to be the full Arnoldi procedure when \( k \) is set to infinity.

**Algorithm 3.1. Incomplete Arnoldi Procedure**

1. Choose a vector \( v_1 \) of norm 1.
2. For \( j = 1, 2, \ldots, m \)
3. \( w = Av_j \);
4. For \( i = \max\{1, j - k + 1\}, \ldots, j \)
5. \( h_{ij} = (w, u_i) \);
6. \( w = w - h_{ij} u_i \);
7. End;
8. \( h_{j+1,j} = \|z\|_2 \);
9. If \( h_{j+1,j} = 0 \) then stop;
10. \( v_{j+1} = w/h_{j+1,j} \);
11. End

The truncated version of GMRES uses this incomplete Arnoldi procedure and is called Quasi-GMRES [13]. The practical implementation of this algorithm allows the solution to be updated at each iteration, and is thus called a 'direct' version, or DQGMRES [51].

To suggest that truncated iterative methods may be effective in cases of near symmetry, we study the asymptotic behavior of the iterates of DQGMRES as the coefficient matrix \( A \) varies from nonsymmetry to (skew) symmetry. We first
decompose \( A \) as

\[
A = S + B
\]

in which \( S \) is symmetric or skew symmetric, and set

\[
\varepsilon = \| B \|_2.
\]

We will first establish asymptotic relations among the variables in the incomplete and full Arnoldi procedures. Then we will apply the incomplete procedure to \( A \), and the full procedure to \( S \), using the superscripts \( I \) and \( F \) to distinguish between the variables appearing in the two procedures. (Note that since \( S \) is (skew) symmetric, the full procedure on \( S \) is the same as the incomplete procedure with \( k \geq 2 \).)

Moreover, if we denote the degree of the minimal polynomial of \( v^I_1 \) with respect to \( S \) by \( \nu \), then \( h^F_{\nu+1} \neq 0 \) and \( h^F_{j+1,j} \neq 0 \) for \( 1 \leq j < \nu \). In the proof of the following lemma, we also use \( \hat{v}^I_j \) and \( \hat{v}^F_j \) to denote the vectors \( w^I \) and \( w^F \) obtained at the end of Line 7 in the incomplete and complete Arnoldi procedures.

**Lemma 3.1** Assume the truncation parameter \( k \geq 2 \). If \( v^I_i = v^I_1 + O(\varepsilon) \), then

\[
h^I_{ij} = h^F_{ij} + O(\varepsilon), \quad v^I_j = v^F_j + O(\varepsilon)
\]

where \( 1 \leq j \leq \nu \) and \( \max\{1, j-k+1\} \leq i \leq j+1 \).

**Proof.** The proof is by induction on the index \( j \). By Lines 5 and 6 of the Arnoldi procedure,

\[
h^I_{11} = (Av^I_1, v^I_1), \quad \hat{v}^I_2 = Av^I_1 - h^I_{11} v^I_1, \quad h^I_{21} = \| \hat{v}^I_2 \|_2,
\]

62
we have

\[ h_{11}^I = (Sv_1^F, v_1^F) + O(\varepsilon) = h_{11}^F + O(\varepsilon), \]

\[ \hat{v}_2^I = Sv_1^F - h_{11}^F v_1^F + O(\varepsilon) = \hat{v}_2^F + O(\varepsilon), \]

\[ h_{21}^I = ||\hat{v}_2^F||_2 + O(\varepsilon) = h_{21}^F + O(\varepsilon) \]

and hence the lemma holds for \( j = 1 \). Assume that the lemma has been proved for \( j < j_0 \leq \nu \). On that hypothesis, we prove it for \( j = j_0 \). By Line 10 of the Arnoldi procedure,

\[ v_{j_0}^I = \hat{v}_{j_0}^I / h_{j_0, j_0-1}^I \]

which yields that

\[ v_{j_0}^I = \frac{\hat{v}_{j_0}^F + O(\varepsilon)}{h_{j_0, j_0-1}^F + O(\varepsilon)} = \frac{\hat{v}_{j_0}^F}{h_{j_0, j_0-1}^F} + O(\varepsilon) = v_{j_0}^F + O(\varepsilon) \]

by the induction hypothesis. Therefore

\[ w^I = Av_{j_0}^I = Sv_{j_0}^F + O(\varepsilon) = w^F + O(\varepsilon) \]

for the \( w^I \) and \( w^F \) in Line 3 of the Arnoldi procedures. Using another induction on the index \( i \) in Lines 5 and 6, and the induction hypothesis on \( j \) and, in the mean time, noting that \( h_{i_0}^F = 0 \) for \( 1 \leq i \leq j_0 - 2 \), we have

\[ h_{i_0}^I = h_{i_0}^F + O(\varepsilon) \]

for \( \max\{1, j_0 - k + 1\} \leq i \leq j_0 \) and

\[ \hat{v}_{j_0+1}^I = \hat{v}_{j_0+1}^F + O(\varepsilon). \]
From the last equation,
\[ h_{j_0+1,j_0}^I = \| \hat{v}_{j_0+1}^I \|_2 = \| \hat{v}_{j_0+1}^F \|_2 + O(\varepsilon) = h_{j_0+1,j_0}^F + O(\varepsilon) \]
and then the induction step is complete. \( \square \)

We now turn to the DQGMRES algorithm. Consider the linear system
\[ Sx = b \]
and denote by \( x_m^G \) and \( x_m^Q \) the approximate solutions by the GMRES and DQGMRES algorithms, respectively. Let \( \mu \) be the degree of the minimal polynomial of the vector \( b - Sx_0 \) with respect to \( S \). A result of the lemma can be stated as follows.

**Theorem 3.2** Given the same initial guess \( x_0 \) to GMRES and DQGMRES with \( k \geq 2 \), then at any given step \( m \) with \( 1 \leq m \leq \mu \),
\[ x_m^Q = x_m^G + O(\varepsilon). \]

**Proof.** By the definitions of DQGMRES and GMRES, we have
\[ x_m^Q = x_0 + \beta^Q V_m^I \left( \left( \tilde{H}_m^I \right)^T \tilde{H}_m^I \right)^{-1} \left( \tilde{H}_m^I \right)^T e_1 \]
and
\[ x_m^G = x_0 + \beta^G V_m^F \left( \left( \tilde{H}_m^F \right)^T \tilde{H}_m^F \right)^{-1} \left( \tilde{H}_m^F \right)^T e_1 \]
where \( \beta^Q = \| b - Ax_0 \|_2 \) and \( \beta^G = \| b - Sx_0 \|_2 \). Since
\[ v_1^I = \frac{1}{\beta^Q} (b - Ax_0) = \frac{1}{\beta^G} (b - Sx_0) + O(\varepsilon) = v_1^F + O(\varepsilon), \]

64
we have by the lemma,

\[ \bar{H}_m^I = \bar{H}_m^E + O(\varepsilon), \quad V_m^I = V_m^E + O(\varepsilon) \]

and therefore the desired equation holds. \( \square \)

If we let \( x_A \) be the exact solution to \( Ax = b \) and \( x_S \) be the exact solution to \( Sx = b \), then it is obvious that \( x_A = x_S + O(\varepsilon) \). Since, on the other hand, \( x_\mu^G = x_S \) we immediately have the following corollary.

**Corollary 3.1** For any initial guess \( x_0 \) and any \( k \geq 2 \),

\[ x_\mu^G = x_A + O(\varepsilon). \]

The corollary suggests that we may use DQGMRES with small \( k \) when \( A \) is nearly symmetric or nearly skew symmetric. It may be possible to generalize the above theory to the class of CG(s) matrices of Faber and Manteuffel [21], i.e., the class of matrices for which it is possible to define Krylov subspace algorithms using \( s \)-term recurrences.

### 3.3 Symmetric preconditioning in Bi-CG

Both left-symmetric and right-symmetric preconditioning of the Bi-CG algorithm are relatively straightforward, and no extra vectors are required. In the right-preconditioned Bi-CG algorithm with coefficient matrix \( A \) and preconditioner \( M \), the preconditioned coefficient matrix of the dual system is \((AM^{-1})^T = M^{-T}AT\),
i.e., the dual residual $r^*$ is the residual of

$$M^{-T}A^T x^* = b^*,$$

which is a left-preconditioned version of some linear system with $A^T$. To develop
the symmetric right-preconditioned Bi-CG algorithm using $M^{-1}$-inner products,
the preconditioned coefficient matrix of the dual system must be the adjoint of
$AM^{-1}$ in the $M^{-1}$ inner product. This is $A^TM^{-1}$ as shown by

$$(AM^{-1}x, y)_{M^{-1}} = (M^{-1}AM^{-1}x, y) = (x, M^{-1}A^TM^{-1}y) = (x, A^TM^{-1}y)_{M^{-1}}.$$

The dual system thus involves the coefficient matrix $A^TM^{-1}$.

Like GMRES, both left and right symmetric preconditioned versions of Bi-CG
are equivalent to the split preconditioned version, and this can be shown by a
change of variables. However, in both left and right symmetric preconditioned
versions, the exact, rather than the split preconditioned residual is available.

The unpreconditioned Bi-CG algorithm cannot have a serious breakdown if $A$
is SPD and $r_0^*$ is chosen to be $r_0$. This is because $r_j^* = r_j$ and $p_j^* = p_j$ for all $j$ and
the vectors $Ap_j, p_j^*$ never become orthogonal. In fact, the cosine

$$\frac{(Ap_j, p_j^*)}{\|Ap_j\|\|p_j^*\|} = \frac{(Ap_j, p_j^*)}{\|p_j\|\|p_j^*\|\|Ap_j\|}$$

can be bounded below by the reciprocal of the condition number of $A$.

Similarly, in the symmetric right-preconditioned version of Bi-CG, if both $A$
and $M$ are SPD, and $r_0^* = r_0$, then $r_j^* = r_j$ and $p_j^* = p_j$ for all $j$, and

$$
\frac{(Ap_j, p_j^*)}{\|Ap_j\|\|p_j^*\|} \geq \text{cond}^{-1}(A)
$$

$$
\frac{(M^{-1}r_j, r_j^*)}{\|M^{-1}r_j\|\|r_j^*\|} \geq \text{cond}^{-1}(M).
$$

We measure the cosines rather than the quantities $(Ap_j, p_j^*)$ and $(M^{-1}r_j, r_j^*)$ because the $p$ and $r$ vectors have magnitudes going to 0 as the algorithms progress. Recall that in the case when $(M^{-1}r_j, r_j^*) = 0$ and $r_j = 0$, we have a lucky breakdown.

For the case of regular right- or left-preconditioning, or if $r_0^* \neq r_0$ in the symmetrically preconditioned cases, then no such lower bounds as the above exist, and the algorithms are liable to break down.

When $A$ is near-symmetric, it is our hypothesis that the probability of breakdown is lower in the symmetrically preconditioned cases, and this will be shown by experiment in the next section.

### 3.4 Numerical Experiments

§5.1 tests the idea of using symmetric preconditionings with truncated iterative methods. §5.2 tests the breakdown behavior of symmetrically preconditioned BiCG.
3.4.1 Truncated iterative methods

To test the idea of using symmetric preconditionings with truncated iterative methods for nearly symmetric systems, we selected a standard fluid flow problem where the degree of symmetry in the matrices is parameterized by the Reynolds number. The flow problem is the two-dimensional square-lid driven cavity, and was discretized by the Galerkin Finite Element method. Rectangular elements were used, with biquadratic basis functions for velocities, and linear discontinuous basis functions for pressure. We considered a segregated solution method for the Navier-Stokes equations, where the velocity and pressure variables are solved separately; the matrices arising from a fully-coupled solution method are otherwise indefinite. In particular, we considered the expression of the conservation of momentum,

\[ \text{Re}(u \cdot \nabla u) = -\nabla p + \nabla^2 u \]

where \( u \) denotes the vector of velocity variables, \( p \) denotes the pressure variable, and \( \text{Re} \) is the Reynolds number. The boundary conditions for the driven cavity problem over the unit square are \( u = (1, 0)^T \) on the top edge of the square, and \( u = (0, 0)^T \) on the other three sides and the corners. The reference pressure specified at the bottom-left corner is 0.

The matrices are the initial Jacobians at each Newton iteration, assuming a zero pressure distribution. For convenience, however, we chose the right-hand sides of the linear systems to be the vector of all ones. A mesh of 20 by 20 elements was used, leading to momentum equation matrices of order 3042 and having 91204
nonzero entries. The nodes corresponding to the boundaries were not assembled into the matrix, and the degrees of freedom were numbered element by element. For Reynolds number 0., the matrix is SPD, and is equal to the symmetric part of the matrices with nonzero Reynolds number.

We generated matrices with Reynolds number less than 10, which gives rise to the nearly symmetric case. For Reynolds number 1., the degree of symmetry measured by

\[
\frac{\|A - A^T\|_F}{\|A + A^T\|_F}
\]

has value $7.5102 \times 10^{-4}$ and this measure increases linearly with the Reynolds number (at least up to $Re = 10$).

In the numerical experiments below, we show the number of matrix-vector products consumed by GMRES$(k)$ and DQGMRES$(k)$ to reduce the actual residual norm to less than $10^{-6}$ of the original residual norm, with a zero initial guess. Several values of $k$ are used. A dagger $(\dagger)$ in the tables indicates that there was no convergence within 500 matrix-vector products. The incomplete Choleski factorization IC(0) of the $Re = 0$ problem was used as the preconditioner in all the problems.

For comparison, we first show in Table 3.1 the results using standard right-preconditioning. Table 3.2 shows the results using right-preconditioning with $M^{-1}$ inner products, or equivalently, split preconditioning. The right-preconditioned methods have a slight advantage in this comparison (by as many as 20 Mat-Vec's),
since they directly minimize the actual residual norm, whereas the symmetrically
preconditioned methods minimize a preconditioned residual norm.

When split preconditioning is used, the actual, or non-preconditioned resid-
ual norm is not available. In this case, for GMRES, we use the preconditioned
residual norm in the stopping criterion, and check that the actual residual norm is
within twice the tolerance at the end of the iterations. For DQGMRES, since an
accurate residual norm estimate is not available within the algorithm, the actual
residual norm was computed and used for the stopping criterion for the purpose
of this comparison. (This slight mismatch in the stopping criterion accounts for
the difference between GMRES(∞) and DQGMRES(2) for the Re = 0 case.) In
practice, the stopping criterion should use the preconditioned residual norm that
is available, not necessarily the actual one.

| Table 3.1: Mat-Vec's for convergence for right-preconditioned methods. |
|-----------------|-----------------|-----------------|-----------------|
| Re.             | GMRES(k)        | DQGMRES(k)      |
| 5               | 10              | ∞               | 2               | 3               | 4               | 5               | 6               | 7               | 8               | 9               | 10              |
| 0               | 232             | 129             | 59              | 76              | 229             | 105             | 72              | 92              | †               | 160             | 83              | 75              |
| 1               | 218             | 126             | 69              | 76              | 276             | 131             | 81              | 94              | †               | 97              | 90              | 79              |
| 2               | 233             | 126             | 70              | 78              | 391             | 258             | 82              | 95              | †               | 98              | 94              | 78              |
| 3               | 208             | 126             | 71              | 87              | 346             | 232             | 85              | 95              | †               | 99              | 97              | 79              |
| 4               | 214             | 128             | 72              | 94              | 345             | †               | 88              | 95              | 477             | 108             | 98              | 86              |
| 5               | 210             | 128             | 72              | 192             | 304             | †               | 91              | 95              | 326             | 128             | 99              | 90              |
| 6               | 214             | 128             | 73              | 446             | 361             | †               | 94              | 97              | 258             | 197             | 100             | 94              |
| 7               | 215             | 129             | 73              | †               | 345             | †               | 97              | 99              | 239             | 229             | 101             | 96              |

The results in Table 3.1 show the irregular performance of DQGMRES(k) for
these small values of k when the preconditioned system is not symmetric. The per-
formance is entirely regular in Table 3.2, where the preconditioned system is near
symmetric. For Reynolds numbers up to 3, the systems are sufficiently symmetric so that DQGMRES(2) behaves the same as DQGMRES with much larger $k$. The performance remains regular until beyond Reynolds number 7, when the number of steps to convergence begins to become irregular, like in the right-preconditioned case.

GMRES with either right or symmetric preconditioning does not show any marked difference in performance; apparently the symmetry of the preconditioned system is not as essential here for this problem. However, the results do show that DQGMRES($k$) with small values of $k$ may perform as well, in terms of number of steps, as GMRES($k$) with large values of $k$, particularly if near-symmetry is preserved. Since the former is much more efficient, the combination of preserving symmetry and truncated iterative methods may result in a much more economical method, as well as the more regular behavior shown above.

We also performed the same experiments with orthogonal projection methods, namely the Full Orthogonalization Method (FOM) and its truncated variant, the
Direct Incomplete Orthogonalization Method (DIOM) [50]. The results were very similar to the results above, and are not shown here. Indeed, the development of the algorithms and the theory above is identical for these methods.

For interest, we also performed tests where an ILU(0) preconditioner was constructed for each matrix and compared right and split preconditioning. For the near-symmetric systems here, there was very little difference in these results compared to using IC(0) constructed from the $Re = 0$ case for all the matrices. Thus the deterioration in performance as the Reynolds number increases is not entirely due to a relatively less accurate preconditioner, but is more due to the increased nonsymmetry and non-normality of the matrices. Although the eigenvalues of the preconditioned matrices are identical, their eigenvectors and hence their degree of non-normality may change completely. Unfortunately, it is difficult to quantitatively relate non-normality and convergence.

3.4.2 Breakdown behavior of Bi-CG

To test the breakdown behavior of Bi-CG, MATLAB was used to generate random matrices of order 300 with approximately 50 percent normally distributed nonzero entries. The matrices were adjusted so that

$$A \leftarrow \frac{A + A^T}{2} - (\sigma_{\text{min}} + 10^{-5})I + \epsilon \frac{A - A^T}{2},$$

72
i.e., the symmetric part was shifted so that the lowest eigenvalue was $10^{-5}$ and then $\varepsilon$ times the skew-symmetric part was added back. The parameter $\varepsilon$ was altered to get varying degrees of nonsymmetry.

For each $\varepsilon$ that we tested, 100 matrices were generated, and the smallest value of the cosines corresponding to the denominators in the algorithms were recorded. In the right-preconditioned case, we recorded the minimum of

$$\frac{(AM^{-1}p_j^*, p_j^*)}{\|AM^{-1}p_j\|\|p_j^*\|} \quad \text{and} \quad \frac{(r_j, r_j^*)}{\|r_j\|\|r_j^*\|}$$

for all $j$, and for the symmetric right-preconditioned case, we recorded the minimum of

$$\frac{(Ap_j, p_j^*)}{\|Ap_j\|\|p_j^*\|} \quad \text{and} \quad \frac{(M^{-1}r_j, r_j^*)}{\|M^{-1}r_j\|\|r_j^*\|}$$

for all $j$. The relative residual norm reduction was $10^{-9}$ when the iterations were stopped. The initial guesses were 0, and $r_0^*$ was set to $r_0$. IC(0) of the symmetric part was used as the preconditioner.

Table 3.3 shows the frequencies of the size of minimum cosines for the right-preconditioned (first row of each pair of rows) and the symmetrically-preconditioned cases (second row of each pair of rows). For example, all 100 minimum cosines were between $10^{-3}$ and $3 \times 10^{-3}$ in the symmetrically-preconditioned case. The average number of Bi-CG steps and the average minimum cosine is also shown. The last column, labeled ‘better’, shows the number of times that the minimum cosine was higher in the improved algorithm.

The Table shows that the right-preconditioned algorithm can produce much
smaller cosines, indicating a greater probability for breakdown. The difference between the algorithms is less as the degree of nonsymmetry is increased. For $\varepsilon = 0.1$, there is almost no difference in the breakdown behavior of the algorithms.

The Table shows that the number of Bi-CG steps is not significantly reduced in the new algorithm, nor is the average minimum cosine of the modified algorithm significantly increased. It is the probability that a small cosine is not encountered that is better.

Table 3.3: Frequencies of minimum cosines for right-preconditioned (first row of each pair of rows) and symmetrically-preconditioned (second row of each pair of rows) Bi-CG.

<table>
<thead>
<tr>
<th>$\varepsilon$</th>
<th>steps</th>
<th>$3e-6$ to $1e-5$</th>
<th>$1e-5$ to $3e-5$</th>
<th>$3e-5$ to $1e-4$</th>
<th>$1e-4$ to $3e-4$</th>
<th>$3e-4$ to $1e-3$</th>
<th>$1e-3$ to $3e-3$</th>
<th>$1e-2$ to $3e-2$</th>
<th>$3e-2$ to $1e-1$</th>
<th>average x10^{-3}</th>
<th>better</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.000</td>
<td></td>
<td>32.51</td>
<td>31.77</td>
<td>1</td>
<td>4</td>
<td>9</td>
<td>19</td>
<td>26</td>
<td>30</td>
<td>11</td>
<td>1.87</td>
</tr>
<tr>
<td>0.005</td>
<td></td>
<td>30.57</td>
<td>29.97</td>
<td>1</td>
<td>2</td>
<td>8</td>
<td>16</td>
<td>34</td>
<td>34</td>
<td>5</td>
<td>3.51</td>
</tr>
<tr>
<td>0.010</td>
<td></td>
<td>29.27</td>
<td>28.94</td>
<td>1</td>
<td>0</td>
<td>3</td>
<td>2</td>
<td>10</td>
<td>29</td>
<td>32</td>
<td>20</td>
</tr>
<tr>
<td>0.050</td>
<td></td>
<td>27.53</td>
<td>27.32</td>
<td>1</td>
<td>3</td>
<td>8</td>
<td>13</td>
<td>36</td>
<td>31</td>
<td>8</td>
<td>4.15</td>
</tr>
<tr>
<td>0.100</td>
<td></td>
<td>26.38</td>
<td>26.42</td>
<td>1</td>
<td>11</td>
<td>18</td>
<td>39</td>
<td>27</td>
<td>4</td>
<td>11</td>
<td>0.88</td>
</tr>
</tbody>
</table>

It is important to note that this behavior only applies when $r_0^*$ is set to $r_0$.

When $r_0^*$ is chosen randomly, there is no gain in the symmetrically-preconditioned algorithm, as shown in Table 3.4.
Table 3.4: Frequencies of minimum cosines when $r_0$ is chosen randomly.

<table>
<thead>
<tr>
<th>$\varepsilon$</th>
<th>steps</th>
<th>3e-6 to 1e-5</th>
<th>1e-5 to 3e-5</th>
<th>3e-5 to 1e-4</th>
<th>3e-4 to 3e-3</th>
<th>3e-3 to 3e-2</th>
<th>3e-2 to 1e-1</th>
<th>average $\times 10^{-3}$</th>
<th>better</th>
</tr>
</thead>
<tbody>
<tr>
<td>0.000</td>
<td>33.05</td>
<td>1</td>
<td>4</td>
<td>11</td>
<td>24</td>
<td>26</td>
<td>30</td>
<td>4</td>
<td>0.92</td>
</tr>
<tr>
<td>32.54</td>
<td>1</td>
<td>1</td>
<td>3</td>
<td>11</td>
<td>24</td>
<td>56</td>
<td>4</td>
<td>1.31</td>
<td></td>
</tr>
</tbody>
</table>

Table 3.5: Steps and minimum cosines for the driven cavity problem.

<table>
<thead>
<tr>
<th>$Re$</th>
<th>Bi-CG steps</th>
<th>min cosines</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>right</td>
<td>symm</td>
</tr>
<tr>
<td>0</td>
<td>70</td>
<td>62</td>
</tr>
<tr>
<td>1</td>
<td>71</td>
<td>68</td>
</tr>
<tr>
<td>2</td>
<td>74</td>
<td>72</td>
</tr>
<tr>
<td>3</td>
<td>73</td>
<td>72</td>
</tr>
<tr>
<td>4</td>
<td>77</td>
<td>72</td>
</tr>
<tr>
<td>5</td>
<td>80</td>
<td>75</td>
</tr>
<tr>
<td>6</td>
<td>80</td>
<td>78</td>
</tr>
<tr>
<td>7</td>
<td>80</td>
<td>80</td>
</tr>
</tbody>
</table>

Table 3.5 shows the number of steps and the minimum cosines for the two algorithms applied to the driven cavity problem described in Section 5.1 above. Figure 3.1 shows a plot of the minimum cosines as the two algorithms progress for the $Re = 1$ problem. Note that the minimum cosines are higher and much smoother in the symmetrically-preconditioned case. In the $Re = 7$ problem, the cosines are still higher, but the smoothness is lost.
Figure 3.1: Minimum cosines in right-preconditioned Bi-CG (solid line) and symmetrically-preconditioned Bi-CG (dashed line) for the $Re = 1$ problem.

3.5 Conclusions

When solving linear systems with matrices that are very close to being symmetric, this paper has shown that it is possible to improve upon the standard practice of using a (nonsymmetric) preconditioner for that matrix along with a left- or right-preconditioned iterative method. The original degree of symmetry may be maintained by using a symmetric preconditioner and an alternate inner product (or split preconditioning, if appropriate). By combining this idea with truncated iterative methods, solution procedures that converge more quickly and require less storage are developed. The truncated methods also seem to become more robust with the truncation parameter $k$ when near-symmetry is maintained. The Bi-CG algorithm also seems to be more robust with respect to serious breakdown when near-symmetry is maintained.
CHAPTER 4

ML($k$)BiCGSTAB: A BiCGSTAB Variant Based on Multiple Lanczos Starting Vectors

We propose a new Krylov subspace solver, which we name ML($k$)BiCGSTAB, for solving the nonsymmetric linear system

$$Ax = b$$

in this chapter. The derivation of the new method is similar to that of BiCGSTAB but based on multiple Lanczos process. ML($k$)BiCGSTAB can be regarded as a bridge between BiCGSTAB and FOM. In fact, both BiCGSTAB and FOM are special cases of ML($k$)BiCGSTAB where $k = 1$ and 2 respectively. In §1, we give our version of the multiple Lanczos process with one right starting vector and $k$ left starting vectors. In §2, a BiCG-like method is derived from the Lanczos method in §1. §3 contains the main derivation for the ML($k$)BiCGSTAB method. Numerical results are given in §4.
4.1 A Lanczos method for \( k \) Left Starting Linear Independent Vectors

Let \( A \) be an \( n \times n \) real matrix and given \( k + 1 \) real vectors \( v_0, q_1, q_2, \cdots, q_k \).

We define

\[
p_{j+k+i} = A^{Tj}q_i
\]  

(4.1)

for \( i = 1, 2, \cdots, k; j = 0, 1, 2, \cdots, \), and suppose the matrices

\[
W_i = \begin{bmatrix}
p_i^T v_0 & p_i^T A v_0 & \cdots & p_i^T A^{l-1} v_0 \\
p_2^T v_0 & p_2^T A v_0 & \cdots & p_2^T A^{l-1} v_0 \\
\vdots & \vdots & & \vdots \\
p_\nu^T v_0 & p_\nu^T A v_0 & \cdots & p_\nu^T A^{l-1} v_0
\end{bmatrix}, \quad l = 1, 2, \cdots, \nu
\]

are nonsingular, where \( \nu \) is the grade of \( v_0 \) with respect to \( A \), i.e., the degree of the minimal polynomial of \( v_0 \) with respect to \( A \).

We now consider a sequence \( \{v_l\}_{l=0,1,\cdots,\nu} \) of vectors from the Krylov space

\[
K(v_0, A) = \text{span}\{v_0, Av_0, A^2v_0, \cdots\}
\]

with the properties

\[
v_l \in A^l v_0 + K_l(v_0, A) \equiv A^l v_0 + \text{span}\{v_0, Av_0, \cdots, A^{l-1} v_0\}
\]

(4.2)

and

\[
v_l \perp \text{span}\{p_1, p_2, \cdots, p_l\}.
\]

(4.3)
The existence and uniqueness of such a sequence is guaranteed by the nonsingularity of the $W_l$'s. In fact, if we express $v_l$ as

$$v_l = A^l v_0 + \gamma_0^{(l)} v_0 + \gamma_1^{(l)} A v_0 + \cdots + \gamma_{l-1}^{(l)} A^{l-1} v_0,$$  

then (4.3) is equivalent to $W_l \gamma^{(l)} = -f$ where $\gamma^{(l)} = [\gamma_0^{(l)}, \gamma_1^{(l)}, \cdots, \gamma_{l-1}^{(l)}]^T$ and $f = [p^T_l A^l v_0, \cdots, p^T_l A v_0]^T$.

Two simple facts can be derived for the sequence $\{v_l\}_{l=0,1,\ldots\nu}$: (a) $v_\nu = 0$ and (b) $v_l \not\in p_{l+1}$ whenever $l < \nu$. To see (a), we note that $\nu$ is the grade of $v_0$ and hence $A^\nu v_0 \in K_\nu(v_0, A)$. Property (4.2) is now reduced to $v_\nu \in K_\nu(v_0, A)$ and thus (a) follows by the uniqueness of $v_\nu$. To prove (b), we assume that $v_l \perp p_{l+1}$. Then combining with property (4.3), we have $v_l \perp \text{span}\{p_1, \cdots, p_{l+1}\}$, and hence $W_{l+1} \tilde{\gamma}^{(l)} = 0$, where $\tilde{\gamma}^{(l)} = [\gamma_0^{(l)}, \cdots, \gamma_{l-1}^{(l)}, 1]^T$ and $\gamma_i^{(l)}$ are defined in (4.4), in contradiction with the nonsingularity of $W_{l+1}$.

Applying (4.4) to itself recursively, we can represent $v_l$ in terms of its previous $v_0, \cdots, v_{l-1}$ as follows,

$$v_l = A v_{l-1} + h_{l-1}^{(l-1)} v_{l-1} + h_{l-2}^{(l-1)} v_{l-2} + \cdots + h_0^{(l-1)} v_0.$$  

Noting that $v_l \perp \text{span}\{p_1, \cdots, p_l\}$, $v_l \not\in p_{l+1}$ and $A^T p_i = p_{k+i}$, and examining in turn

$$p^T_i v_l = p^T_i A v_{l-1} + p^T_i h_{l-1}^{(l-1)} v_{l-1} + p^T_i h_{l-2}^{(l-1)} v_{l-2} + \cdots + p^T_i h_0^{(l-1)} v_0$$

for $i = 1, 2, \cdots, l-k-1$, we find all the coefficients zero except $h_{l-1}^{(l-1)}, h_{l-2}^{(l-1)}, \cdots, h_{m_l}^{(l-1)}$ where $m_l = \max(l - k - 1, 0)$. Thus we obtain a $k+2$ term recursion relationship.
for \( \{v_l\}_{l=1,\ldots,\nu} \) of the form,

\[
v_l = A v_{l-1} + h_{l-1}^{(l-1)} v_{l-1} + h_{l-2}^{(l-1)} v_{l-2} + \cdots + h_{m_l}^{(l-1)} v_{m_l}.
\]

If we set \( V_\nu = [v_0, v_1, \ldots, v_{\nu-1}] \) and \( H_\nu = (h_{ij})_{i,j=1,\ldots,\nu} \), the \( \nu \times \nu \) Hessenberg matrix with \( h_{j+1,j} = 1; h_{ij} = -h_{l-1}^{(j-1)} \) for \( m_j + 1 \leq i \leq j \); \( h_{ij} = 0 \) otherwise, and if we recall that \( v_\nu = 0 \), then the recurrence relations (4.5) can be written in matrix form as

\[
A V_\nu = V_\nu H_\nu.
\]

Moreover, set \( P_\nu = [p_1, p_2, \ldots, p_\nu] \) and apply \( P_\nu^T \) to (4.6) from the left,

\[
P_\nu^T A V_\nu = P_\nu^T V_\nu H_\nu.
\]

Because of (4.3) and the fact that \( v_l \not\in p_{l+1} \), \( P_\nu^T V_\nu \) is a lower triangular matrix with all the entries nonzero in its diagonal. Comparing the corresponding principal blocks of both sides of (4.7), we obtain a condition, which guarantees the nonsingularity of \( H_\nu \), that if the matrices

\[
S_l = \begin{bmatrix}
p_1^T A v_0 & p_1^T A^2 v_0 & \cdots & p_1^T A^l v_0 \\
p_2^T A v_0 & p_2^T A^2 v_0 & \cdots & p_2^T A^l v_0 \\
\vdots & \vdots & \ddots & \vdots \\
p_l^T A v_0 & p_l^T A^2 v_0 & \cdots & p_l^T A^l v_0
\end{bmatrix}, \quad l = 1, 2, \ldots, \nu
\]

are nonsingular, so are the principal blocks of \( H_\nu \).

Finally, we can see from (4.2) that

\[
K_{l+1}(v_0, A) = \text{span}\{v_0, v_1, \ldots, v_l\}, \quad l = 0, 1, \ldots, \nu - 1.
\]
Since the dimension of $K_\nu(v_0, A)$ is $\nu$, the vectors $\{v_l\}_{l=0,1,\ldots,\nu-1}$ are linear independent.

Summing up the above discussions, we conclude that

**Theorem 4.1** Given $v_0, q_1, q_2, \cdots q_k$ and $\nu$, $p_i$'s, $W_\nu$ and $S_\nu$ as defined above. If the principal submatrices of $W_\nu$ are nonsingular, there exist an $n \times \nu$ matrix $V_\nu = [v_0, v_1, \cdots, v_{\nu-1}]$ of rank $\nu$, whose first column is $v_0$, and an $\nu \times \nu$ Hessenberg matrix $H_\nu$, which has upper bandwidth $k$ and all the entries $h_{j+1,j} = 1$ in its lower subdiagonal, such that

$$v_l \perp \text{span}\{p_1, p_2, \cdots, p_l\}, \quad v_l \not\in p_{l+1}, \quad l = 0, 1, \cdots, \nu - 1,$$

and

$$AV_\nu = V_\nu H_\nu. \quad (4.9)$$

Such $V_\nu$ and $H_\nu$ are unique. Furthermore, if the principal submatrices of $S_\nu$ are nonsingular, so are the principal submatrices of $H_\nu$.

In most of cases, the conditions of Theorem 4.1 are satisfied when the $q_i$'s are chosen randomly from a continuous distribution.

Our procedure for generating the vectors $v_i$'s and $p_i$'s are closely related to the multiple starting vector Lanczos method in [1]. In fact, the columns of $V_\nu$ are exactly the same as the right Lanczos vectors in [1]. However, the $p_i$'s are different from the left Lanczos vectors in [1]; in fact they are not orthogonal to the
$v_i$'s. It turns out we do not need the full bi-orthogonality property in deriving our extensions of BiCG and BiCGSTAB.

4.2 ML($k$)BiCG: A BiCG Variant Based on Multiple Starting Vectors

Lanczos

We now turn to the linear system

$$Ax = b$$

(4.10)

and we shall derive, based on Theorem 1, a projection method by borrowing the techniques used in the derivation of the Conjugate Gradient algorithm from the symmetric Lanczos procedure [p.176, 12]. Even though this is a well known procedure, in our case there is some difference from the standard case and therefore we include the derivation here for completeness and clarity.

Suppose an initial guess $x_0$ to (4.10) is given. We set in Theorem 1 $v_0 = b - Ax_0$. At the $l$th step of our projection method, we seek an approximate solution $x_l$ with

$$x_l \in x_0 + \text{span}\{v_0, v_1, \ldots, v_{l-1}\}$$

(4.11)

and

$$r_l \equiv b - Ax_l \perp \text{span}\{p_1, p_2, \ldots, p_l\},$$

(4.12)

\footnote{We do not assume the matrix $A$ is nonsingular throughout the paper unless where specified. We just require that the assumptions in Theorem 4.1 hold.}
where the $v_i$'s and $p_i$'s are as defined in Theorem 4.1.

Since $P_l^T V_l$, a lower triangular matrix with all diagonal entries nonzero, is nonsingular, it can be shown in the same manner as in deriving CG from the Lanczos basis that $x_l$ is uniquely determined by conditions (4.11) and (4.12) and has the following expression,

$$x_l = x_0 + V_l H_l^{-1} e_1,$$

(4.13)

with the corresponding residual $r_l$ is in the same direction as $v_l$, where $P_l = [p_1, \ldots, p_l]$, $V_l = [v_0, \ldots, v_{l-1}]$, $H_l$ is the $l \times l$ principal submatrix of $H_\nu$, and $e_1$ is the first column of the $l \times l$ identity matrix. From (4.11) and (4.8), we have $r_l \in v_0 - \text{span} \{Av_0, A^2v_0, \ldots, A^lv_0\}$ and hence $r_l \neq 0$ if $l < \nu$ since $\nu$ is the grade of $v_0$. Moreover, $r_\nu = 0$ from (4.13) and (4.9).

Letting $r_l = \xi_l v_l$ for some scalar $\xi_l$ which is not zero whenever $i < \nu$ and setting $\Lambda_l = \text{diag} \{\xi_0, \xi_1, \ldots, \xi_{l-1}\}$, then (4.13) can be rewritten as

$$x_l = x_0 + R_l \Lambda_l^{-1} H_l^{-1} e_1,$$

where $R_l = [r_0, r_1, \ldots, r_{l-1}]$. Write the $LDU$ decomposition of $H_l \Lambda_l$ as

$$H_l \Lambda_l = L_l D_l U_l,$$

which exists due to the nonsingularities of the principal submatrices of $H_l \Lambda_l$, and define

$$G_l \equiv [g_0, g_1, \ldots, g_{l-1}] = R_l U_l^{-1}, \quad z_l = D_l^{-1} L_l^{-1} e_1.$$
Following again the derivation for CG, we have

\[
\begin{bmatrix}
 z_{l-1} \\
 \alpha_l
\end{bmatrix}
\]

for some \( \alpha_l \). As a result, \( x_l \) can be updated as

\[
x_l = x_0 + G_l z_l = x_0 + G_{l-1} z_{l-1} + \alpha_l g_{l-1} = x_{l-1} + \alpha_l g_{l-1}
\]

(4.14)

and hence

\[
r_l = r_{l-1} - \alpha_l A g_{l-1}.
\]

(4.15)

On the other hand, since \( G_{i+1} = R_{i+1} U_{i+1}^{-1} \), \( g_i \) can be computed from the previous \( g_i \)'s and \( r_l \) by the update

\[
g_i = r_l + \beta_i^{(i)} g_{i-1} + \beta_i^{(i)} g_{i-2} + \cdots + \beta_i^{(i)} g_{\tilde{m}_i} u_{\tilde{m}_i}
\]

(4.16)

where \( \tilde{m}_i = \max(l - k, 0) \) and where \( -\beta_i^{(i)} \) are the nonzero entries of the last column of \( U_{i+1} \).

To compute the coefficients \( \alpha_l \) and \( \beta_i^{(i)} \) in (4.14), (4.15) and (4.16), we need the \( A \)-orthogonality of the vectors \( g_i \)'s and \( p_i \)'s. Since

\[
P_i^T A g_i = P_i^T A R_i U_i^{-1} = P_i^T A V_i A U_i^{-1} = P_i^T (V_i H_i + v_i e_i^T) A U_i^{-1}
\]

\[
= P_i^T V_i H_i A U_i^{-1} = P_i^T V_i L_i D_i U_i^{-1} = P_i^T V_i L_i D_i,
\]

where \( e_i \) denotes the last column of the \( l \times l \) identity matrix, and since \( P_i^T V_i \) is nonsingular and lower triangular, we have

\[
p_i^T A g_j = 0, \quad i < j
\]

(4.17)
\[ p_i^T Ag_{i-1} \neq 0. \]

Thus, utilizing this information, we examine the following equations derived from (4.15),
\[ p_i^T r_i = p_i^T r_{i-1} - \alpha_i p_i^T Ag_{i-1} \]
and
\[ p_i^T Ag_i = p_i^T Ar_i + \beta_{i-1}^{(l)} p_i^T Ag_{i-1} + \beta_{i-2}^{(l)} p_i^T Ag_{i-2} + \cdots + \beta_{\tilde{m}_i}^{(l)} p_i^T Ag_{\tilde{m}_i} \]
for \( i \) going from \( \tilde{m}_i + 1 \) to \( l \). Then we get
\[ \alpha_i = \frac{p_i^T r_{i-1}}{p_i^T Ag_{i-1}} \] (4.18)
and
\[ \beta_{i-1}^{(l)} = -\frac{p_i^T Ar_i + p_i^T \sum_{j=\tilde{m}_i}^{i-2} \beta_j^{(l)} Ag_j}{p_i^T Ag_{i-1}}, \quad i = \tilde{m}_i + 1, \ldots, l. \] (4.19)

Now, putting the relations (4.14), (4.15), (4.16), (4.18) and (4.19) together, we have

**Algorithm 4.1. ML(k)BiCG**

1. Choose an initial guess \( x_0 \) and \( k \) vectors \( q_1, q_2, \ldots, q_k \).
2. Compute \( r_0 = b - Ax_0 \) and set \( p_1 = q_1, g_0 = r_0 \).
3. For \( l = 1, 2, \ldots, \) until convergence:
   4. \( \alpha_l = \frac{p_l^T r_{l-1}}{p_l^T Ag_{l-1}}; \)
   5. \( x_l = x_{l-1} + \alpha_l g_{l-1}; \)
   6. \( r_l = r_{l-1} - \alpha_l Ag_{l-1}; \)
   7. For \( s = \max(l-k, 0), \ldots, l-1 \)
8. \[ \beta_i^{(l)} = -p^{T}_{s+1}A \left( r_l + \sum_{l-1}^{s} \beta_i^{(l)} g_s \right) / p^{T}_{s+1} A g_s; \]

9. \textit{End}

10. \[ g_s = r_l + \sum_{l=\max(l-k,0)}^{l-1} \beta_i^{(l)} g_s; \]

11. \textit{Compute} \( p_{l+1} \text{ according to } (4.1) \)

12. \textit{End}

It is worthwhile to remark on two special cases where \( k = 1 \) and \( k \geq \nu \). If \( k = 1 \), then \( p_l = A^{p_{l-1}} q_1 \) and conditions (4.11), (4.12) become

\[ x_l \in x_0 + \text{span}\{v_0, v_1, \cdots, v_{l-1}\}, \quad r_l \perp K_l(q_1, A^T), \]

which are exactly what the BiCG approximate solution \( x_l^{\text{BiCG}} \) needs to satisfy. As a result, Algorithm 4.1 is equivalent to the BiCG algorithm mathematically. On the other hand, when \( k \geq \nu \), (4.11) and (4.12) reduce to

\[ x_l \in x_0 + \text{span}\{v_0, v_1, \cdots, v_{l-1}\}, \quad r_l \perp \text{span}\{q_1, q_2, \cdots, q_l\} \]

for \( 1 \leq l \leq \nu \). If, at the \( l \)th step of the computations, we choose \( (p_{l+1} =) q_{l+1} = r_l \) while setting \( (p_1 =) q_1 = r_0 \) beforehand, then Algorithm 4.1 is mathematically equivalent to the FOM algorithm.

From its derivation, we can state the following result about Algorithm 4.1.

\textbf{Theorem 4.2} Under the assumptions of Theorem 4.1, ML(\( k \)) BiCG does not break down by zero division before step \( \nu \) and the approximate solution \( x_{\nu} \) at step \( \nu \) is exact to the system (4.10).
4.3 ML(\(k\))BiCGSTAB: A BiCGSTAB Variant Based on Multiple Starting Vectors Lanczos

The implementation of Algorithm 4.1 requires the use of \(A^T\) to compute \(p_{l+1}\) in Line 11. In practice, however, the transpose of \(A\) is not always available, for instance if the matrix is not formed explicitly and the matvec product is only given as an operator. But this difficulty can be overcome by adopting the techniques in the derivations of CGS and BiCGSTAB. In this section, we give a transpose-free version of Algorithm 4.1 which we call ML(\(k\))BiCGSTAB.

We first rearrange the outer for loop of Algorithm 4.1 into a form more convenient for our development. Let \(l = jk + i\) and let the index \(i\) vary from 1 to \(k\) and \(j\) starting with 0. Then we convert the loop in \(l\) into doubly nested loops in \(j\) and \(i\) respectively. By moving the case where \(i = 1\) outside the \(i\)-loop, we rewrite the \(l\)-loop (omitting Lines 5 and 11) of Algorithm 4.1 as,

1. For \(j = 0, 1, 2, \cdots\)

2. \[\alpha_{jk+1} = p_{jk+1}^T r_{(j-1)k+k} / p_{jk+1}^T A g_{(j-1)k+k};\]

3. \[r_{jk+1} = r_{(j-1)k+k} - \alpha_{jk+1} A g_{(j-1)k+k};\]

4. For \(i = 1, 2, \cdots, k\)

5. For \(s = \max((j-1)k + i, 0), \cdots, jk + i - 1\)

6. \[\beta_{s}^{(jk+i)} = -p_{s+1}^T A \left( r_{jk+i} + \sum_{t=\max((j-1)k+i,0)}^{s-1} \beta_{t}^{(jk+i)} g_{t} \right) / p_{s+1}^T A g_{s};\]
in which Lines 5-8 can be again expanded into:

1. For $s = \max((j-1)k+i,0), \cdots, (j-1)k+k-1$

2. $\beta_s^{(j_k+i)} = -p_{s+1}^T A \left( r_{j_k+i} + \sum_{t=\max((j-1)k+i,0)}^{s-1} \beta_t^{(j_k+i)} g_t \right) / p_{s+1}^T A g_s$;

3. End

4. $\beta_{(j-1)k+k}^{(j_k+i)} = -p_{j_k+1}^T A \left( r_{j_k+i} + \sum_{t=\max((j-1)k+i,0)}^{(j-1)k+k-1} \beta_t^{(j_k+i)} g_t \right) / p_{j_k+1}^T A g_{(j-1)k+k}$;

5. For $s = j_k+1, \cdots, j_k+i-1$

   $\beta_s^{(j_k+i)} = -p_{s+1}^T A \left( r_{j_k+i} + \sum_{t=\max((j-1)k+i,0)}^{(j-1)k+k-1} \beta_t^{(j_k+i)} g_t + \sum_{t=j_k+1}^{s-1} \beta_t^{(j_k+i)} g_t \right) / p_{s+1}^T A g_s$;
7. End

8. \[ g_{jk+i} = r_{jk+i} + \sum_{s=\max((j-1)k+i,0)}^{(j-1)k+k} \beta_{s}^{(jk+i)} g_s + \sum_{s=jk+1}^{jk+i-1} \beta_{s}^{(jk+i)} g_s; \]

and further into:

1. If \( j = 0 \)

2. \[ \beta_0^{(i)} = -p_1^T Ar_i / p_1^T Ag_0; \]

3. For \( s = 1, \ldots, i-1 \)

4. \[ \beta_s^{(i)} = -p_{s+1}^T A \left( r_i + \beta_0^{(i)} g_0 + \sum_{t=1}^{s-1} \beta_t^{(i)} g_t \right) / p_{s+1}^T A g_s; \]

5. End

6. \[ g_i = r_i + \sum_{s=0}^{i-1} \beta_s^{(i)} g_s; \]

7. Else

8. For \( s = i, \ldots, k-1 \)

9. \[ \beta_{(j-1)k+s}^{(jk+i)} = -p_{(j-1)k+s+1}^T A \left( r_{jk+i} + \sum_{t=i}^{s-1} \beta_{(j-1)k+t}^{(jk+i)} g_{(j-1)k+t} \right) / p_{(j-1)k+s+1}^T A g_{(j-1)k+s}; \]

10. End

11. \[ \beta_{(j-1)k+k}^{(jk+i)} = -p_{jk+1}^T A \left( r_{jk+i} + \sum_{t=i}^{k-1} \beta_{(j-1)k+t}^{(jk+i)} g_{(j-1)k+t} \right) / p_{jk+1}^T A g_{(j-1)k+k}; \]
For $s = 1, \cdots, i-1$

$$\rho^{(j,k+i)}_{jk+s} = -p^T_{jk+s+1} A \left( r_{jk+i} + \sum_{t=i}^{k} \rho^{(j,k+i)}_{(j-1)k+t} g_{(j-1)k+t} + \sum_{t=1}^{s-1} \rho^{(j,k+i)}_{jk+t} g_{jk+t} \right) / p^T_{jk+s+1} A g_{jk+s};$$

End

$$g_{jk+i} = r_{jk+i} + \sum_{s=1}^{k} \rho^{(j,k+i)}_{(j-1)k+s} g_{(j-1)k+s} + \sum_{s=1}^{i-1} \rho^{(j,k+i)}_{jk+s} g_{jk+s};$$

End

We now introduce an auxiliary polynomial $\psi_l(\lambda)$ defined by the recurrence relation,

$$\psi_0(\lambda) = 1,$$

$$\psi_l(\lambda) = (\rho_l \lambda + 1) \psi_{l-1}(\lambda), \quad l = 1, 2, \cdots,$$

where $\rho_l$ is a free parameter. This polynomial $\psi_l(\lambda)$ was first used by van der Vorst in the derivation of BiCGSTAB [60]. If we express $\psi_l(\lambda)$ in terms of the power basis,

$$\psi_l(\lambda) = \eta_l^{(l)} \lambda^l + \cdots + \eta_1^{(l)} \lambda + \eta_0^{(l)},$$

then it is clear that $\eta_l^{(l)} = \rho_1 \rho_2 \cdots \rho_l$ and $\eta_0^{(l)} = 1$.

Next, we define the following vectors, analogous to those defined in BiCGSTAB,

$$\tilde{\pi}_{jk+i} = \psi_j(A)r_{jk+i}, \quad \pi_{jk+i} = \psi_{j+1}(A)r_{jk+i},$$

90
\[ \tilde{\omega}_{jk+i} = \psi_j(A)g_{jk+i}, \quad \omega_{jk+i} = \psi_{j+1}(A)g_{jk+i}, \]

for \( i = 1, 2, \ldots, k; j = 0, 1, \ldots \), and set \( \pi_0 = \omega_0 = r_0 (= g_0) \). Our goal is to define \( \pi_{jk+i} \) to be the residual of our new method. The other three vectors are needed in deriving a recurrence for \( \pi_{jk+i} \). By recalling (4.1), (4.12) and (4.17), we find that the scalars \( \alpha_i \)'s and \( \beta_i \)'s in Algorithm 4.1 can be computed via these new vectors. The derivations are quite complicated and therefore we give the details in the Appendix while summarizing only the results here. In fact, we have

\[ \alpha_{jk+1} = \frac{q_1^T \pi_{(j-1)k+k}}{q_1^T A \omega_{(j-1)k+k}} \]

for \( 0 \leq j, i = 1 \);

\[ \alpha_{jk+i+1} = \frac{\rho_{j+1} q_i^T \tilde{\pi}_{jk+i}}{q_{i+1}^T (\omega_{jk+i} - \tilde{\omega}_{jk+i})} \]

for \( 0 \leq j, 1 \leq i < k \);

\[ \beta_0^{(i)} = -\frac{q_1^T \pi_i}{\rho_1 q_1^T A \omega_0} \]

for \( 1 \leq i \leq k, j = 0 \);

\[ \beta_s^{(i)} = -\frac{q_{s+1}^T (\pi_i + \beta_0^{(i)} \rho_1 A \omega_0 + \sum_{t=1}^{s-1} \beta_t^{(i)} (\omega_t - \tilde{\omega}_i))}{q_{s+1}^T (\omega_s - \tilde{\omega}_s)} \]

for \( 1 \leq s < i \leq k, j = 0 \);

\[ \beta_{(j-1)k+s}^{(j+1)} = -\frac{q_{s+1}^T (\pi_{jk+i} + \sum_{t=i}^{s-1} \beta_{(j-1)k+t}^{(j+1)} (\omega_{(j-1)k+t} - \tilde{\omega}_{(j-1)k+t}))}{q_{s+1}^T (\omega_{(j-1)k+s} - \tilde{\omega}_{(j-1)k+s})} \]

for \( 1 \leq j, 1 \leq i \leq s \leq k - 1 \);

\[ \beta_{(j-1)k+k}^{(j+1)} = -\frac{q_1^T (\pi_{jk+i} + \rho_{j+1} \sum_{t=i}^{k-1} \beta_{(j-1)k+t}^{(j+1)} A \omega_{(j-1)k+t})}{\rho_{j+1} q_1^T A \omega_{(j-1)k+k}} \]
for \( 1 \leq j, 1 \leq i \leq k; \)

\[
\beta_{jk+i}^{(j k + i)} = -q_s^T (\pi_{jk+i} + \rho_{jk+i} \sum_{t=1}^{s} \beta_{(j-1)k+t}^{(j k + i)} A \omega_{(j-1)k+t} + \sum_{t=1}^{s-1} \beta_{jk+t}^{(j k + i)} (\omega_{jk+t} - \tilde{\omega}_{jk+t})) / q_s^T (\omega_{jk+s} - \tilde{\omega}_{jk+s})
\]

for \( 1 \leq j, 1 \leq s < i \leq k. \)

The vectors \( \bar{\pi}_{jk+i}, \pi_{jk+i}, \tilde{\omega}_{jk+i} \) and \( \omega_{jk+i} \) themselves can be updated according to Algorithm 4.1 as follows,

\[
\bar{\pi}_{jk+1} = \psi_j(A) r_{jk+1}
\]

\[
= \psi_j(A)(r_{(j-1)k+k} - \alpha_{jk+1} A g_{(j-1)k+k})
\]

\[
= \pi_{(j-1)k+k} - \alpha_{jk+1} A \omega_{(j-1)k+k}
\]

for \( 0 \leq j, i = 1; \)

\[
\pi_{jk+1} = \psi_{j+1}(A) r_{jk+1}
\]

\[
= (\rho_{j+1} A \psi_j(A) + \psi_j(A)) r_{jk+1}
\]

\[
= \rho_{j+1} A \bar{\pi}_{jk+1} + \bar{\pi}_{jk+1}
\]

for \( 0 \leq j, i = 1; \)

\[
\bar{\pi}_{jk+i+1} = \psi_j(A) r_{jk+i+1}
\]

\[
= \psi_j(A)(r_{jk+i} - \alpha_{jk+i+1} A g_{jk+i})
\]

92
\[\begin{align*}
&= \psi_j(A)r_{jk+i} - \frac{\alpha_{jk+i+1}}{\rho_{j+1}}(\psi_{j+1}(A) - \psi_j(A))g_{jk+i} \\
&= \tilde{\pi}_{jk+i} - \frac{\alpha_{jk+i+1}}{\rho_{j+1}}(\omega_{jk+i} - \tilde{\omega}_{jk+i})
\end{align*}\]

for \(1 \leq i < k, 0 \leq j;\)

\[\pi_{jk+i+1} = \psi_{j+1}(A)r_{jk+i+1}\]

\[\begin{align*}
&= \psi_{j+1}(A)(r_{jk+i} - \alpha_{jk+i+1}Ag_{jk+i}) \\
&= \pi_{jk+i} - \alpha_{jk+i+1}A\omega_{jk+i}
\end{align*}\]

for \(1 \leq i < k, 0 \leq j;\)

\[\tilde{\omega}_i = \psi_0(A)g_i\]

\[\begin{align*}
&= \psi_0(A)\left(r_i + \sum_{s=0}^{i-1} \beta_s^{(i)}g_s\right) \\
&= \tilde{\pi}_i + \beta_0^{(i)}\omega_0 + \sum_{s=1}^{i-1} \beta_s^{(i)}\tilde{\omega}_s
\end{align*}\]

for \(1 \leq i \leq k, j = 0;\)

\[\omega_i = \psi_1(A)g_i\]

\[\begin{align*}
&= \psi_1(A)\left(r_i + \sum_{s=0}^{i-1} \beta_s^{(i)}g_s\right) \\
&= \psi_1(A)r_i + \beta_0^{(i)}\psi_1(A)g_0 + \sum_{s=1}^{i-1} \beta_s^{(i)}\psi_1(A)g_s
\end{align*}\]
\[ \begin{align*}
&= \pi_i + \beta_0^{(i)} (\rho_1 A \omega_0 + \omega_0) + \sum_{s=1}^{i-1} \beta_s^{(i)} \omega_s \\
\text{for } 1 \leq i \leq k, j = 0; \\
\tilde{\omega}_{jk+i} &= \psi_j(A) g_{jk+i} \\
&= \psi_j(A) \left( r_{jk+i} + \sum_{s=1}^k \beta_{(j-1)k+s} g_{(j-1)k+s} + \sum_{s=1}^{i-1} \beta_{jk+s}^{(j+i)} g_{jk+s} \right) \\
&= \tilde{\pi}_{jk+i} + \sum_{s=1}^k \beta_{(j-1)k+s} \omega_{(j-1)k+s} + \sum_{s=1}^{i-1} \beta_{jk+s}^{(j+i)} \tilde{\omega}_{jk+s} \\
\text{for } 1 \leq j, 1 \leq i \leq k; \\
\omega_{jk+i} &= \psi_{j+1}(A) g_{jk+i} \\
&= \psi_{j+1}(A) \left( r_{jk+i} + \sum_{s=1}^k \beta_{(j-1)k+s} g_{(j-1)k+s} + \sum_{s=1}^{i-1} \beta_{jk+s}^{(j+i)} g_{jk+s} \right) \\
&= \psi_{j+1}(A) r_{jk+i} + \sum_{s=1}^k \beta_{(j-1)k+s} (\rho_{j+1} A + I) \psi_{j}(A) g_{(j-1)k+s} + \\
&\quad \sum_{s=1}^{i-1} \beta_{jk+s}^{(j+i)} \psi_{j+1}(A) g_{jk+s} \\
&= \pi_{jk+i} + \sum_{s=1}^k \beta_{(j-1)k+s} (\rho_{j+1} A \omega_{(j-1)k+s} + \omega_{(j-1)k+s}) + \sum_{s=1}^{i-1} \beta_{jk+s}^{(j+i)} \omega_{jk+s} \\
\text{for } 1 \leq j, 1 \leq i \leq k, \text{ where } I \text{ is the } n \times n \text{ identity matrix.}
\end{align*} \]

The formulas we have just derived constitute the main operations of the ML(k)BiCGSTAB algorithm, which we summarize as follows. 

94
1. Set $\pi_0 = \omega_0 = r_0$.

2. For $j = 0, 1, 2, \cdots$

3. $\alpha_{jk+1} = \frac{q^T \pi_{(j-1)k+k}}{q_1^T A \omega_{(j-1)k+k}}$;

4. $\tilde{\pi}_{jk+1} = \pi_{(j-1)k+k} - \alpha_{jk+1} A \omega_{(j-1)k+k}$;

5. Choose $\rho_{j+1} \neq 0$

6. $\pi_{jk+1} = \rho_{j+1} A \tilde{\pi}_{jk+1} + \tilde{\pi}_{jk+1}$;

7. For $i = 1, 2, \cdots, k$

8. If $j = 0$

9. $\beta_0^{(i)} = -\frac{q^T \pi_i}{q_1^T A \omega_0}$;

10. For $s = 1, \cdots, i - 1$

11. $\beta_s^{(i)} = -\frac{q^T \left( \pi_i + \beta_0^{(i)} \rho_1 A \omega_0 + \sum_{t=1}^{s-1} \beta_t^{(i)} (\omega_i - \tilde{\omega}_i) \right)}{q^T_{s+1} (\omega_s - \tilde{\omega}_s)}$;

12. End

13. $\tilde{\omega}_i = \tilde{\pi}_i + \beta_0^{(i)} \omega_0 + \sum_{s=1}^{i-1} \beta_s^{(i)} \tilde{\omega}_s$;

14. $\omega_i = \pi_i + \beta_0^{(i)} (\rho_1 A \omega_0 + \omega_0) + \sum_{s=1}^{i-1} \beta_s^{(i)} \omega_s$;

15. Else

16. For $s = i, \cdots, k - 1$
17. \[ \beta_{(j,k+i)}^{(j,k+i)} = -q_{s+1}^T \left( \pi_{j,k+i} + \sum_{t=1}^{s} \beta_{j,k+t}^{(j,k+i)} (\omega_{j,k+t} - \bar{\omega}_{j,k+t}) \right) q_{s+1}^T (\omega_{j,k+s} - \bar{\omega}_{j,k+s}); \]

18. **End**

19. \[ \beta_{(j,k+k)}^{(j,k+i)} = -q_{i+1}^T \left( \pi_{j,k+i} + \rho_{j+1} \sum_{t=1}^{k} \beta_{j,k+t}^{(j,k+i)} A \omega_{j,k+t} \right) \rho_{j+1} q_{i+1}^T A \omega_{j,k+k}; \]

20. **For** \( s = 1, \ldots, i-1 \)

21. \[ \beta_{j,k+s}^{(j,k+i)} = -q_{s+1}^T \left( \pi_{j,k+i} + \rho_{j+1} \sum_{t=1}^{k} \beta_{j,k+t}^{(j,k+i)} A \omega_{j,k+t} + \sum_{t=1}^{s} \beta_{j,k+t}^{(j,k+i)} (\omega_{j,k+t} - \bar{\omega}_{j,k+t}) \right) q_{s+1}^T (\omega_{j,k+s} - \bar{\omega}_{j,k+s}); \]

22. **End**

23. \[ \omega_{j,k+i} = \pi_{j,k+i} + \sum_{s=1}^{k} \beta_{j,k+s}^{(j,k+i)} (\rho_{j+1} A \omega_{j,k+s} + \omega_{j,k+s}) + \sum_{s=1}^{i-1} \beta_{j,k+s}^{(j,k+i)} \omega_{j,k+s}; \]

24. **End**

25. **If** \( i < k \)

26. \[ \alpha_{j,k+i+1} = \rho_{j+1} q_{i+1}^T \pi_{j,k+i} \rho_{j+1} q_{i+1}^T (\omega_{j,k+i} - \bar{\omega}_{j,k+i}); \]

27. \[ \pi_{j,k+i+1} = \pi_{j,k+i} - \frac{\alpha_{j,k+i+1}}{\rho_{j+1}} (\omega_{j,k+i} - \bar{\omega}_{j,k+i}); \]

28. \[ \omega_{j,k+i+1} = \pi_{j,k+i} - \alpha_{j,k+i+1} A \omega_{j,k+i}; \]

29. **End**

30. **End**
Some simplifications can be made to these operations, for instance, (a) resetting the scalar $\alpha_{j+1}$ in Line 27 to be $q^T_{i+1} \tilde{\pi}_{j+1} / q^T_{i+1} (\omega_{j+1} - \tilde{\omega}_{j+1})$ since $\alpha_{j+1}$ is only used in Lines 28 and 29 and the factor $\rho_{j+1}$ will be cancelled in Line 28; (b) merging Lines 9-14 and Lines 19-24 by adding a conditional control "if $j \geq 1$" in Line 16 and treating $\beta_l^{(j)}$ with $l < 0$ as zero; such $\beta_l^{(j)}$'s will appear in Lines 19-24 when $j = 0$; (c) introducing the auxiliary vector

$$d_{j+1} = \omega_{j+1} - \tilde{\omega}_{j+1},$$

which can be updated by using Lines 23 and 24 as

$$d_{j+1} = \pi_{j+1} - \tilde{\pi}_{j+1} + \rho_{j+1} \sum_{s=1}^{k} \beta_{(j-1)k+s}^{(j+1)} A \omega_{(j-1)k+s} + \sum_{s=1}^{i-1} \rho_{j+s}^{(j+1)} d_{j+s}.$$

With these changes, we rewrite Lines 8-30 as,

1. For $s = i, \ldots, k-1$ and $j \geq 1$

2. $\beta_{(j-1)k+s}^{(j+1)} = -q^T_{i+1} \left( \pi_{j+1} + \sum_{t=i}^{s-1} \rho_{j-1}^{(j+1)} A \omega_{(j-1)k+s} \right) / q^T_{i+1} d_{j-1}k+s$;

3. End

4. $\beta_{(j-1)k+k}^{(j+1)} = -q^T_{i} \left( \pi_{j+1} + \rho_{j+1} \sum_{t=i}^{k-1} \beta_{(j-1)k+k}^{(j+1)} A \omega_{(j-1)k+k} \right) / \rho_{j+1} q^T_{i} A \omega_{(j-1)k+k}$;

97
5. For $s = 1, \cdots, i - 1$

\[ \beta_{jk+s}^{(jk+i)} = -\frac{q_{s+1}^T (\pi_{jk+i} + \rho_{j+1} \sum_{t=1}^{k} \beta_{(j+1)k+t}^{(j+1)} A \omega_{(j-1)k+t} + \sum_{t=1}^{s-1} \beta_{jk+t}^{(j+1)} d_{jk+t})}{q_{s+1}^T d_{jk+s}}; \]

7. End

8. \[ d_{jk+i} = \pi_{jk+i} - \tilde{\pi}_{jk+i} + \rho_{j+1} \sum_{s=1}^{k} \beta_{(j-1)k+s}^{(j+1)} A \omega_{(j-1)k+s} + \sum_{s=1}^{i-1} \beta_{jk+s}^{(j+1)} d_{jk+s}; \]

9. \[ \omega_{jk+i} = \pi_{jk+i} + \sum_{s=1}^{k} \beta_{(j-1)k+s}^{(j+1)} (\rho_{j+1} A \omega_{(j-1)k+s} + \omega_{(j-1)k+s}) + \sum_{s=1}^{i-1} \beta_{jk+s}^{(j+1)} \omega_{jk+s}; \]

10. If $i < k$

11. \[ \alpha_{jk+i+1} = \frac{q_{i+1}^T \tilde{\pi}_{jk+i}}{q_{i+1}^T d_{jk+i}}; \]

12. \[ \tilde{\pi}_{jk+i+1} = \tilde{\pi}_{jk+i} - \alpha_{jk+i+1} d_{jk+i}; \]

13. \[ \pi_{jk+i+1} = \pi_{jk+i} - \rho_{j+1} \alpha_{jk+i+1} A \omega_{jk+i}; \]

14. End

in which Lines 1-9 can be further written as,

1. \[ z_d = \tilde{\pi}_{jk+i}; \quad z_{\omega} = \pi_{jk+i}; \quad z_{A \omega} = 0; \]

2. For $s = i, \cdots, k - 1$ and $j \geq 1$

3. \[ \beta_{(j-1)k+s}^{(j+1)} = -\frac{q_{s+1}^T z_d}{q_{s+1}^T d_{(j-1)k+s}}; \]

4. \[ z_d = z_d + \beta_{(j-1)k+s}^{(j+1)} d_{(j-1)k+s}; \]
5. \( z_\omega = z_\omega + \beta_{(j-1)k+k}^{(j+i)} \omega_{(j-1)k+k} \)

6. \( z_{A\omega} = z_{A\omega} + \beta_{(j-1)k+k}^{(j+i)} A\omega_{(j-1)k+k} \)

7. \textbf{End}

8. \( \beta_{(j-1)k+k}^{(j+i)} = -\frac{q_i^T \left( \pi_{jk+i} + \rho_{j+1} z_{A\omega} \right)}{\rho_{j+1} q_i^T A\omega_{(j-1)k+k}} \)

9. \( z_\omega = z_\omega + \beta_{(j-1)k+k}^{(j+i)} \omega_{(j-1)k+k} \)

10. \( z_{A\omega} = \rho_{j+1} \left( z_{A\omega} + \beta_{(j-1)k+k}^{(j+i)} A\omega_{(j-1)k+k} \right) \)

11. \( z_d = \pi_{jk+i} + z_{A\omega} \)

12. \textit{For} \( s = 1, \ldots, i - 1 \)

13. \( \beta_{jk+s}^{(j+i)} = -\frac{q_s^T z_d}{q_s^T d_{jk+s}} \)

14. \( z_d = z_d + \beta_{jk+s}^{(j+i)} d_{jk+s} \)

15. \( z_\omega = z_\omega + \beta_{jk+s}^{(j+i)} \omega_{jk+s} \)

16. \textbf{End}

17. \( d_{jk+i} = z_d - \pi_{jk+i} \)

18. \( \omega_{jk+i} = z_\omega + z_{A\omega} \)
As the approximate solution $x_l$ at step $l (= jk + i)$ of the ML($k$)BiCGSTAB algorithm, we define

$$x_{jk+l+1} = x_{jk+l} - \rho_{jk+l} \tilde{x}_{jk+l} + \alpha_{jk+l} \omega_{jk+l}$$  \hspace{1cm} (4.20)

for $j \geq 0$ and

$$x_{jk+i+1} = x_{jk+i} + \rho_{jk+i} \alpha_{jk+i+1} \omega_{jk+i}$$  \hspace{1cm} (4.21)

for $j \geq 0, k > i \geq 1$. One can readily verify by induction that $\pi_l$ is the residual vector of $x_l$, that is, $\pi_l = b - Ax_l$.

We are now ready to state the ML($k$)BiCGSTAB algorithm formally. As part of the algorithm, we choose the parameter $\rho_{jk+l}$ to minimize the 2-norm of the vector $\pi_{jk+l} = \rho_{jk+l} A \tilde{x}_{jk+l} + \tilde{x}_{jk+l}$, i.e., $\rho_{jk+l} = -\tilde{x}_{jk+l}^T A \tilde{x}_{jk+l} / \| A \tilde{x}_{jk+l} \|^2$. Noting that the data $q_i^T A \omega_{jk+i}$, $q_i^T d_{jk+i}$ and $q_i^T d_{jk+i}$ are repeatedly used inside each loop of the control variable $j$, to save the computational cost, we introduce a variable $c_{jk+i'}$ such that $c_{jk+i'} = q_i^T d_{jk+i}$ if $1 \leq i' \leq k$ and $c_{jk+i} = q_i^T A \omega_{jk+i}$ if $i' = k$. Moreover, since $\pi_l$ is the residual of $x_l$, we relabel $\pi_l$ as $r_l$. We also relabel $\omega_l$ and $\tilde{x}_l$ as $q_l$ and $u_l$ respectively.

**Algorithm 4.2. ML($k$)BiCGSTAB**

1. Choose an initial guess $x_0$ and $k$ vectors $q_1, q_2, \ldots, q_k$.
2. Compute $r_0 = b - Ax_0$ and set $g_0 = r_0$.
3. For $j = 0, 1, 2, \ldots$
4. \hspace{1cm} $u_{jk+i} = A g_{jk+i}$
5. \hspace{1cm} $c_{jk+i} = q_i^T u_{jk+i}$
6. \hspace{1cm} $\alpha_{jk+i} = q_i^T r_{jk+i} c_{jk+i}$

100
7. \( u_{j+k+1} = r_{(j-1)k+k} - \alpha_{j+k+1} u_{(j-1)k+k} \)
8. \( \rho_{j+1} = -u_{j+k+1}^T A u_{j+k+1} / \| A u_{j+k+1} \|^2 \)
9. \( x_{j+k+1} = x_{(j-1)k+k} - \rho_{j+1} u_{j+k+1} + \alpha_{j+k+1} \theta_{(j-1)k+k} \)
10. \( r_{j+k+1} = \rho_{j+1} A u_{j+k+1} + u_{j+k+1} \)
11. \( \text{For } i = 1, 2, \ldots, k \)
12. \( z_d = u_{j+k+i}; z_g = r_{j+k+i}; z_w = 0 \)
13. \( \text{For } s = i, \ldots, k-1 \text{ and } j \geq 1 \)
14. \( \rho_{(j-1)k+s}^{(j+k+i)} = -q_{s+1}^T z_d / c_{(j-1)k+s} \)
15. \( z_d = z_d + \rho_{(j-1)k+s}^{(j+k+i)} d_{(j-1)k+s} \)
16. \( z_g = z_g + \rho_{(j-1)k+s}^{(j+k+i)} \theta_{(j-1)k+s} \)
17. \( z_w = z_w + \rho_{(j-1)k+s}^{(j+k+i)} w_{(j-1)k+s} \)
18. \( \text{End} \)
19. \( \rho_{(j-1)k+k}^{(j+k+i)} = \frac{q_1^T (r_{j+k+i} + \rho_{j+1} z_w)}{\rho_{j+1} c_{(j-1)k+k}} \)
20. \( z_g = z_g + \rho_{(j-1)k+k}^{(j+k+i)} \theta_{(j-1)k+k} \)
21. \( z_w = \rho_{j+1} \left( z_w + \rho_{(j-1)k+k}^{(j+k+i)} w_{(j-1)k+k} \right) \)
22. \( z_d = r_{j+k+i} + z_w \)
23. \( \text{For } s = 1, \ldots, i-1 \)
24. \( \rho_{j+k+s}^{(j+k+i)} = -q_{s+1}^T z_d / c_{j+k+s} \)
25. \( z_d = z_d + \rho_{j+k+s}^{(j+k+i)} d_{j+k+s} \)
26. \( z_g = z_g + \rho_{j+k+s}^{(j+k+i)} \theta_{j+k+s} \)
27. \( \text{End} \)
28. \( d_{j+k+i} = z_d - u_{j+k+i} \)
29. \( g_{j+k+i} = z_g + z_w \)
30. \( \text{If } i < k \)
31. \( c_{j+k+i} = q_{i+1}^T d_{j+k+i} \)
32. \[ \alpha_{jk+i+1} = q^T_{j+1} u_{jk+i+1}/c_{jk+i}; \]
33. \[ u_{jk+i+1} = u_{jk+i} - \alpha_{jk+i+1} d_{jk+i}; \]
34. \[ x_{jk+i+1} = x_{jk+i} + \rho_{j+1} \alpha_{jk+i+1} g_{jk+i}; \]
35. \[ w_{jk+i} = A q_{jk+i}; \]
36. \[ r_{jk+i+1} = r_{jk+i} - \rho_{j+1} \alpha_{jk+i+1} w_{jk+i}; \]
37. \[ \text{End} \]
38. \[ \text{End} \]
39. \[ \text{End} \]

The denominators in the ML(k)BiCGSTAB algorithm appear only in the evaluations of \( \alpha \)'s and \( \beta \)'s and from the process of their derivations, these denominators differ from their counterparts in Algorithm 4.1, but can not be zero if the denominators in Algorithm 4.1 are not zero, assuming the \( \rho_j \)'s defined in Algorithm 4.2 are nonzero. Hence, under the assumption that \( \rho_{j+1} \neq 0 \) for all \( j \), we can see that if Algorithm 4.1 does not break down by zero division at some step, then neither does ML(k)BiCGSTAB at the same step. Moreover, since the residual vector \( r_{jk+i}^2 \) is by definition the vector \( \psi_{j+1}(A)r_{jk+i}^1 \), where \( r_{jk+i}^1 \) and \( r_{jk+i}^2 \) denote the corresponding residual vectors in Algorithms 4.1 and 4.2 respectively, and since \( r_{jk+i}^1 \in K_{jk+i+1}(v_0, A) \), \( r_{jk+i}^2 \in K_{jk+i+j+2}(v_0, A) \). Thus it is possible that \( r_{jk+i}^2 \) vanishes when \( jk+i+j+1 \geq \nu \), or \( jk+i \geq \nu - j - 1 \). On the other hand, \( r_{\nu}^1 \) must be zero because \( r_{\nu}^1 = 0 \).

**Theorem 4.3** Under the assumptions of Theorem 4.1 and suppose \( \rho_{j+1} \neq 0 \) for all \( j \), the ML(k)BiCGSTAB algorithm does not break down by zero division before
step \( \nu \) and an exact solution \(^2\) to (4.10) is obtained at or before step \( \nu \).

A similar remark to the one at the end of §3 can also be made here. Mathematically, ML(1)BiCGSTAB is equivalent to BiCGSTAB since it was established based on BiCG by using exactly the same techniques used in deriving BiCGSTAB.

In the case where \( k \geq \nu \), we can obtain an exact solution in the first loop of \( j \), i.e., \( j = 0 \), and the algorithm now can be regarded as a FOM algorithm (with the \( q_i \)'s appropriately chosen), for the reasons stated in the following. Since

\[
\begin{align*}
r_i^1 & \in v_0 - \text{span}\{Av_0, A^2v_0, \ldots, A^i v_0\} \\
\end{align*}
\]

from §3, we have

\[
\begin{align*}
r_i^2 & = \psi_1(A)r_i^1 \in r_i^F - \text{span}\{Ar_i^0, A^2 r_i^0, \ldots, A^i r_i^0\},
\end{align*}
\]

where \( v_0 = b - Ax_0 \), \( r_i^F = \psi_1(A) v_0 \) and \( r_i^1 \) and \( r_i^2 \) are the residual vectors of Algorithm 4.1 and ML(\( k \))BiCGSTAB respectively and \( 1 \leq i \leq \nu \). Thus if \( A \) is nonsingular and recalling that \( \psi_1(\lambda) = \rho_1 \lambda + 1 \), then

\[
\begin{align*}
x_i^2 & \in x_i^F + \text{span}\{r_i^F, A r_i^0, \ldots, A^{i-1} r_i^0\}
\end{align*}
\]

where \( x_i^2 \) denotes the approximate solution of ML(\( k \))BiCGSTAB, defined by (4.20) and (4.21), with residual \( r_i^2 \) and where \( x_i^F = x_0 - \rho_1 v_0 \) and \( r_i^F = b - Ax_i^F \). If at step \( i \) of the ML(\( k \))BiCGSTAB algorithm, we choose \( q_i(= A^0 q_1 = p_1) = \psi_1(A^T) \psi_1(A) v_0 \)

\(^2\)If the coefficient matrix \( A \) is singular, the system (4.10) may have more than one solutions.
and $q_i (= A^e q_i = p_i) = \psi_1(A^T)r_{i-1}^2$ for $2 \leq i' \leq i$, then

$$(r_0^F)^T r_i^2 = (\psi_1(A^T) \psi_1(A)v_0)^T r_i^1 = p_{i+1}^T r_i^1 = 0, \quad 1 \leq i,$$

and

$$(r_i^2)^T r_i^2 = (r_i^2)^T \psi_1(A)r_i^1 = p_{i+1}^T r_i^1 = 0, \quad 1 \leq i' \leq i - 1$$

by (4.12). In other words,

$$r_i^2 \perp \text{span}\{r_0^F, r_1^2, \cdots, r_{i-1}^2\}, \quad 1 \leq i \leq \nu.$$

As a result, the ML(k)BiCGSTAB ($k \geq \nu$) algorithm with the special choices for $q_i$'s described above is mathematically equivalent to the FOM algorithm defined by

$$x_i^F \in x_0^F + \text{span}\{r_0^F, A r_0^F, \cdots, A^{i-1} r_0^F\}$$

and

$$r_i^F \perp \{r_0^F, r_1^F, \cdots, r_{i-1}^F\},$$

where the initial guess $x_0^F$ and residual $r_0^F$ are defined as above.

It is quite straightforward to give a preconditioned version of ML(k)BiCGSTAB. Suppose we are solving the right-preconditioned system,

$$A M^{-1} y = b, \quad y = M x.$$

Directly applying the ML(k)BiCGSTAB algorithm to the system $A M^{-1} y = b$ and then recovering the $x$-approximation from the $y$-approximation
with the relation $y_i = Mx_i$ yields the following algorithm.

**Algorithm 4.3. ML(k)BiCGSTAB with Preconditioning**

1. Choose an initial guess $x_0$ and $k$ vectors $q_1, q_2, \ldots, q_k$.
2. Compute $r_0 = b - Ax_0$ and set $g_0 = r_0$.
3. For $j = 0, 1, 2, \ldots$
   4. $\tilde{g}(j-1)k+k = M^{-1}g(j-1)k+k$;
   5. $w(j-1)k+k = A\tilde{g}(j-1)k+k$;
   6. $c(j-1)k+k = q_1^T w(j-1)k+k$;
   7. $\alpha_{j+1} = q_1^T r(j-1)k+k / c(j-1)k+k$;
   8. $u_{j+1} = r(j-1)k+k - \alpha_{j+1} w(j-1)k+k$;
   9. $\tilde{u}_{j+1} = M^{-1}u_{j+1}$;
   10. $\rho_{j+1} = -u_{j+1}^T A\tilde{u}_{j+1} / \| A\tilde{u}_{j+1} \|^2$;
   11. $x_{j+1} = x(j-1)k+k - \rho_{j+1} \tilde{u}_{j+1} + \alpha_{j+1} \tilde{g}(j-1)k+k$;
   12. $r_{j+1} = \rho_{j+1} A\tilde{u}_{j+1} + u_{j+1}$;
   13. For $i = 1, 2, \ldots, k$
   14. $z_d = u_{j+1}$; $z_g = r_{j+1}$; $z_w = 0$;
   15. For $s = i, \ldots, k-1$ and $j \geq 1$
   16. $\beta_{j+1}^{(j+k+i)} = -q_1^T z_d / c(j-1)k+s$;
   17. $z_d = z_d + \beta_{j+1}^{(j+k+i)} d(j-1)k+s$;
   18. $z_g = z_g + \beta_{j+1}^{(j+k+i)} z_d$;
   19. $z_w = z_w + \beta_{j+1}^{(j+k+i)} w(j-1)k+s$;
   20. End
   21. $\rho_{j+1}^{(j+k+i)} = -q_1^T (r_{j+1} + \rho_{j+1} z_w) / \rho_{j+1} c(j-1)k+k$;
   22. $z_g = z_g + \rho_{j+1}^{(j+k+i)} g(j-1)k+k$;
\[ z_w = \rho_{j+1} \left( z_w + \beta_{(j+1)k+k}^{(j+1)} u_{(j+1)k+k} \right); \]

\[ z_d = r_{j+k+i} + z_w; \]

\[ \text{For } s = 1, \ldots, i - 1 \]

\[ \beta_{(j+k+i)}^{(j+1)} = -q_{s+1}^{(j+1)} \tilde{z}_d / c_{j+k+i}; \]

\[ z_d = z_d + \beta_{(j+k+i)}^{(j+1)} d_{j+k+i}; \]

\[ z_g = z_g + \beta_{(j+k+i)}^{(j+1)} g_{j+k+i}; \]

End

\[ d_{j+k+i} = z_d - u_{j+k+i}; \]

\[ g_{j+k+i} = z_g + z_w; \]

If \( i < k \)

\[ c_{j+k+i} = q_{i+1}^T d_{j+k+i}; \]

\[ \alpha_{j+k+i+1} = q_{i+1}^T u_{j+k+i} / c_{j+k+i}; \]

\[ u_{j+k+i+1} = u_{j+k+i} - \alpha_{j+k+i+1} d_{j+k+i}; \]

\[ \tilde{g}_{j+k+i} = M^{-1} g_{j+k+i}; \]

\[ x_{j+k+i+1} = x_{j+k+i} + \rho_{j+1} \alpha_{j+k+i+1} \tilde{g}_{j+k+i}; \]

\[ w_{j+k+i} = A \tilde{g}_{j+k+i}; \]

End

End

End

With suitable changes of variables, it may be shown that both the left and split preconditioning versions of ML(k)BiCGSTAB also lead to Algorithm 4.3 provided that \( q_1, q_2, \ldots, q_k \) are appropriately chosen. For the concepts of left, right and split preconditioning, one is referred to [50].
<table>
<thead>
<tr>
<th>Preconditioning ($M^{-1}v$)</th>
<th>$1 + 1/k$</th>
<th>Vector addition</th>
<th>3</th>
</tr>
</thead>
<tbody>
<tr>
<td>Matvec ($Av$)</td>
<td>$1 + 1/k$</td>
<td>Saxpy</td>
<td>$2.5k + 3.5 + 1/k$</td>
</tr>
<tr>
<td>dot product</td>
<td>$k + 2$</td>
<td>Scalar operation</td>
<td>$k + 3 - 1/k$</td>
</tr>
<tr>
<td>Scalar-vector</td>
<td>1</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Table 4.1: Average cost per step of the preconditioned ML($k$)BiCGSTAB.

Each loop of the control variable $j$ in Algorithm 4.3 involves solving $k + 1$ systems with coefficient matrix $M$, $k + 1$ matrix-vector multiplications with $A$, $k^2 + 2k$ dot products, $2.5k^2 + 3.5k + 1$ saxpy's, $3k$ vector additions, $k$ scalar-vector multiplications and $k^2 + 3k - 1$ scalar operations. Since there are $k$ steps in each loop of $j$, the average cost per step can be calculated and is listed in Table 4.1. Regarding the storage, the data $\{q_1, \ldots, q_k\}$, $\{d_{(j-1)k+i}, \ldots, d_{jk+i-1}\}$, $\{g_{(j-1)k+i}, \ldots, g_{jk+i-1}\}$ and $\{w_{(j-1)k+i}, \ldots, w_{jk+i-1}\}$ are used in the process at step $jk + i$ and hence they must be stored. Since they dominate the memory when $k$ is large, the storage of the algorithm is about $4kn$. We note that when $k = 1$ the cost is the same as BiCGSTAB’s and for large $k$ the cost tends to that of FOM.

4.4 Numerical Experiments

In this section, we shall illustrate the numerical convergence behavior of ML($k$)BiCGSTAB. We shall compare ML($k$)BiCGSTAB to BiCG, BiCGSTAB and GMRES($m$) [48] on a test suite of matrices from the Harwell-Boeing Collection [16]. For the implementation of these latter three methods, we used the versions described in [8]. All
the experiments were run in MATLAB 4.2c on a SUN Sparc station with machine precision about $10^{-16}$. As for the initial guesses and right-hand sides, we always chose $x_0 = 0$ and $b = [1, 1, \ldots, 1]^T$. For the initial vectors $q_1, q_2, \ldots, q_k$ in the ML($k$)BiCGSTAB algorithm, we first chose $k$ random vectors with independent and identically distributed entries from a normal distribution with mean 0 and variance 1 and then made them orthogonal to each other by using the modified Gram-Schmidt algorithm [31]. With $q_i$'s chosen randomly, the conditions of Theorem 4.1 are satisfied in most of cases and therefore we expect that the chance of the algorithm encountering breakdown is very rare. The iteration was stopped as soon as the true relative error $\|b - Ax\|_2/\|b\|_2$ was less than $10^{-7}$. Finally, all the figures plot the true relative residual versus the number of matrix-vector multiplies taken.

We ran all four methods, on a representative group of matrices from the Harwell-Boeing collection. The results are summarized in Table 4.2. We observe that, in terms of number of matvec's, ML(50)BiCGSTAB and ML(100)BiCGSTAB are always better than the other four methods, at least for this collection of matrices. The only exception is the matrix watt2 where only BiCG and GMRES converged. We can also see that ML($k$)BiCGSTAB for $k = 25$ is almost as good as $k = 50$, whereas $k = 100$ does not give much improvement over $k = 50$ in most cases. We believe that the improvement of ML($k$)BiCGSTAB over BiCGSTAB can be attributed to the use of multiple starting vectors. In prin-
inciple, ML(\(k\))BiCGSTAB can never be better than full GMRES, but as we can see from the table, it can be much better than restarted GMRES. We can also see from the table that ML(\(k\))BiCGSTAB and BiCG tend to converge and diverge more or less on the same subset of matrices, but ML(\(k\))BiCGSTAB typically requires many fewer matvec's when they all converge.

Next, we present the convergence history for three matrices from Table 4.2. These matrices are described below. We have used ML(30)BiCGSTAB in these examples.

Example 1. This example is the first matrix named IMPCOL D from the CHEMIMP group of the Harwell-Boeing collection. The order of the matrix is 425 and it has 1339 nonzero entries. In this example, no preconditioner was used and the convergence curves are plotted in Figure 4.1. BiCGSTAB encounter a breakdown after 450 matvec's.

Example 2. The matrix is the second one named ORSIRRI from the OILGEN group. The order of the matrix is 1030 and the number of nonzero entries is 6858. We first run the algorithms without preconditioning and then with ILU(0) preconditioning. The results are shown in Figures 4.2 and 4.3 respectively.

Example 3. This is the HOR131 matrix from the NNCENG group. The order is 434 and it has 4710 nonzero entries. The ILU(0) preconditioner was used and the result is plotted in Figure 4.4.

We observe that when all four methods converge (as in Examples 2 and 3),
<table>
<thead>
<tr>
<th>Matrix</th>
<th>Order</th>
<th>BiCG</th>
<th>Bicgstab</th>
<th>GMRES(100)</th>
<th>ML(25)</th>
<th>ML(50)</th>
<th>ML(100)</th>
</tr>
</thead>
<tbody>
<tr>
<td>1186bus</td>
<td>1138</td>
<td>4748</td>
<td>5872</td>
<td>1966</td>
<td>1384</td>
<td>1395</td>
<td></td>
</tr>
<tr>
<td>bcspwr06</td>
<td>1454</td>
<td>5810</td>
<td>14246</td>
<td>3187</td>
<td>2720</td>
<td>1899</td>
<td></td>
</tr>
<tr>
<td>bcsstk08</td>
<td>1074</td>
<td>15844</td>
<td>—</td>
<td>3859</td>
<td>1902</td>
<td>1242</td>
<td></td>
</tr>
<tr>
<td>bcsstk14</td>
<td>1806</td>
<td>29294</td>
<td>—</td>
<td>—</td>
<td>13315</td>
<td>6336</td>
<td></td>
</tr>
<tr>
<td>bcsstk19</td>
<td>817</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>bcsstm27</td>
<td>1224</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>can1054</td>
<td>1054</td>
<td>9008</td>
<td>—</td>
<td>4058</td>
<td>3126</td>
<td>2606</td>
<td></td>
</tr>
<tr>
<td>dwt1005</td>
<td>1005</td>
<td>1178</td>
<td>2934</td>
<td>673</td>
<td>625</td>
<td>645</td>
<td></td>
</tr>
<tr>
<td>cris1776</td>
<td>1176</td>
<td>1426</td>
<td>1530</td>
<td>666</td>
<td>532</td>
<td>499</td>
<td></td>
</tr>
<tr>
<td>fs5414</td>
<td>541</td>
<td>2738</td>
<td>2640</td>
<td>728</td>
<td>469</td>
<td>403</td>
<td></td>
</tr>
<tr>
<td>gr3090</td>
<td>900</td>
<td>76</td>
<td>52</td>
<td>38</td>
<td>40</td>
<td>40</td>
<td></td>
</tr>
<tr>
<td>gre1107</td>
<td>1107</td>
<td>—</td>
<td>b</td>
<td>—</td>
<td>8676</td>
<td>3262</td>
<td></td>
</tr>
<tr>
<td>hor131</td>
<td>434</td>
<td>—</td>
<td>—</td>
<td>1945</td>
<td>1268</td>
<td>1048</td>
<td></td>
</tr>
<tr>
<td>impcold</td>
<td>425</td>
<td>—</td>
<td>b</td>
<td>1619</td>
<td>910</td>
<td>597</td>
<td></td>
</tr>
<tr>
<td>jagmesh2</td>
<td>1009</td>
<td>1726</td>
<td>2958</td>
<td>—</td>
<td>1152</td>
<td>995</td>
<td>1129</td>
</tr>
<tr>
<td>jpwk991</td>
<td>991</td>
<td>100</td>
<td>58</td>
<td>49</td>
<td>55</td>
<td>33</td>
<td>55</td>
</tr>
<tr>
<td>lns511</td>
<td>511</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>lock1074</td>
<td>1068</td>
<td>o</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>lskp1270</td>
<td>1270</td>
<td>2492</td>
<td>4458</td>
<td>1628</td>
<td>1591</td>
<td>1445</td>
<td></td>
</tr>
<tr>
<td>mahindas</td>
<td>1258</td>
<td>—</td>
<td>b</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>mefe</td>
<td>765</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>mac1374</td>
<td>1374</td>
<td>—</td>
<td>b</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>nos3</td>
<td>900</td>
<td>494</td>
<td>384</td>
<td>1968</td>
<td>251</td>
<td>249</td>
<td>246</td>
</tr>
<tr>
<td>orsir1</td>
<td>1030</td>
<td>2068</td>
<td>3318</td>
<td>1270</td>
<td>838</td>
<td>781</td>
<td>772</td>
</tr>
<tr>
<td>plat1919</td>
<td>1919</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>pores2</td>
<td>1224</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>sleyr3</td>
<td>1000</td>
<td>o</td>
<td>—</td>
<td>o</td>
<td>o</td>
<td>o</td>
<td></td>
</tr>
<tr>
<td>sherman2</td>
<td>1080</td>
<td>—</td>
<td>b</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>watli1</td>
<td>1856</td>
<td>19406</td>
<td>—</td>
<td>1131</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
<tr>
<td>west0899</td>
<td>989</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td>—</td>
<td></td>
</tr>
</tbody>
</table>

Table 4.2: Comparison of Methods on a Representative Group of Matrices from the Harwell-Boeing Collection. ML(25), ML(50) and ML(100) stand for ML(25)Bicgstab, ML(50)Bicgstab and ML(100)Bicgstab respectively. The numbers in the table are number of matvec's. "n" means no convergence within 20n matvec's for BiCG and 10n for the other methods, "b" denotes breakdown, "o" denotes overflow.
Figure 4.1: Example 1: with no preconditioning

Figure 4.2: Example 2: with no preconditioning
Figure 4.3: Example 2: with ILU(0) preconditioning

Figure 4.4: Example 3: with ILU(0) preconditioning
ML(30)BiCGSTAB requires approximately the same or fewer matvec's than the other three methods. In fact, it can be significantly faster than the other three methods, as in Example 2. Moreover, ML(30)BiCGSTAB manages to converge when the other three methods fail, as in Example 1.

Finally, we present some numerical results to demonstrate the dependence of the performance of ML(\(k\))BiCGSTAB on the value of \(k\). In Figure 4.5, we plot the number of matvec's (scaled by \(1/10n\)) versus \(k\) for the two matrices ORSIRR1 and HOR131. In order to illustrate the improvement of ML(\(k\))BiCGSTAB over BiCGSTAB for \(k > 1\), we plot for \(k = 1\) the number of matvec's for BiCGSTAB instead of for the mathematically equivalent ML(1)BiCGSTAB. We observe that for both matrices there is a dramatic improvement in performance as \(k\) increases from 1. This behavior is typical for the matrices that we have tested and this can be partially observed from Table 4.2. Thus the advantage of ML(\(k\))BiCGSTAB can be realized even for small values of \(k\). On the other hand, we can also see that for large enough values of \(k\) (e.g. \(k > 10\) for ORSIRR1 and \(k > 30\) for HOR131), the performance is not sensitive to the value of \(k\). Thus, it is not crucial to choose an optimal value of \(k\) as long as \(k\) is large enough. We have also found that the performance is not sensitive to the specific choice of the random starting vectors \(q_i\)'s, provided that \(k\) is large enough. However, we should caution that the performance could be sensitive to the choice of \(q_i\)'s for small values of \(k\).

More testings are of course needed to better understand and assess the per-
Figure 4.5: Number of matvec's / 10n vs. $k$ for the matrices HOR131 (a) and ORSIR1 (b). "o" denotes no convergence within 10n matvec's. For $k = 1$, we plotted the number of matvec's for BiCGSTAB. Note that there is a dramatic improvement in performance as $k$ increases from 1, but that for $k \geq 30$, the performance is not sensitive to $k$. 
formance of ML\((k)\)BiCGSTAB, but we hope we have at least demonstrated the potential advantages of this new method.

4.5 Appendix

Here we give the detailed derivation of the coefficients \(\alpha_i\)'s and \(\beta_i\)'s of the ML\((k)\)BiCGSTAB algorithm.

\[
\alpha_{jk+1} = \frac{\eta_j^{(i)} p_{jk+i}^T r_{(j-1)k+k}}{\eta_j^{(i)} p_{jk+i}^T A g_{(j-1)k+k}} = \frac{\sum_{s=0}^{j} \eta_s^{(i)} p_{sk+i+1}^T r_{(j-1)k+k}}{\sum_{s=0}^{j} \eta_s^{(i)} p_{sk+i+1}^T A g_{(j-1)k+k}}
\]

\[
= \frac{\sum_{s=0}^{j} \eta_s^{(i)} q_{i+1}^T A^s r_{(j-1)k+k}}{\sum_{s=0}^{j} \eta_s^{(i)} q_{i+1}^T A^s+1 g_{(j-1)k+k}} = \frac{q_{i+1}^T \psi_j(A) r_{(j-1)k+k}}{q_{i+1}^T A \psi_j(A) g_{(j-1)k+k}}
\]

\[
= \frac{q_{i+1}^T \pi_{(j-1)k+k}}{q_{i+1}^T A \omega_{(j-1)k+k}}
\]

for \(0 \leq j, i = 1\);

\[
\alpha_{jk+i+1} = \frac{\eta_j^{(i)} p_{jk+i+1}^T r_{jk+i}}{\eta_j^{(i)} p_{jk+i+1}^T A g_{jk+i}} = \frac{\sum_{s=0}^{j} \eta_s^{(i)} p_{sk+i+1}^T r_{jk+i}}{\sum_{s=0}^{j} \eta_s^{(i)} p_{sk+i+1}^T A g_{jk+i}}
\]

\[
= \frac{\sum_{s=0}^{j} \eta_s^{(i)} q_{i+1}^T A^s r_{jk+i}}{\sum_{s=0}^{j} \eta_s^{(i)} q_{i+1}^T A^s+1 g_{jk+i}} = \frac{q_{i+1}^T \psi_j(A) r_{jk+i}}{q_{i+1}^T A \psi_j(A) g_{jk+i}}
\]

\[
= \frac{\rho_{j+i+1} q_{i+1}^T \psi_j(A) r_{jk+i}}{q_{i+i+1}^T (\psi_{j+1}(A) - \psi_j(A)) g_{jk+i}} = \frac{\rho_{j+i+1} q_{i+1}^T \tilde{\pi}_{jk+i}}{q_{i+i+1}^T (\omega_{jk+i} - \tilde{\omega}_{jk+i})}
\]
for $0 \leq j, 1 \leq i < k$;

$$
\beta_0^{(i)} = -\frac{p_i^T(\eta_1^{(1)} A r_i + \eta_0^{(1)} r_i)}{\eta_1^{(1)} p_i^T A g_0} = -\frac{q_i^T \psi_1(A) r_i}{\rho_1 q_i^T A g_0}
$$

$$
= -\frac{q_i^T \pi_i}{\rho_1 q_i^T A \omega_0}
$$

for $1 \leq i \leq k, j = 0$;

$$
\beta_s^{(i)} = -\frac{p_{s+1}^T (\eta_1^{(1)} A r_i + \beta_0^{(i)} \eta_1^{(1)} A g_0 + \sum_{t=1}^{s-1} \beta_t^{(i)} \eta_1^{(1)} A g_t)}{\eta_1^{(1)} p_{s+1}^T A g_s}
$$

$$
= -\frac{p_{s+1}^T (\psi_1(A) r_i + \beta_0^{(i)} \eta_1^{(1)} A g_0 + \sum_{t=1}^{s-1} \beta_t^{(i)} (\psi_1(A) - \psi_0(A)) g_t)}{p_{s+1}^T (\psi_1(A) - \psi_0(A)) g_s}
$$

$$
= -\frac{q_{s+1}^T (\pi_i + \beta_0^{(i)} \rho_1 A \omega_0 + \sum_{t=1}^{s-1} \beta_t^{(i)} (\omega_t - \bar{\omega}_t))}{q_{s+1}^T (\omega_s - \bar{\omega}_s)}
$$

for $1 \leq s < i \leq k, j = 0$;

$$
\beta_{(j-1)k+s}^{(j)} = -\frac{\eta_j^{(j)} p_{(j-1)k+s+1}^T A r_{(j-1)k+s} + \eta_j^{(j)} \sum_{t=i}^{s-1} \beta_t^{(j-1)k+t} p_{(j-1)k+s+1}^T A g_{(j-1)k+t}}{\eta_j^{(j)} p_{(j-1)k+s+1}^T A g_{(j-1)k+s}}
$$

$$
= -\frac{\eta_j^{(j)} p_{(j-1)k+s+1}^T r_{(j-1)k+i} + \rho_j \sum_{t=i}^{s-1} \beta_t^{(j-1)k+t} \eta_j^{(j-1)} p_{(j-1)k+s+1}^T A g_{(j-1)k+t}}{\rho_j \eta_j^{(j-1)} p_{(j-1)k+s+1}^T A g_{(j-1)k+s}}
$$

$$
= -\frac{\sum_{u=0}^{j} \eta_u^{(j)} p_{uk+s+1}^T r_{(j-1)k+i} + \rho_j \sum_{t=i}^{s-1} \beta_t^{(j-1)k+t} \sum_{u=0}^{j-1} \eta_u^{(j-1)} p_{uk+s+1}^T A g_{(j-1)k+t}}{\rho_j \sum_{u=0}^{j-1} \eta_u^{(j-1)} p_{uk+s+1}^T A g_{(j-1)k+s}}
$$

$$
= -\frac{\sum_{u=0}^{j} \eta_u^{(j)} q_{s+1}^T A u r_{(j-1)k+i} + \rho_j \sum_{t=i}^{s-1} \beta_t^{(j-1)k+t} \sum_{u=0}^{j-1} \eta_u^{(j-1)} q_{s+1}^T A u+1 g_{(j-1)k+t}}{\rho_j \sum_{u=0}^{j-1} \eta_u^{(j-1)} q_{s+1}^T A u+1 g_{(j-1)k+s}}
$$

116
\[
\begin{align*}
q_{j+1}^{r_i} & \psi_j(A) p_{j+1}^{r_i} + \rho_j \sum_{t=i}^{j-1} \beta_{(j-1)k+t}^{(j-1)i+1} q_{j+1}^{r_i} A \psi_{j-1}(A) g_{(j-1)k+t} \\
& \quad - \frac{q_{j+1}^{r_i} \psi_j(A) p_{j+1}^{r_i} + \sum_{t=i}^{j-1} \beta_{(j-1)k+t}^{(j-1)i+1} q_{j+1}^{r_i} (\psi_j(A) - \psi_{j-1}(A)) g_{(j-1)k+t}}{q_{j+1}^{r_i} (\psi_j(A) - \psi_{j-1}(A)) g_{(j-1)k+t}} \\
& = - \frac{q_{j+1}^{r_i} \left( \pi_{j+k+i} + \sum_{t=i}^{j-1} \beta_{(j-1)k+t}^{(j-1)i+1} (\omega_{(j-1)k+t} - \tilde{\omega}_{(j-1)k+t}) \right)}{q_{j+1}^{r_i} (\omega_{(j-1)k+s} - \tilde{\omega}_{(j-1)k+s})}
\end{align*}
\]

for \(1 \leq j, 1 \leq i \leq s \leq k-1;\)

\[
\beta_{(j-1)k+k}^{(j-1)i+1} = - \frac{\eta_{j+1}^{(j+1)} p_{j+1}^{r_i} A r_{j+k+i} + \sum_{t=i}^{j-1} \beta_{(j-1)k+t}^{(j-1)i+1} q_{j+1}^{r_i} A g_{(j-1)k+t}}{\eta_{j+1}^{(j+1)} p_{j+1}^{r_i} A g_{(j-1)k+k}}
\]

\[
= - \frac{\sum_{u=0}^{j+1} \eta_u^{(j+1)} p_{u+1}^{r_i} A r_{j+k+i} + \rho_{j+1} \sum_{t=i}^{j-1} \beta_{(j-1)k+t}^{(j-1)i+1} q_{j+1}^{r_i} A g_{(j-1)k+t}}{\rho_{j+1} \sum_{u=0}^{j} \eta_u^{(j)} q_{u+1}^{r_i} A g_{(j-1)k+k}}
\]

\[
= - \frac{\sum_{u=0}^{j+1} \eta_u^{(j+1)} q_{j+1}^{r_i} A u r_{j+k+i} + \rho_{j+1} \sum_{t=i}^{j-1} \beta_{(j-1)k+t}^{(j-1)i+1} \sum_{u=0}^{j} \eta_u^{(j)} q_{u+1}^{r_i} A u g_{(j-1)k+t}}{\rho_{j+1} \sum_{u=0}^{j} \eta_u^{(j)} q_{u+1}^{r_i} A u g_{(j-1)k+k}}
\]

\[
\beta_{(j-1)k+k}^{(j-1)i+1} = - \frac{q_{j+1}^{r_i} \psi_j(A) p_{j+1}^{r_i} A r_{j+k+i} + \rho_{j+1} \sum_{t=i}^{j-1} \beta_{(j-1)k+t}^{(j-1)i+1} q_{j+1}^{r_i} A \psi_{j-1}(A) g_{(j-1)k+t}}{\rho_{j+1} q_{j+1}^{r_i} A \psi_{j-1}(A) g_{(j-1)k+k}}
\]

\[
= - \frac{\sum_{t=i}^{j-1} \beta_{(j-1)k+t}^{(j-1)i+1} A \omega_{(j-1)k+t}}{\rho_{j+1} q_{j+1}^{r_i} A \omega_{(j-1)k+k}}
\]

for \(1 \leq j, 1 \leq i \leq k;\)

\[
\beta_{j+k+s}^{(j-1)i+1} = - \left( \eta_{j+1}^{(j+1)} p_{j+k+s+1}^{r_i} A r_{j+k+i} + \eta_{j+1}^{(j+1)} \sum_{t=i}^{k} \beta_{(j-1)k+t}^{(j-1)i+1} p_{j+k+s+1}^{r_i} A g_{(j-1)k+t} \right)
\]
\[ \eta^{(j+1)}_{j+1} \sum_{t=1}^{s-1} \beta^{(j+k+i)}_{j+k+t} P_{j+k+s+1} A g_{j+k+t} + \eta^{(j+1)}_{j+1} P_{j+k+s+1} T A g_{j+k+s} \]

\[ = - \left( \eta^{(j+1)}_{j+1} P_{(j+1)k+s+1} T A g_{j+k+i} + \rho^{(j+1)}_{j+1} \sum_{t=1}^{k} \beta^{(j+k+i)}_{(j-1)k+t} \eta^{(j)}_{j} P_{j+k+s+1} A g_{(j-1)k+t} + \rho^{(j+1)}_{j+1} \eta^{(j)}_{j} A g_{j+k+s} \right) \]

\[ = - \left( \sum_{u=0}^{j+1} \eta^{(j+1)}_{u} P_{u+k+s+1} T A g_{j+k+i} + \rho^{(j+1)}_{j+1} \sum_{t=1}^{k} \beta^{(j+k+i)}_{(j-1)k+t} \sum_{u=0}^{j} \eta^{(j)}_{u} P_{u+k+s+1} A g_{(j-1)k+t} + \rho^{(j+1)}_{j+1} \sum_{u=0}^{j} \eta^{(j)}_{u} P_{u+k+s+1} A g_{j+k+s} \right) \]

\[ = - \left( \sum_{u=0}^{j+1} \eta^{(j+1)}_{u} q_{s+1} T A u^{+1} g_{j+k+i} + \rho^{(j+1)}_{j+1} \sum_{t=1}^{k} \beta^{(j+k+i)}_{(j-1)k+t} \sum_{u=0}^{j} \eta^{(j)}_{u} q_{s+1} T A u^{+1} g_{(j-1)k+t} + \rho^{(j+1)}_{j+1} \sum_{u=0}^{j} \eta^{(j)}_{u} q_{s+1} T A u^{+1} g_{j+k+s} \right) \]

\[ = - \left( q_{s+1} T \psi^{(j)}_{j+1}(A) r_{j+k+i} + \rho^{(j+1)}_{j+1} \sum_{t=1}^{k} \beta^{(j+k+i)}_{(j-1)k+t} q_{s+1} T A \psi^{(j)}_{j+1}(A) g_{(j-1)k+t} + \rho^{(j+1)}_{j+1} q_{s+1} T A \psi^{(j)}_{j+1}(A) g_{j+k+s} \right) \]

\[ = - \left( q_{s+1} T \psi^{(j)}_{j+1}(A) r_{j+k+i} + \rho^{(j+1)}_{j+1} \sum_{t=1}^{k} \beta^{(j+k+i)}_{(j-1)k+t} q_{s+1} T A \psi^{(j)}_{j+1}(A) g_{(j-1)k+t} + \rho^{(j+1)}_{j+1} q_{s+1} T A \psi^{(j)}_{j+1}(A) g_{j+k+s} \right) \]
\[
\sum_{t=1}^{s-1} \beta_{jk+t}^{(j+k+i)} q_{s+1}^T (\psi_{j+1}(A) - \psi_j(A)) g_{jk+t} \Bigg/ q_{s+1}^T (\psi_{j+1}(A) - \psi_j(A)) g_{jk+s}
\]

\[
= -q_{s+1}^T \left( \pi_{jk+i} + \rho_{j+1} \sum_{t=1}^{k} \beta_{(j-1)k+t}^{(j+k+i)} A \omega_{(j-1)k+t} + \sum_{t=1}^{s-1} \beta_{jk+t}^{(j+k+i)} (\omega_{jk+t} - \tilde{\omega}_{jk+t}) \right) q_{s+1}^T (\omega_{jk+s} - \tilde{\omega}_{jk+s})
\]

for \( 1 \leq j, 1 \leq s < i \leq k \).
Bibliography


[40] D. Lê, Benefits and costs of solving Gaussian elimination related problems using randomization on block, systolic, and recursive block decomposition schemes, Ph.D. Thesis, 1995, Department of Computer Science, University of California, Los Angeles, CA 90095.


